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(Working Papers)

Credit supply and productivity growth

by Francesco Manaresi and Nicola Pierri

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# CREDIT SUPPLY AND PRODUCTIVITY GROWTH

by Francesco Manaresi\* and Nicola Pierri\*\*

## Abstract

We study the impact of bank credit supply on firm output and productivity. By exploiting a matched firm-bank database which covers all the credit relationships of Italian corporations over more than a decade, we measure idiosyncratic supply-side shocks to firms' credit availability. We use our data to estimate a production model augmented with financial frictions and show that an expansion in credit supply leads firms to increase both their inputs and their output (value added and revenues) for a given level of inputs. Our estimates imply that a credit crunch will be followed by a productivity slowdown, as experienced by most OECD countries after the Great Recession. Quantitatively, the credit contraction between 2007 and 2009 could account for about a quarter of the observed decline in Italy's total factor productivity growth. The results are robust to an alternative measurement of credit supply shocks that uses the 2007-08 interbank market freeze as a natural experiment to control for assortative matching between borrowers and lenders. Finally, we investigate possible channels: access to credit fosters IT-adoption, innovation, exporting, and the adoption of superior management practices.

**JEL Classification:** D22, D24, G21.

**Keywords:** credit supply, productivity growth, export, management, IT adoption.

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# 1 Introduction<sup>1</sup>

Does lenders' credit supply affect borrowing firms' productivity and, if so, how?

Aggregate productivity growth has declined in most OECD economies over the last decade, as illustrated by figure 1. While financial crises are found to induce strong and persistent recessions,<sup>2</sup> it is still an open question whether credit supply (or lack thereof) played a major role in generating (and/or sustaining) this productivity slowdown.<sup>3</sup>

In this paper, we estimate the effect of idiosyncratic changes in the supply of credit faced by Italian firms on their total factor productivity (TFP) growth. We focus on Italy because of the availability of detailed loan- and firm-level data on credit, inputs, and output. [Khwaja & Mian \(2008\)](#), [Chodorow-Reich \(2013\)](#), and [Amiti & Weinstein \(2017\)](#) exploit lender-borrower connections to provide evidence that negative shocks experienced by banks diminish credit supplied to borrowing firms and constrain those firms' investment and employment.<sup>4</sup> This paper extends the previous literature by looking at the impact of credit on productivity and by tracing its channels.<sup>5</sup>

The sign of the causal relationship between the availability of external finance and productivity is

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<sup>2</sup>Several authors found that financial crises, and the Great Recession in particular, are different than other recessions. See, for instance, [Cerra & Saxena \(2008\)](#), [Reinhart & Rogoff \(2009\)](#), [Reinhart & Rogoff \(2014\)](#), [Jordà \*et al.\* \(2013\)](#), and [Oulton & Sebastiá-Barriel \(2013\)](#). A contrasting view is expressed by [Stock & Watson \(2012\)](#).

<sup>3</sup>[Fernald \*et al.\* \(2017\)](#) document that the disappointing recovery of output after the Great Recession in the US is mainly due to low TFP dynamics, although they find it implausible that this is generated by financial shocks. Alternative explanations for the productivity slowdown are low business dynamism (e.g., [Decker \*et al.\* \(2014\)](#) and [Davis & Haltiwanger \(2014\)](#)), mismeasurement of digital goods (e.g. [Mokyr \(2014\)](#), [Feldstein \(2015\)](#), [Byrne \*et al.\* \(2016\)](#)), slowdown of technological progress (e.g. [Gordon \(2016\)](#), [Gordon \*et al.\* \(2015\)](#), [Bloom \*et al.\* \(2016\)](#), [Cette \*et al.\* \(2016\)](#)), and weak demand conditions ([Anzoategui \*et al.\* \(2016\)](#)). [Adler \*et al.\* \(2017\)](#) argue for the interaction of several factors, from greater uncertainty to an aging workforce. Focusing on Italy, [Hall \*et al.\* \(2008\)](#) underline the lack of product innovation as a pre-crisis productivity problem.

<sup>4</sup>See also [Bentolila \*et al.\* \(2013\)](#), [Greenstone \*et al.\* \(2014\)](#), [Cingano \*et al.\* \(2016\)](#), and [Bottero \*et al.\* \(2015\)](#).

<sup>5</sup>Some papers have studied how aggregate credit availability impacted TFP via a reallocation of inputs among heterogeneous firms. [Gopinath \*et al.\* \(2017\)](#) and [Cette \*et al.\* \(2016\)](#) argue that low interest rates have triggered unfavorable resource reallocations in southern Europe and [Linarello \*et al.\* \(2018\)](#) study the impact of credit supply on allocative efficiency among Italian firms. Yet, the direct impact of the credit crunch on firm's productivity is yet to be assessed.

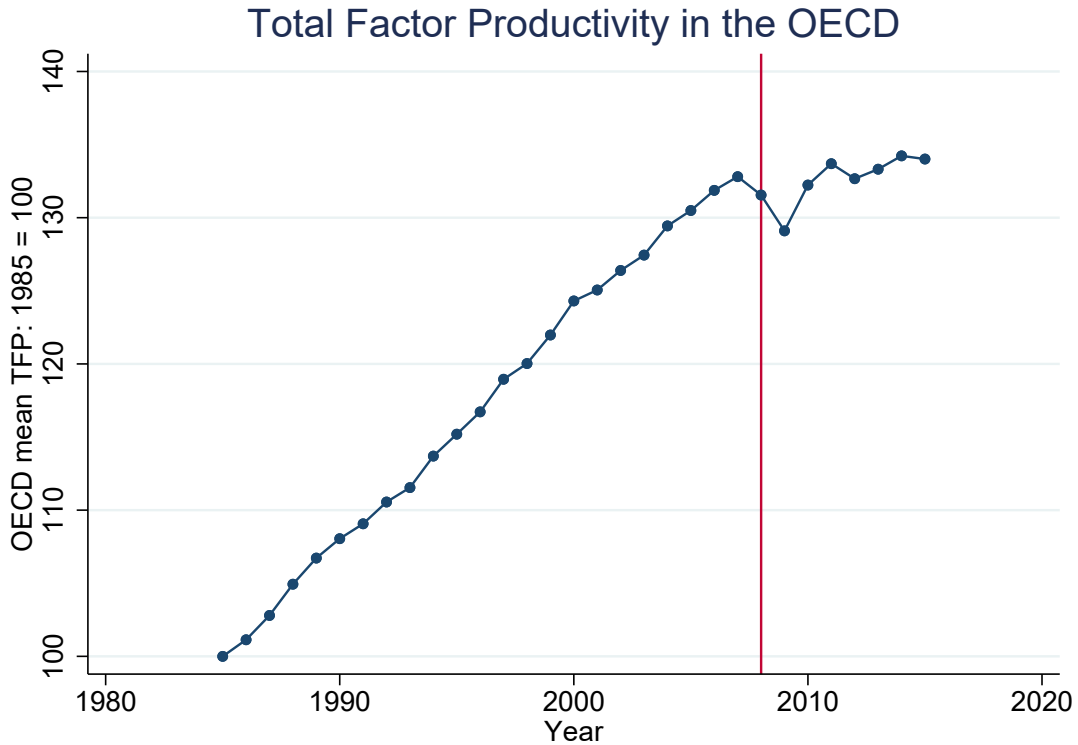


Figure 1: TFP provided by OECD (<https://data.oecd.org/lprdty/multifactor-productivity.htm>). Values in 1985 are normalized to 100 for each Country, with the (unweighted) average plotted. All countries with data since 1985 are included: AUS, BEL, CAN, DEU, DNK, ESP, FIN, FRA, GBR, IRL, ITA, JPN, KOR, NLD, PRT, SWE, and USA. The vertical line indicates year=2008.

theoretically and empirically ambiguous. Standard models of financial frictions assume that agents have an exogenous productivity, implying that credit constraints affect output only via reductions in the amount of capital used in production. Richer models can generate either a negative or a positive relationship. On the one hand, being forced to operate with fewer resources might spur innovation (Field, 2003)<sup>6</sup> and abundance might induce managers to stint their efforts or aggravate agency problems (Jensen, 1986). Moreover, if firms can choose between several business opportunities, they may be more likely to invest first in the most profitable ones. Then, as credit constraints become slacker, the marginal project may be of worse quality. Ates & Saffie (2016) apply the same intuition to firm entry rather than investment selection. All these factors may induce a negative relationship between credit supply and firm productivity and explain the empirical regularity that credit booms are often followed by sharp declines in output (see Schularick & Taylor (2012) and Gorton & Ordoñez (2016)).

On the other hand, credit availability may have positive effects on firm productivity, as it might

<sup>6</sup>He documents that the years after the Great Depression were the most technologically progressive decade of modern American History. In particular, he states: “In other sectors, for example railroads, the disruptions of financial intermediation and very low levels of capital formation associated with the downturn fostered a search for organizational innovations that enabled firms to get more out of what they had.”



support productivity-enhancing strategies.<sup>7</sup> Firms facing tighter credit constraints might invest less in R&D because of liquidity risk (Aghion *et al.*, 2010) and might acquire fewer intangible assets because it is more difficult to use them as collateral (Garcia-Macia, 2015). Credit-constrained firms might undertake less radical innovation (Caggese, 2016), while Midrigan & Xu (2014) emphasize the role of fixed costs. Additionally, negative credit shocks might hurt small firms by forcing managers/entrepreneurs to divert time and effort away from productivity improvements in order to create relationships with new lenders (“managerial inattention”).

We contribute to the relevant literature on four dimensions. First, we combine firm-bank matched data on credit granted by all financial intermediaries to all Italian incorporated firms over the period 1997-2013, with detailed balance-sheet information for a large sample of around 70,000 firms, to provide a complete picture of firm access to bank credit together with high-quality data on inputs acquisition and output for both large and small firms. Importantly, we are in a position to credibly study firm-level financial constraints without limiting our analysis to syndicated loans or public companies.

Second, we identify idiosyncratic credit supply shocks by exploiting two alternative empirical strategies: one based on bank-firm relationships and the other on a natural experiment. Unlike previous empirical studies on the link between finance and productivity, ours does not rely on self-reported measures of credit constraints, (potentially endogenous) proxies for financial strength, or local and industry-specific shocks (which might correlate with demand/technology dynamics).<sup>8</sup> Our main empirical strategy decomposes the growth rate of credit of each bank-firm pair into firm-year and bank-year components. The bank-year component reveals how different banks change the quantity of credit granted *to the same firm* and captures shocks to bank supply. This additive decomposition, closely related to the ones developed by Amiti & Weinstein (2017) and Greenstone *et al.* (2014), rests on assumptions related to the matching between banks and firms and the structure of substitution/complementarity between lenders. We provide novel tests for these hypotheses.

To aggregate bank-specific credit supply shocks at the firm level, we exploit the stickiness of bank-firm relations. Because of relationship lending, one lender’s expansion or contraction of credit disproportionately affects its existing borrowers. As a result, two firms serving the same market might experience different shocks to their ability to finance their operations because of pre-existing credit

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<sup>7</sup>Hall & Lerner (2010) survey some evidence about the difficulties of financing innovation.

<sup>8</sup>Like us, Huber (2017) and Franklin *et al.* (2015) examine the impact of credit supply on firms, using the lending contraction of German and UK banks to undertake a thorough estimation of the impact on output and employment and showing large negative impacts. They do not, however, look at total factor productivity, which is the focus of this paper. They both present results on labor productivity (value added per worker - see column (3) in Table VI of Huber (2017)), but provide evidence regarding the declining capital share and/or material inputs rather than declining TFP.

relations with different lenders.<sup>9</sup> We therefore average bank shocks at the firm level, using lagged credit shares as weights, to obtain a firm-specific credit supply shock. These shocks allow us to study the effect of credit on firm output and productivity both in “normal times” and during recessions.

Quantitatively, we find that a 1% increase in credit granted raises value-added TFP growth by around 0.1% and revenue TFP growth by 0.02-0.03%.<sup>10</sup> During the financial crisis of 2007-09, credit growth shrank by around 12%: our estimates imply that a similar supply-driven credit crunch would have induced between 12.5% and 30% of the average drop in firm TFP experienced by Italian firms during that period. The effect of credit on TFP growth lasts up to two years and does not revert afterwards, so that the impact on TFP is persistent over time and can partly explain the sluggish productivity growth after the financial crisis. Large firms and firms with more lending relationships, which are probably more easily able to substitute away from contracting lenders, are largely unaffected by credit supply. Effects are stronger in sectors where bank credit is more important; that is, manufacturing and industries characterized by higher leverage. Our results imply that a credit crunch can generate a productivity slowdown by depressing firm-level TFP. This effect may persist for several years.

To rule out the possibility that results are driven either by assortative matching between firms and banks or other confounding factors or by some forms of reverse causality, we use a second empirical strategy, which exploits the freezing of the interbank market in 2007-08 as a natural experiment. This shock affected Italian banks differently (and unexpectedly) according to their pre-crisis reliance on this source of funding.<sup>11</sup> We show that firms for which the credit crunch hit harder through their lenders experienced a lower growth rate of productivity afterwards. Firm exposure to the interbank market shock is found to be uncorrelated with pre-crisis growth potential and sensitivity to business cycle. This alternative identification strategy confirms the causal link between credit supply and productivity. Its estimated magnitude is significantly larger than the baseline estimates, suggesting that the productivity effects are stronger during financial turmoil.

Third, we argue that the standard production function estimation methods would not allow one to identify the causal effect of credit supply on productivity (see [De Loecker \(2013\)](#) for a conceptually

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<sup>9</sup>The relevance of this phenomenon has been documented in several countries, including the US ([Chodorow-Reich, 2013](#)), Italy ([Sette & Gobbi, 2015](#)), Spain ([Jiménez & Ongena, 2012](#)) and Pakistan ([Khwaja & Mian, 2008](#)).

<sup>10</sup>Since we do not observe firm-level output prices, productivity is the amount of revenues or value added (not the quantity of goods) generated for a given amount of inputs. In section 3.2, we clarify our terminology in relation to previous literature. We refer to the residuals of a production function as “revenue productivity” when output is measured by (log) net revenues and as “value added productivity” when output is measured by (log) value added. Measures of productivity estimated with revenues and quantities are usually found to be highly correlated.

<sup>11</sup>[Cingano et al. \(2016\)](#) show that firms that in 2006 were borrowing from banks more reliant on the interbank market experienced a stronger credit crunch, and that this, in turn, reduced investments.

analogous case regarding the effect of exporting on efficiency). Therefore, we enrich the production function estimation by allowing for heterogeneous credit constraints affecting both input acquisition and productivity dynamics.<sup>12</sup>

Fourth, we augment our dataset with information from administrative and survey-based sources in order to show that several productivity-enhancing activities, such as R&D, patenting, export, innovation, adoption of information technology (IT), and introduction of superior management practices, are stimulated by credit availability. These strategies increase productivity both in the short-run (e.g., IT-adoption) and in the long-run (e.g., R&D). Therefore, their sensitivity to credit can explain the immediate effects of a credit supply shock on TFP and also suggests that there are additional effects over a longer horizon. Finally, we discuss some indirect evidence that is consistent with the “managerial inattention” hypothesis.

Our results imply that disrupting access to external funds depresses output above and beyond the observable contraction of investments. This contributes to the theoretical literature on the aggregate effects of financial frictions (Brunnermeier *et al.*, 2012) and to the empirical investigation of frictions and investment decisions (see Fazzari *et al.* (1988) and Rauh (2006)).

Our findings are also an important complement to the literature on the misallocation of production factors. This strand of research has been thriving in recent years, in particular, since the seminal paper by Hsieh & Klenow (2009).<sup>13</sup> It studies how frictions—financial ones in particular—affect overall productivity by shaping the allocation of capital and other inputs between firms for a given distribution of idiosyncratic productivity. We show such financial frictions alter the location of productivity distribution. Therefore, any empirical investigation of the effect of a change in financial conditions on productivity should take into account *jointly* the impact on the allocative efficiency of inputs *and* the direct effect on firms’ productive efficiency. Our results also imply that part of the vast heterogeneity in firms’ productivity, which has been consistently found in several empirical works (Syverson, 2011), may be traced back to unequal access to external funds.

We show that the relationship between credit supply and productivity is positive and concave.

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<sup>12</sup> We build on the literature on estimation of the production function with control functions: Olley & Pakes (1996), Levinsohn & Petrin (2003), Akerberg *et al.* (2015), De Loecker & Warzynski (2012), De Loecker (2011), De Loecker & Scott (2016), Gandhi *et al.* (2011). In particular, Shenoy (2017) studies estimation of the production function when firms face heterogeneous and unobservable constraints that distort input acquisition but not productivity. Ferrando & Ruggieri (2015) and Peters *et al.* (2017a) are also related to our paper, since they estimate a production function and allow firm financial strength to affect productivity dynamics.

<sup>13</sup>A non-exhaustive list includes Bartelsman *et al.* (2013), Moll (2014), Asker *et al.* (2014), Midrigan & Xu (2014), Chaney *et al.* (2015), Buera & Moll (2015), Di Nola (2015), Gamberoni *et al.* (2016), Calligaris *et al.* (2016), Whited & Zhao (2016), Borio *et al.* (2016), Besley *et al.* (2017), Hassan *et al.* (2017), Gopinath *et al.* (2017), Schivardi *et al.* (2017), and Lenzu & Manaresi (2017). *Review of Economic Dynamics* had a special issue on “Misallocation and Productivity” in January 2013.

Negative shocks have larger effects than positive ones and credit supply is particularly important during a financial crisis. These empirical results highlight the fact that it is not only the *quantity* of credit that matters for productivity, but also its *stability*. Consequently, a credit crunch is likely to have a larger effect on TFP growth than a credit expansion of the same magnitude. Volatility of the banking sector’s credit supply is detrimental to firm productivity.

A large literature is interested in the link between finance and firm productivity. For instance, see [Schiantarelli & Sembenelli \(1997\)](#), [Gatti & Love \(2008\)](#), [Butler & Cornaggia \(2011\)](#), [Ferrando & Ruggieri \(2015\)](#), [Levine & Warusawitharana \(2014\)](#), and recent papers by [Duval \*et al.\* \(2017\)](#), [Dörr \*et al.\* \(2017\)](#), [Cavalcanti & Vaz \(2017\)](#), and [Mian \*et al.\* \(2017\)](#). Other papers study the impact of credit on specific productivity-enhancing strategies, such as R&D ([Bond \*et al.\* \(2005\)](#), [Aghion \*et al.\* \(2012\)](#), and [Peters \*et al.\* \(2017a\)](#)), innovation ([Benfratello \*et al.\* \(2008\)](#) and [Caggese \(2016\)](#)), intangible investments ([Garcia-Macia \(2015\)](#) and [de Ridder \(2016\)](#)), and exporting ([Paravisini \*et al.\* \(2014\)](#) and [Buono & Formai \(2013\)](#)). Access to other sources of external funds can also affect productive investments: for instance, [Bernstein \(2015\)](#) documents how IPOs change innovation strategies in the United States.

The paper proceeds as follows. Section 2 presents the data sources, discusses sample selection, and provides descriptive statistics of the main variables. Section 3.1 describes the estimation of idiosyncratic credit supply shocks. Section 3.2 presents a partial-equilibrium model of firm production with heterogeneous credit constraints, which is used to back out firm-level productivity. Section 4 shows that credit supply affects firm input acquisition and output. Section 5 contains our main results and deals with their robustness, heterogeneity, and persistence. Section 6 presents additional evidence from the 2007-08 collapse of the interbank market. Section 7 investigates the mechanisms driving the effect of credit supply on productivity. Section 8 concludes.

## 2 Data

To perform our empirical analysis, we combine detailed balance-sheet data with loan-level data from the Italian Credit Register and survey-based information on productivity-enhancing activities.

### 2.1 Firm balance-sheets: The CADS dataset

The Company Accounts Data System (CADS) is a proprietary database administered by CERVED-Group Ltd. for credit risk evaluation. It has collected detailed balance-sheet and income statement

information on non-financial corporations since 1982 and it is the largest sample of Italian firms for which data on actual investment flows are observed; net revenues of CADS firms account for about 70% of the total revenues of the private non-financial sector. Because this database is used by banks for credit decisions, the data are carefully controlled.

We estimate production functions for firms sampled in CADS from 1998 to 2013. Firm-level capital series are computed applying the perpetual-inventory method (PIM) on book-value of capital, investments, divestments, and sector-level deflators and depreciation rates.<sup>14</sup> Operating value added and intermediate expenditures are recorded in nominal values in profit-and-loss statements; we convert them in real terms using sector-level deflators from National Accounts. The baseline measure of labor is the wage bill, deflated using the consumer price index (CPI). Expenditures on intermediate inputs are deflated using a combination of sector-level deflator and regional-level CPI.<sup>15</sup> Throughout the paper, we use a Nace Rev.2 two-digit definition of industry. In addition, in a robustness exercise (section 5.1), we show that our main results are very similar if we use a finer four-digit definition.

From CADS, we also collect information on firm characteristics such as age, cash-flow, liquidity, assets, and leverage (total debt over assets). Their lagged values are used throughout the analysis in section 5 as firm-level time-varying controls.

## 2.2 Firm-bank matched data: The Italian Credit Register

The Italian Credit Register (CR), owned by the Bank of Italy, collects individual data on borrowers with total exposures (both debt and collateral) above €30,000<sup>16</sup> towards any intermediary operating in the country (including banks, other financial intermediaries providing credit, and special-purpose vehicles).<sup>17</sup> The CR contains data on the outstanding bank debt of each borrower, categorized into loans backed by accounts receivable, term loans, and revolving credit lines. CR data can be matched to CADS using each firm’s unique tax identifier.

For all the credit relationships of any Italian incorporated firm and any intermediary between 1998 and 2013, we measure net credit flows as the yearly growth rate (delta-log) of total outstanding

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<sup>14</sup>See [Lenzu & Manaresi \(2017\)](#) for details on PIM. We thank Francesca Lotti for providing capital series for an early version of this paper.

<sup>15</sup>Because some inputs might be bought on national rather than local markets, we assume that the price of intermediate inputs is the arithmetic mean of national price and national price deflated by local CPI.

<sup>16</sup>For instance, a borrowing firm with debt of €20,000 towards a bank appears in the CR if it also provides guarantees worth at least €10,000 to any another bank. The threshold was €75,000 before 2009.

<sup>17</sup>Following previous literature ([Amiti & Weinstein, 2017](#)), we include all financial intermediaries in the main analysis. We use the generic term “bank” for all of them. In a robustness exercise, available upon request, we show that our results are unchanged if we exclude firms which rely heavily on credit from non-bank intermediaries ( $\approx 0.33\%$  of total observations).

debt. We do not differentiate between different kinds of credit (for instance credit line versus loan), because the choice of which type of credit to increase/decrease is ultimately the result of strategic bargaining between banks and firms. We also focus on credit granted rather than on credit used, as the latter is more strongly affected by credit demand.

## 2.3 Additional data sources

While the baseline estimate of the effect of credit supply on productivity exploits CADS and CR, further enquiries into the channels that drive this effect and several robustness checks of our analyses rely on additional data sources.

To test whether estimates of credit supply shocks are robust to assortative matching between firms and banks (see section 3.1), we control for past interest rates charged by banks to firms. This information is available from the TAXIA database, administered by the Bank of Italy, for a large sample of Italian banks (encompassing over 70% of all credit granted to the Italian economy). Interest rates are computed as the ratio of interest expenditures to the quantity of credit used.

For our study of the consequences of the 2007-2008 interbank market collapse as an exogenous change in credit supply (section 6), we obtain information on banks assets, ROA, liquidity, capital ratio, and their interbank liabilities and assets from the Supervisory reports.

In Section 7, we study the relevance of specific productivity-enhancing activities that are fostered by credit supply. These include IT-adoption, R&D expenditures, patenting, and export. Such information is difficult to identify using balance-sheet data, because reporting by firms is generally non-compulsory. For this reason, we complement CADS with two sources of data. Data on IT-adoption, R&D, and export come from the INVIND Survey, administered by the Bank of Italy. INVIND is a panel of around 3,000 firms, representative of Italian firms with more than 20 employees and active in manufacturing and private services. For patent applications to the European Patent Office, we use the PatStat database. In particular, we exploit a release prepared by the Italian Association of Chambers of Commerce (UnionCamere), which matches all patent applications made during 2000-2013 with the tax identifiers of all Italian firms. We also obtain data on management practices for more than 100 manufacturing companies from the World Management Survey (see section 7).

## 2.4 Sample selection and descriptive statistics

Our main analysis is based on two samples. We use (a) a relationship-level dataset, in which an observation corresponds to a bank-firm-year triplet, to identify credit supply shocks and (b) a firm-level dataset, in which observations correspond to firm-year pairs and credit supply shocks are aggregated across banks, to estimate production functions.

The relationship-level dataset is based on the CR data. It consists of all relationships between incorporated firms and financial intermediaries during 1997-2013. The resulting dataset consists of 13,895,537 observations and is composed of 852,196 unique firms and one 1,008 banks per year.

To estimate production functions, we consider all firms in CADS that report positive revenues, capital, labor cost, and intermediate expenditures, so that a revenue production function can be estimated. As a result, we exclude around one-fifth of the original CADS dataset: the final sample consists of 76,542 firms, corresponding to 656,960 firm-year observations. This dataset is used to estimate all the baseline regressions. Table 1 reports the main variables from the firm-level dataset for both the whole sample and for manufacturers.

To provide preliminary descriptive evidence that bank credit is a relevant source of finance for Italian firms, we study the credit intensity of firms' activity. We define the credit intensity of firm  $i$  at time  $t$  as the ratio of total credit granted at the end of year  $t - 1$  to the net revenues of year  $t$ . On average, manufacturers are granted 43 cents for each euro of revenues generated, while this figure is only 34 cents for non-manufacturers. Appendix figure A.1 shows that credit-intense companies are larger in non-manufacturing sectors, but not in manufacturing. Appendix figure A.2 shows that industries with a higher capital-to-labor ratio are more credit-intensive.

## 3 Theoretical Framework

We investigate the relation between credit supply and productivity. As a first step, we consider an empirical model to disentangle idiosyncratic shocks to credit supply from shocks to credit demand and shocks to the general economic context (section 3.1). We then build a model of production with heterogeneous credit constraints to recover firm TFP (section 3.2).

### 3.1 Credit supply shocks

We define a credit supply shock as any change in bank-specific factors affecting a bank's ability and willingness to provide credit to firms. Banks are heterogeneous in their exposure to different

macroeconomic risks (Begenau *et al.*, 2015). This heterogeneity can arise because of differences in liabilities, assets or capital.<sup>18</sup>

Total credit granted to firm  $i$  at the end of year  $t$  equals the sum of credit granted by all existing intermediaries  $b$ :  $C_{i,t} = \sum_b C_{i,b,t}$ . We define firm  $i$  and bank  $b$  to have a pre-existing lending relation in period  $t$  if and only if  $C_{i,b,t-1} > 0$ . Credit granted  $C_{i,b,t}$ , is an equilibrium quantity which depends on both supply and demand factor, as well as on aggregate shocks. We collect all the observable and unobservable factors that determine the idiosyncratic supply of credit to corporations from bank  $b$  in year  $t$  into the vector  $S_{b,t}$ . For instance, bank-specific capital, cost of funds, and lending strategies may all be components of  $S_{b,t}$ . Similarly, let  $D_{i,t}$  be the vector of observables and unobservables shaping firm  $i$ 's demand for credit and its desirability as a borrower, such as productivity, size, and leverage. In addition, credit may be affected by firm-bank specific factors, such as the length of the pre-existing lending relationship or the quantity of credit previously provided by the bank to the firm (affecting the incentive to evergreen). We collect these match-specific covariates in the vector  $X_{i,b,t}$ . Finally, aggregate factors affecting all intermediaries and borrowers, such as aggregate demand or the monetary and fiscal stance, are collected in  $J_t$ .

### Assumption 1

$\exists$  some smooth, unknown function  $C(\cdot)$  such that:

$$\frac{C_{i,b,t}}{C_{i,b,t-1}} = \frac{C(J_t, D_{i,t}, S_{b,t}, X_{i,b,t})}{C(J_{t-1}, D_{i,t-1}, S_{b,t-1}, X_{i,b,t-1})} \quad (1)$$

While this assumption is very general, it nonetheless limits the substitution patterns amongst different lenders. Indeed, it rules out the impact of other banks' idiosyncratic shocks  $S_{b',t}$  on credit granted by  $b$  to  $i$ . In appendix A.1, we show that the exclusion of other banks' supply from equation (1) does not significantly affect our estimate of idiosyncratic credit supply shocks.

Log-linearizing equation (1) yields:

$$\Delta c_{i,b,t} = j_t + \Delta d'_{i,t} c_1 + \Delta s'_{b,t} c_2 + \Delta x'_{i,b,t} c_3 + approx_{i,b,t} \quad (2)$$

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<sup>18</sup>For instance, Khwaja & Mian (2008) show that the Pakistani banks that relied more on dollar deposits experienced stronger liquidity shocks after the unexpected nuclear tests in 1998. Chodorow-Reich (2013) uses US banks' connections to Lehman Brothers and exposure to mortgage-backed securities as an instrument for their financial health. In section 6, we exploit heterogeneity in reliance on the Interbank market as a source of exogenous variation during the credit crunch in Italy.



We define the credit supply shock of bank  $b$  in period  $t$  to be  $\Delta s'_{b,t}c_2$ . The idiosyncratic credit supply shock experienced by firm  $i$  in period  $t$  is a function of  $\Delta s'_{b,t}c_2$  for all the previously connected banks. Decomposition (2) can be written as:

$$\Delta c_{i,b,t} = j_t + d_{i,t} + \phi_{b,t} + \epsilon_{i,b,t}$$

where:  $j_t$  is the mean growth rate of credit in the economy,  $\phi_{b,t}$  is the change in credit granted explained by bank  $b$ 's supply factors,  $d_{i,t}$  is the change in credit granted explained by firm  $i$  factors, and  $\epsilon_{i,b,t}$  is the sum of a matching specific shock  $\Delta x'_{i,b,t}c_t^3$  and the approximation error  $approx_{i,b,t}$ .

### Assumption 2

$$\epsilon_{i,b,t} \perp \mathcal{D}_i, \mathcal{S}_b$$

where  $\mathcal{D}$  and  $\mathcal{S}$  are sets of dummy variables indicating the identities of the borrower and lender.

Furthermore, without loss of generality, we normalize  $E[d_{i,t}] = E[\phi_{b,t}] = 0$ . We apply OLS to estimate equation (3.1). Under assumption 2, the bank $\times$ year fixed effects ( $\phi_{b,t}$ ) are unbiased estimates of  $\Delta s_{b,t}$ . We focus on corporations having multiple relations in order to estimate bank-idiosyncratic shocks by exploiting within-firm-and-time variability. This allows us to condition for time-varying observables and unobservables at the borrower level.<sup>19</sup>

[Amiti & Weinstein \(2017\)](#) (AW hereafter) study the identification of model (3.1). They show that assumption 2 holds without loss of generality, as long as one is willing to conveniently “relabel” the firm and bank fixed effects. That is, one can write the idiosyncratic component  $\Delta x_{i,b,t}$  as  $\Delta x_{i,b,t} = a_{i,t} + b_{b,t} + e_{i,b,t}$ , where  $a$  and  $b$  are the linear projections of  $\Delta x_{i,b,t}$  on dummies for bank and firm identity and  $e_{i,b,t}$  is uncorrelated with these dummies by construction. Therefore, bank fixed effects in (3.1) correspond to  $\phi_{b,t}^{AW} = \phi_{b,t} + c^3 \cdot b_{b,t}$ , which are the parameters of interest in AW’s empirical analysis. In fact, AW show that the idiosyncratic match-specific terms do not affect the bank aggregate lending. In our study, however, we are interested in identifying the role of pure supply-side factors,  $\Delta s_{b,t}$ , so that the orthogonality assumption (assumption 2) does not come without loss of

<sup>19</sup>Because we are using a delta-log approximation, the expected values are intended to be conditional on credit by bank  $b$  to firm  $i$  being positive in both  $t$  and  $t - 1$ . In a robustness exercise, available upon request, we compute the model by measuring growth rates as suggested by [Davis et al. \(1996\)](#) ( $\Delta c_{i,b,t} = 2 \cdot \frac{C_{i,b,t} - C_{i,b,t-1}}{C_{i,b,t} + C_{i,b,t-1}}$ ), which we can compute as long as credit is positive in either  $t$  or  $t - 1$  or both.

generality. In particular, it limits the interaction between demand and supply shocks (which enter the approximation error) and restricts the correlation between match-level covariates and bank or firm factors. We argue in appendix A.1 that this assumption is testable: we focus on two potential source of omitted variables in  $\epsilon_{i,b,t}$  which may bias our estimate of supply-side shocks: substitution (or complementarity) patterns (such as those discussed in assumption 1) and relation characteristics. We show that our results on the impact of credit supply shocks on productivity are unaffected by the inclusion of these controls in the estimation of credit supply shocks. We therefore rely on the simpler specification in equation (3.1) for our main analysis.

In this paper, we study how borrowers’ inputs acquisition and output production are affected by lenders’ supply. Consequently, the cornerstone of the empirical strategy is a firm-level measure of credit supply shocks. To move from the bank-level measure of equation (3.1) to its firm-level counterpart, we rely on the intuition of the “lending channel” (Khwaja & Mian, 2008): borrower-lender relationships are valuable because they help mitigate information asymmetry, limited commitment, or other problems which might generate credit rationing (Petersen & Rajan, 1994). Consequently, they are sticky: changes in credit supplied by a bank have a disproportionately large effect on the firms with which it already has established credit relations.<sup>20</sup> Obviously, a firm connected to a bank whose supply contracts can always apply to another bank for credit (see below). Yet, as long as credit from an unconnected bank is less likely or more costly, substitution between lenders will be imperfect. The empirical relevance of this phenomenon has been shown in several previous studies. We exploit this well-established fact to identify firm-specific credit supply shocks.

As a simple benchmark, we assume that the strength of a firm-bank relationship is proportional to the amount of credit granted. Therefore, we measure the shock to credit supply faced by firm  $i$  in period  $t$  as

$$\phi_{i,t} = \sum_b \phi_{b,t} \cdot \frac{C_{b,i,t-1}}{\sum_{b'} C_{b',i,t-1}} \quad (3)$$

A histogram of  $\phi_{i,t}$  is provided in figure 2. Although the estimation of  $\phi_{b,t}$  is performed considering only firms with multiple banking relations, the variable  $\phi_{i,t}$  is defined for all firms which have some

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<sup>20</sup>Further causes of stickiness may encompass personal or political connections between firms’ and banks’ managers. Stickiness may be considered a credit market friction because it may prevent credit to flow to the firms with the best investment opportunities. Our analysis abstracts from any welfare consequences of relationship lending and focuses on one of its empirical implications.

credit granted in year  $t - 1$ .

Two empirical findings validate this measure of credit supply shocks. First, we expect a positive supply shock to decrease the number of loan applications to new lenders, while we expect a positive demand shock to increase these applications. Appendix A.2 shows that an increase of our measures of credit supply shock is indeed associated with fewer loan applications on both the intensive and extensive margin. Second, appendix D shows that our measure responds negatively to the freeze of the interbank market, which was the trigger of the credit crunch in Italy (see section 6 for details). In appendix A.3, we study the relation between credit supply shocks and some determinants of bank credit supply, such as the crowding-out of sovereign debt, M&A episodes and balance-sheet strength, and we present qualitative results in line with economic intuition and previous literature.

### 3.2 Production with heterogeneous financial frictions

We propose an empirical model to estimate firms’ production functions and recover their idiosyncratic productivity. We augment the classical production function estimation framework with a control function (Akerberg *et al.*, 2007) by adding two elements: a set of credit constraints and a modified law of motion for productivity dynamics. This section presents the main elements of the model; details can be found in appendix B.1. Uppercase letters denote variables in levels, while lowercase letters denote natural logarithms.

Firm  $i$  operating in industry  $s$ , in year  $t$ , combines capital ( $k_{i,t}$ ), labor ( $l_{i,t}$ ), and intermediate inputs ( $m_{i,t}$ )—which are also referred to as “materials”—to generate sales, ( $Y_{i,t}$ ) according to an industry-specific production function  $f(\cdot)$ , known up to a set of parameters  $\beta_s$ . Each firm has an idiosyncratic Hicks-neutral productivity  $\omega_{i,t}$ :

$$Y_{i,t} = \exp\{\omega_{i,t} + f(l_{i,t}, k_{i,t}, m_{i,t}, \beta_s)\}$$

As is common in the literature (Olley & Pakes, 1996), we assume that productivity can be decomposed into a structural component ( $\tilde{\omega}_{i,t}$ ) and an i.i.d. error term ( $\epsilon_{i,t}^Y$ ), which is unknown to the firm when production decisions are made:

$$\omega_{i,t} = \tilde{\omega}_{i,t} + \epsilon_{i,t}^Y$$

Intermediate inputs are flexibly chosen every period in order to maximize variable profits (sales minus cost of labor and intermediate inputs). Then, if firm  $i$  is unconstrained, the amount of materials  $m^{unc}$  will solve:

$$\frac{\partial \exp\{f(l_{i,t}, k_{i,t}, m^{unc}, \beta) + \tilde{\omega}_{i,t}\}}{\partial m} = P_{p,t}^M \quad (4)$$

where  $P_{p,t}^M$  is the price of materials faced by firm  $i$ , which might depend on its location  $p$ .<sup>21</sup> In section 4, we provide evidence that firms acquire less inputs when they receive negative credit supply shocks. Relying on the first-order condition in (4) would be misleading if firms face heterogeneous credit constraints. Therefore, we allow for the possibility that intermediate inputs (and other inputs) face financially generated constraints:

$$m_{i,t} \leq m_{i,t}^{max} = \log \frac{1}{P_{p,t}^M} K_{i,t-1} \cdot \Gamma(B_{i,t-1}, \phi_{i,t}, \tilde{\omega}_{i,t})$$

where  $B_{i,t-1}$  is previous-period debt and  $\Gamma$  is an unknown function. Similar constraints<sup>22</sup> are standard in the literature on financial frictions, such as Moll (2014), Buera & Moll (2015), and Gopinath *et al.* (2017), and they can be micro-founded by several market failures. We innovate by allowing them to depend on firm TFP and credit supply shocks. The results of the paper hold if we exclude credit rationing and, alternatively, if we assume that firms face heterogeneous costs of external funds. High-productivity firms might be considered more reliable borrowers and might therefore be allowed to borrow more, *ceteris paribus*. We thus assume that  $\Gamma$  is strictly increasing in its third argument. The quantity of intermediate inputs acquired by firm  $i$  is:

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<sup>21</sup>Gandhi *et al.* (2011) show that most of the estimation procedures based on the control function approach fail to identify the elasticity of output with respect to the flexible inputs (e.g., intermediate inputs). De Loecker & Scott (2016) argue that a researcher can overcome this non-identification result under the assumption that firms face heterogeneous and autocorrelated input prices. The authors rely on firm-level wages to estimate their model. However, heterogeneity in wages might reflect heterogeneous worker quality or productivity. We, instead, allow local price shocks to affect input real prices and recover all the production function parameters.

<sup>22</sup>Capital is the only potentially constrained input in most models.

$$m_{i,t} = \min\{m_{i,t}^{max}, m_{i,t}^{unc}\} := m(x_{i,t}, \tilde{\omega}_{i,t}, \phi_{i,t}) \quad (5)$$

where  $m(\cdot)$  is unknown and  $x_{i,t}$  is a vector containing firm-level inputs (capital, lagged capital, and labor), prices, and lagged debt. Under standard assumptions, the optimal value of materials is increasing in productivity  $\tilde{\omega}_{i,t}$ , equation 5 can therefore be “inverted” (see [Olley & Pakes \(1996\)](#) and [Levinsohn & Petrin \(2003\)](#)). That is, there exists an (unknown) function  $h$  such that:

$$\tilde{\omega}_{i,t} = h(x_{i,t}, m_{i,t}, \phi_{i,t})$$

Therefore, log sales can be written as:

$$y_{i,t} = \Psi(x_{i,t}, m_{i,t}, \phi_{i,t}) + \epsilon_{i,t}^Y$$

where  $\Psi(x_{i,t}, m_{i,t}, \phi_{i,t}) = h(x_{i,t}, m_{i,t}, \phi_{i,t}) + f(l_{i,t}, k_{i,t}, m_{i,t}, \beta_s)$ . Following the previous literature, we assume a law of motion for productivity:

$$E_t[\omega_{i,t} | \mathcal{I}_{t-1}] = E_t[\omega_{i,t} | \omega_{i,t-1}, \phi_{i,t-1}] = g_t(\omega_{i,t-1}, \phi_{i,t-1}) \quad (6)$$

where  $\mathcal{I}_{t-1}$  is the firm information set at  $t-1$  and  $g_t(\cdot)$  is unknown. We innovate by allowing credit supply to affect productivity dynamics. It would not be correct to estimate the production function without including financial frictions in the productivity dynamics and regress the productivity residuals on financial variables. An analogous problem is highlighted in [De Loecker \(2013\)](#) discussion of the measurement of productivity gains from exporting. Let us also define the productivity innovation as  $\zeta_{i,t} := \tilde{\omega}_{i,t} - E[\tilde{\omega}_{i,t} | \mathcal{I}_{t-1}]$ . Equation (6) implies moment conditions:

$$E [\zeta_{i,t} | \mathcal{I}_{t-1}] = E [\zeta_{i,t} | z_{i,t-1}] = E \left[ \begin{array}{c} \Psi_{i,t} - f(l_{i,t}, k_{i,t}, m_{i,t}, \beta) - \\ g_t(\Psi_{i,t-1} - f(l_{i,t-1}, k_{i,t-1}, m_{i,t-1}, \beta), \phi_{i,t-1}) \end{array} \middle| z_{i,t-1} \right] = 0 \quad (7)$$

where  $z_{i,t-1}$  contains lagged values of investments, labor, materials, and other variables. Estimation of the model is performed in two stages. In the first stage, we estimate the function  $\Psi$  as  $\Psi_{i,t} = E[y_{i,t} | x_{i,t}, m_{i,t}, \phi_{i,t}]$ . In the second stage, we rely on (7) to estimate the structural parameter of interest  $\beta_s$ . Table A.2 presents some descriptive statistics. Finally, we can back out firm-level productivity as residuals from  $\omega_{i,t} = y_{i,t} - f(k_{i,t}, l_{i,t}, m_{i,t}, \beta_s)$  or, in the value-added case,  $\omega_{i,t} = va_{i,t} - f(k_{i,t}, l_{i,t}, \beta_s)$ .

As detailed in section 2, we observe balance-sheet and income statements but do not observe firm-level output prices. Therefore, this paper is about the ability of firms to transform inputs into sales and value added and not (only) about their technical efficiency. Our measure of productivity is referred to as “productivity” in several empirical studies, such as [Olley & Pakes \(1996\)](#), and as  $tfpr^{rr}$  (or “regression-residual total factor revenue productivity”) in [Foster \*et al.\* \(2017\)](#). Furthermore, our measure of productivity is proportional to the empirical estimate of (log)  $TFPQ$  (or “total factor quantity productivity”) in [Hsieh & Klenow \(2009\)](#). Our choice in this regard is somewhat constrained, as no firm-level data on product-level prices are available to economists for a sufficiently large number of Italian firms. Appendix C.1 provides a more detailed treatment of the topic and contains a brief discussion of the pros and cons of relying on revenues to estimate TFP.

## 4 Credit Supply Shocks and Firm Production

Is a firm’s production affected by the credit supply of its lenders? If credit frictions are not important, the amount of credit a firm receives should be unaffected by the supply shocks of its lenders. In a frictionless world, a firm’s policy function might be affected by aggregate financial conditions but should not be shaped by the idiosyncratic shocks hitting any specific lender. Therefore, we estimate:

$$\Delta x_{i,t} = \psi_i + \psi_{p,s,t} + \gamma \cdot \phi_{i,t} + \eta_{i,t} \quad (8)$$

where  $x_{i,t}$  is either the log of total credit granted to firm  $i$  or a measure of output (log value added or net revenues) produced by firm  $i$  during year  $t$  or a measure of (log) input. The  $\psi$  terms are firm and year $\times$ industry $\times$ province fixed effects. The former control for firm-specific unobserved heterogeneity which might affect both financial conditions and production. The latter capture local<sup>23</sup> and sectoral demand and technology shocks, which might create spurious correlation between credit supply and firm dynamics. Results are presented in Table 2. Firms connected with banks expanding their supply of credit show higher growth of credit received, inputs acquired, and output produced than to other firms operating in the same market. The elasticity of credit granted with respect to the firm-level credit supply shock is approximately equal to 1. This allows for simple interpretation of the magnitude of all the main specification of this paper: a one-percentage-point increase in  $\phi_{i,t}$  is the change of credit supply necessary to increase the average credit granted one percent.

The effect of an expansion of credit supply is stronger for value added than for capital accumulation. Net revenues respond almost as much as capital. Labor and intermediate inputs are found to be much less sensitive to credit supply shocks than output and capital are, from both the economic and the statistical point of view. Capital investments are likely to be fully paid up front, while expenditure for materials or labor can sometimes be delayed until some cash flow has been generated from the production. For instance, wages are usually paid at the end of the month. Therefore, it is not surprising that these inputs are less sensitive to changes in a firm's ability to access external finance. To understand whether the effect on inputs is sufficiently large to rationalize the impact on output or, conversely, whether productivity is responding to credit shocks, we need to rely on the elasticities of output to inputs estimated in section 3.2.

## 5 The Effect of Credit Supply on Firm Productivity Growth

Is firm productivity growth affected by the credit supply of its lenders? After identifying firm-level measures of credit supply shocks (section 3.1) and measuring TFP (section 3.2), we now tackle the main research question by estimating the model:

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<sup>23</sup>A province is a local administrative unit, approximately of the size of a US county. CADS reports the province in which each firm is headquartered.

$$\Delta\omega_{i,t} = \psi_i + \psi_{p,s,t} + \gamma \cdot \phi_{i,t} + \eta_{i,t} \quad (9)$$

where:  $\Delta\omega_{i,t}$  is the growth (delta log) of the Hicks-neutral productivity for firm  $i$  between years  $t - 1$  and  $t$  and  $\phi_{i,t}$  is the weighted average of credit supply shocks of  $i$ 's previous-period lenders. The  $\psi$  terms are firm and year  $\times$  industry  $\times$  province fixed effects. The former control for firm-specific unobserved heterogeneity which might affect both financial conditions and production.<sup>24</sup> The latter capture local and sectoral demand and technology shocks, which might create spurious correlation between credit supply and firm dynamics. Results are shown in Table 3. One observation is one firm per year in CADS for 1998-2013, subject to the selection criteria detailed in section 2.4. In each column, we consider productivity growth as obtained from a different production function estimation. The two columns on the left use value added as a measure of output, while productivity in columns 3 and 4 is based on net deflated revenues. Columns 1 and 3 are based on the Cobb-Douglas functional form, while 2 and 4 are based on Trans-Log production functions. The top panel presents results for the whole economy, while the bottom panel focuses on manufacturers. All specifications clearly show that an increase in credit supply boosts productivity growth. A credit supply shock of one percentage point induces an increase in the growth rate or value-added productivity of approximately one-tenth of a percentage point for the whole economy and 0.13 points for manufacturing.<sup>25</sup> The effect on the revenue based measures of productivity is between 0.02 and 0.03 percentage points. The difference between the size of the effect of credit supply on value-added productivity growth and the size of its effect on revenue productivity growth can be partially explained by the fact that, in our sample, the standard deviation of the former is more than three times that of the latter.

The magnitude of the effects is economically large. For instance, the drop in the total growth rate of credit granted between 2007 and 2009 is around 12% in our sample. Over the same period, (mean) value-added productivity growth declined by a bit more than 8% and revenue productivity growth declined by 1%. Therefore, if the drop in credit was fully driven by supply, it would explain between 12% and 30% of the productivity drop over the same period. These figures are likely to be conservative estimates; below we show that the productivity effects of credit shock are persistent

<sup>24</sup>For instance, Malacrino (2016) shows that firms founded by wealthier owners have different dynamics of profitability and growth over their life-cycle.

<sup>25</sup>As shown in Table 2, a 1% increase in supply shock is the change in supply which causes credit granted to increase by 1%.



and that credit supply is particularly valuable during financial turmoil.

Appendix figure A.8 reports the bootstrapped distribution of the estimated effect of credit supply shock on productivity. The production functions are re-estimated for each bootstrap sample. All coefficients are above zero. This finding indicates that the sampling error in estimating productivity dynamics does not distort statistical inferences based on Table 3.

## 5.1 Robustness

This paper argues for a causal interpretation of the estimated relations between credit supply and firm productivity growth. We provide a broad set of robustness exercises to support this claim. Table 4 contains the relative results for the Cobb-Douglas revenue productivity case.<sup>26</sup> Column (1) reports the baseline estimate (as in Table 3). Column (2) adds a set of lagged controls: a polynomial in assets size and the ratios of value added, cash flow, liquidity, and bank debt to assets. The inclusion of such controls has negligible impact on the estimated coefficients.

Estimates of equation (9) face both identification-related threats and measurement threats. This section deals with the latter and with potential problems related to the estimation of the productivity dynamics. Measurement error issues are discussed in appendix C.3. Analogously to the “peer effect” literature (Bramoullé *et al.*, 2009), three main threats may hamper our identification strategy of credit supply shocks based on firm-bank connections: reverse causality, correlated unobservables, and assortative matching. That is,  $\phi_{i,t}$  can be correlated with the error term in equation (9) because (a) connected agents are subject to correlated shocks, (b) lenders might decrease credit supply when expecting their borrowers to experience lower productivity growth, or (c) banks which are expanding their supply of credit are more likely to establish lending relations with firms that are increasing their productivity.<sup>27</sup> The productivity shocks received by sizable borrowers might be the very reason why their lenders contract the supply of credit. That is, if banks have information about the future profitability of some particularly significant borrowers, they might preemptively decrease the supply of credit to all borrowers. We define an “important” borrower as any firm which, at any point between 1997 and 2013, accounts for more than 1% of the credit granted by any of its lenders. We then estimate model (9) excluding such firms. Results are reported in column (3) of Table 4, which shows that the estimated effect of credit supply shocks on productivity growth is unaffected by the

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<sup>26</sup>Results for value-added productivity and revenue translog productivity are in Tables A.11 and A.12. They all show remarkable stability across specifications.

<sup>27</sup>For instance, Bonaccorsi di Patti & Kashyap (2017) document that the banks which recover earlier from distress are the ones which were quicker to cut their credit to risky borrowers. Notice that the additive growth rate model allows for assortative matching in levels.

exclusion of the borrowers that are most likely to lead to reverse causality, thus mitigating this concern.

A further concern is that connected borrowers and lenders might be affected by correlated unobservable shocks. In particular, the output market of the borrower might overlap with the lender’s collection or lending market. For instance, a drop in local house prices might contemporaneously lower consumption and also affect the value of collateral backing lenders’ loans. Since we measure revenue-based productivity, any demand shock might increase markups and be picked up as a change in productivity. To investigate the relevance of correlated unobservables for our results, we compare specifications with two different fixed-effects structures:

$$\Delta\omega_{i,t} = \psi_i + \psi_{p,s,t} + \gamma \cdot \phi_{i,t} + \eta_{i,t}$$

$$\Delta\omega_{i,t} = \psi_i + \psi_p + \psi_s + \psi_t + \gamma \cdot \phi_{i,t} + \eta_{i,t}$$

The first specification includes industry×province×year fixed effects, which aim to control for demand and technology shocks. The second includes only industry, province, and year fixed effects; it therefore allows only for nationwide economic fluctuations. Results are reported in columns (1) and (5) of Table 4. The magnitude of the coefficient is remarkably stable across the two specifications, despite the fact that the inclusion of the finer grid of fixed effects *doubles* the  $R^2$ . This finding reveals that, if any unobservable is affecting both credit supply shocks and productivity, then it must be orthogonal with respect to location or industry. Since credit activity is indeed concentrated at the local (and/or industry) level, this is extremely unlikely to happen. Consequently, we can reasonably conclude that correlated unobservables are not driving our results. A formal econometric treatment of this intuitive argument is provided by [Altonji \*et al.\* \(2005\)](#) and [Oster \(2016\)](#). In appendix C.2, we provide bounding sets for the coefficient of interest, following [Oster \(2016\)](#), and show that they do not contain zero. Therefore, our results are “robust” to the presence of unobservable shocks. Furthermore, column (4) of Table 4 shows that firm fixed effects, while useful to control for firm-level unobservable characteristics, are not essential to our results.

Column (6) of Table 4 adopts an alternative measure of credit supply shocks, which controls for match-level characteristics<sup>28</sup> (see appendix A.1 for details). The estimated effect of credit supply on

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<sup>28</sup>Namely, the size of the loan relative to the borrower’s total credit received, size of the loan relative to the lender’s

productivity growth is similar to that in the baseline specification of column (1), providing no evidence that assortative matching explains our results. Finally, section 6 exploits a natural experiment to confirm that credit supply affects productivity; this, together with relative placebo tests, should eliminate residual concerns.

The bank-level credit supply shocks are computed using information on all borrowers. Therefore, if firm  $i$  has a lending relation with bank  $b$ , then its credit supply is estimated from a linear regression including observations relative to the amount of credit granted by  $b$  to  $i$  (see section 3.1). This could generate problems in small samples. Therefore, we estimate an alternative set of bank-level credit supply shocks using a “split sample” procedure.<sup>29</sup> Column (7) presents estimates of the baseline specification using the “split sample” credit supply shock as an instrument. The similarity between estimates in columns (1) and (7) confirm that, since we rely on the universe of credit relations, this (potential) finite-sample bias is not a concern.

Estimation of production function parameters is a difficult exercise involving several (strong) assumptions, such as the absence of measurement error on inputs and a Markovian structure for the productivity dynamics. We perform several exercises to show that the specific modeling choices of section 3.2 do not affect the estimated effect of credit supply on productivity growth either qualitatively or in terms of its magnitude. First, we re-estimate both the production function and equation (9), using a finer four-digit industry classification (the baseline uses two-digit classification). Results are reported in column (8) of Table 4, which mitigates the concern that heterogeneity in the shape of the production function is a main driver of the baseline specification. Second, we re-estimate the production function by controlling for endogenous exit as in [Olley & Pakes \(1996\)](#). Column (9) of Table 4 shows that the magnitude of the relation between credit supply shocks and productivity is unchanged.<sup>30</sup> Furthermore, we compare our results to traditional estimation techniques. Column (10) of Table 4 reports results from the production function estimated with the cost-share procedure ([Foster et al. , 2017](#)).<sup>31</sup> Results are in the ballpark of the baseline estimation.

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total credit granted, interest rate, length of the lending relationship, type of credit instrument used, and the occurrence of past non-performing loans.

<sup>29</sup>That is, we divide all firms into subsamples A and B. For each bank, we estimate two credit supply shocks,  $\phi_{b,t}^A$  and  $\phi_{b,t}^B$ , using data about credit given to firms belonging to only one subsample at time. Then, we compute firm-level idiosyncratic shocks as the weighted average of the bank-level credit supply shocks estimated with data on firms of the other subsample. For instance, if firm  $i$  belongs to subsample A, we estimate its credit supply shocks as  $\phi_{i,t} = \sum_b w_{i,b,t-1} \cdot \phi_{b,t}^B$  where  $w_{i,b,t-1}$  is the share of credit to firm  $i$  granted by bank  $b$  in the previous period.

<sup>30</sup>An alternative model of production with heterogeneous credit constraints is presented in section B.2, together with relative results, which do not differ significantly from the baseline estimation.

<sup>31</sup>Under the Cobb-Douglas functional form, the ratio of the expenditures on any flexible input to income is equal to the elasticity of the output with respect to this input. Therefore, assuming labor and intermediate inputs are fully flexible, we estimate their sectoral elasticities as the median share of expenditure on each input over the total revenue (or value added). The elasticity of capital is given by 1 (constant returns to scale) minus the elasticities of labor and

An alternative approach is to refrain from estimating the production function and, instead, study how the estimated effect of credit supply shocks on productivity varies as a function of the unknown parameters of the production function. The simplest production function is a Cobb-Douglas in value added:

$$va_{i,t} = \omega_{i,t} + \rho \cdot (\beta_k \cdot k_{i,t} + (1 - \beta_k) \cdot l_{i,t})$$

where  $\rho$  disciplines the returns to scale and  $\beta_k$  is the (relative) elasticity of value added to capital. Then, given a pair  $(\tilde{\rho}, \tilde{\beta}_k)$ , we can back out productivity as

$$\omega_{i,t}(\tilde{\rho}, \tilde{\beta}_k) = va_{i,t} - \tilde{\rho} \cdot (\tilde{\beta}_k \cdot k_{i,t} + (1 - \tilde{\beta}_k) \cdot l_{i,t})$$

and estimate  $\gamma(\tilde{\rho}, \tilde{\beta}_k)$  as the coefficient of

$$\Delta\omega_{i,t}(\tilde{\rho}, \tilde{\beta}_k) = \psi_i + \psi_{p,s,t} + \gamma(\tilde{\rho}, \tilde{\beta}_k) \cdot \phi_{i,t} + \eta_{i,t} \tag{10}$$

We let  $\rho$  vary from 0.3 to 2 and  $\beta_k$  from 0.01 to 0.9, so that our grid encompasses any plausible values of the return to scale and the elasticity of value added to capital. Results are presented in graphical form in figure 5, showing that we find a positive (and statistically significant) effect of credit supply shocks on value-added productivity growth for any point on the grid. Moreover, while higher values of the parameters tend to decrease the point estimates,  $\gamma(\tilde{\rho}, \tilde{\beta}_k)$  stays between 0.07 and 0.1 within the whole support.

The collection of evidence reported in this section clarifies that any misspecification of the pro-  


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intermediates. Foster *et al.* (2017) describe the theoretical and empirical differences between the cost-share approach and the control function. We divide intermediate inputs into expenditure for services and expenditure for materials in order to show that merging them together does not drive the baseline results of the paper. Doing so, we lose some observations, since not all income statements report expenditure on the two items separately.

duction function estimation, although it might bias the point-estimate of the effect of credit supply on productivity, is unlikely to change its magnitude significantly.

## 5.2 Heterogeneity

Are all firms equally affected by credit supply shocks? A firm’s size might be a good predictor of its ability to find alternative sources of credit in case current lenders dry up. Furthermore, larger firms are less likely to be credit-constrained in the first place. Therefore, for each year, we compute an indicator for whether or not a firm is in the top quartile of the size distribution in terms of asset value or number of employees. Then, we estimate the equation:

$$\Delta\omega_{i,t} = \psi_i + \psi_{s,p,t} + (\gamma + \gamma_{big} \cdot Big_{i,t-1}) \cdot \phi_{i,t} + \psi_{Big} \cdot Big_{i,t-1} + \eta_{i,t}$$

Results are reported in columns (1) and (2) of Table 5, which refer to Cobb-Douglas revenue productivity. The parameter  $\gamma_{big}$  is estimated to be negative, indicating that large firms are less affected by credit supply shocks. The difference between the two groups is much larger and statistically significant in manufacturing.

Furthermore, we are interested in understanding whether having a larger number of lenders might help firms find sources of finance in case of negative credit supply shocks. Therefore, we estimate the model by allowing the coefficient to be different for firms in the bottom quartile for number of lending relations during the previous period.<sup>32</sup> Results in column (3) document that borrowers with fewer lenders are much more affected by credit supply shocks.

An important dimension of the relevance of credit supply shocks is firms’ reliance on external funds. We classify industries as above and below the median according to the mean leverage (debt over assets) in the sample. Column (4) of Table 5 shows that the effect of credit supply shocks on revenue productivity is stronger in sectors with high leverage. Perhaps surprisingly, we do not find any significant pattern when analyzing heterogeneity according to sectoral cash flow over assets (see column (5)).

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<sup>32</sup>A few seminar participants suggested differentiating the effect of credit shocks between firms with one and with multiple lending relationships. Unfortunately, less than 5% of the observations in CADS have only one lender, so the relative coefficient would not be reliably estimated.

### 5.3 Persistence

The effect of credit supply on productivity is persistent. We define the innovation to the credit supply as  $\zeta_{i,t}^\phi := \phi_{i,t} - E[\phi_{i,t}|\phi_{t-1}]$ . Then, we estimate the model

$$\omega_{i,t} = \psi_i + \psi_{p,s,t} + \sum_{\tau=-T}^T \gamma_\tau \cdot \zeta_{i,t}^\phi + \eta_{i,t} \quad (11)$$

We choose  $T = 3$ , since our empirical strategy is not fit to estimate the regression at a longer horizon.<sup>33</sup> Figure 4 graphically displays the coefficients,  $\gamma_\tau$ , for firms active in manufacturing (bottom panel) and all industries (top panel). They document that the peak in productivity is experienced one year after the shock and that the effect remains positive and significant for at least four years. This finding underlines that a temporary credit contraction can have persistent effects on productivity. It also rules out the potential concern that the effect we measure on revenue productivity is short-lived and due to factor hoarding caused by adjustment costs of labor and capital.

We do not find any statistically significant pre-trend. Our main results rely on pre-existing lending relations being orthogonal with respect to non-financial productivity shocks. Therefore, the absence of a pre-trend supports the claim that credit supply shocks have a causal effect on productivity.

### 5.4 Concavity of the credit-productivity relationship

The main goal of this paper is to measure and explain the productivity effects of changes in the quantity of credit supplied, focusing on its first moment: is *more* credit bad or good? This section, instead, investigates the shape of the relation between productivity and credit supply shocks, in order to understand whether higher moments of the distribution of credit supply shocks might have an impact on average firm productivity.

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<sup>33</sup> The within-firm estimator, while allowing us to control for firm unobserved heterogeneity, creates a mechanical negative correlation between observation means at different lags. In fact, regression of firm productivity on past productivity yields a coefficient between .9 and .98 if no fixed effects are included and between .3 and .4 if the standard set of fixed effects is included. Therefore, a shock to productivity of magnitude  $1 \cdot m$ , is expected to show up as a change in productivity of only  $0.03 \cdot m$  to  $0.06 \cdot m$  after 3 years.

We divide  $\phi_{i,t}$  into quintiles  $q = 1, 2, 3, 4, 5$  and estimate:

$$\Delta\omega_{i,t} = \psi_i + \psi_{p,s,t} + \sum_{q=1, q \neq 3}^5 \gamma^q \cdot \mathbf{1}(\phi_{i,t} \in q) + \eta_{i,t}$$

where  $\mathbf{1}(\phi_{i,t} \in q)$  is an indicator function taking value 1 iff the credit supply shock of firm  $i$  in year  $t$  belongs to the  $q$ th quintile of its distribution; the third (or median) quintile  $q = 3$  is the omitted category with  $\gamma^3 = 0$ . Results are shown in graphical form in figure 6. The relation between credit supply and revenue productivity seems to be concave. That is, firms connected with banks with a relatively low supply of credit experience lower revenue productivity growth than their competitors; firms connected to banks with a particularly strong increase in credit do not grow at a particularly high rate. It is important not to be connected with banks experiencing bad credit supply shocks, but it is not useful to be connected with banks increasing their supply of credit particularly quickly. To strengthen this intuition, we re-estimate equation (11), which is used to study the persistence of credit supply shocks, by differentiating between positive and negative shocks. Figure 7 presents the results in graphical form. The coefficients relative to negative credit supply shocks are shown with negative values. The effect of credit supply shocks on productivity is driven by firms connected with banks experiencing relatively negative credit supply dynamics. Additionally, we argue in section 6 that credit supply shocks are particularly important when credit dries up.

These empirical findings imply that an increase in credit supply cannot undo the harm of a negative shock of the same size. Therefore, it is not only the quantity of credit that matters, but also the stability of its provision. This analogously suggests that a credit crunch followed (or preceded) by a credit expansion of the same magnitude leads to a net loss in average firm productivity. We conclude that the volatility of the banking sector's supply can be detrimental to firm productivity.

## 6 The Interbank Market Collapse as a Natural Experiment

The credit supply shock derived in section 3.1 has the value of being general, in that it can be attributed to all firms (both multiple- and single-borrowers) and measured in any year for which there is bank-firm data on credit granted. This feature is exploited in section 7. Furthermore, the panel variation of  $\phi_{i,t}$  is essential for production function estimation (see section 3.2). However, since its construction relies on firm-bank connections, estimates of equation (9) might suffer from

the identification problems highlighted in section 5.1. Although we have already discussed several robustness exercises to mitigate such concerns, here we propose an alternative strategy to strengthen the robustness of our results. We use the 2007-2008 market collapse of the interbank market as a specific “natural experiment” in which credit supply shifts were arguably exogenous with respect to firm observed and unobserved characteristics.<sup>34</sup> In addition, such variation came unexpectedly both to lenders and to borrowers, thus overcoming the problem of assortative matching.

The interbank market is a critical source of funding for banks: it allows them to readily fill liquidity needs of different maturities through secured and unsecured contracts. Total gross interbank funding accounted for over 13% of total assets of Italian banks at the end of 2006. Market transactions began shrinking in July 2007, when fears about the spread of toxic assets in banks’ balance sheets made the evaluation of counterparty risk extremely difficult (Brunnermeier, 2009); the situation worsened further after Lehman’s default in September 2008. As a consequence, total transactions among banks fell significantly. In Italy, in particular, they plummeted from €24bn. in 2006 to €4.8bn. at the end of 2009. At the same time, the cost of raising funds in the interbank market rose sharply: the Euribor-Eurepo spread, which was practically zero until August 2007, reached over 50 basis points for all maturities in the subsequent year. It then increased by five times after the Lehman crisis and remained well above 20 basis points in the following years. Two recent papers have exploited the collapse of the interbank market as a source of exogenous shock to credit supply. Iyer *et al.* (2013) used Spanish data to show that bank *pre-crisis* exposure to the interbank shock, as measured by the ratio of interbank liabilities to assets, was a significant predictor of a drop in credit granted during the crisis. Cingano *et al.* (2016) focus on CADS data for Italy to show that this drop had a significant negative effect on firms’ capital accumulation. These researchers reported results of several empirical tests showing that banks’ pre-crisis exposure was not correlated with their borrowers’ characteristics, such as investment opportunities and firm growth potential, thus making this variable particularly suitable to instrument the impact of credit supply on firms’ outcomes. We focus on the period 2007-2009, when credit dried up the most. Subsequently, ECB interventions partially offset the impact of the interbank market shock. Our measure of firm exposure to the credit supply tightening is the average 2006 interbank exposure of Italian banks at the firm level, using firms’ specific credit shares in 2006 as weights. Because firm exposure is time-invariant, we use cross-sectional variation. We include observations over a three-year window. Formally, for each firm  $i$  active in industry  $s$  and province  $p$  over the years  $t \in [2007, 2009]$ , we estimate the equation:

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<sup>34</sup>This is not the first paper to rely on natural experiments to identify idiosyncratic credit supply shocks; see, for instance, Khwaja & Mian (2008), Chodorow-Reich (2013), and Paravisini *et al.* (2014).



$$\Delta\omega_{i,t} = \psi_{p,s,t} + \gamma \cdot INTBK_{i,2006} + \eta_{i,t} \quad (12)$$

where  $\omega_{i,t}$  is firm idiosyncratic productivity,  $INTBK_{i,2006}$  is the pre-crisis reliance on the interbank market, and  $\psi$  is a set of province $\times$ industry $\times$ year fixed effects. Results are shown in Table 6. Firms whose lenders were more reliant on the interbank market in 2006 had significantly lower revenue and value-added productivity growth during the credit crunch. This strengthens the causal interpretation of the relations between credit supply and productivity growth documented in section 5. A 1% increase in average bank dependence on the interbank market results in an approximately .05% decrease in average value-added productivity growth and an approximately .02% decrease in revenue productivity growth. Consequently, the same interbank shock which decreases credit growth by 1% also decreases value-added productivity of 0.25% and revenue productivity by one-tenth of a percent for the whole sample. These effects are between two and five times larger than the baseline estimate from Table 3, suggesting that accessing a reliable source of credit supply is particularly important during financial turmoil.

## 6.1 Placebo and robustness tests

Estimation of (12) provides evidence that firms hit harder by the credit crunch decrease their relative productivity. What if banks relying more heavily on the interbank market were just matched to worst borrowers? To remove this concern, we run equation (12) including only years *before* the freeze of the interbank market; that is,  $t \in [2004, 2006]$ . Results, shown in columns (1)—(4) of Table 7, show that firms more exposed to the freeze of the interbank market did not have statistically different growth rates of productivity before the credit crunch. Additional results show that firms more exposed to the interbank shock were not more sensitive to business-cycle fluctuation before 2007. Details are in appendix D. We implement an additional placebo test. That is, we investigate the effect of a hypothetical freeze of the interbank market in 2003. For  $t \in [2003, 2005]$  we estimate the model:

$$\Delta\omega_{i,t} = \psi_{p,s,t} + \gamma \cdot INTBK_{i,2002} + \eta_{i,t}$$

Columns (5)—(8) of Table 7 show that the placebo collapse is not a significant predictor of firms'

subsequent productivity growth.

The collection of evidence presented in this section should eliminate the concern that the relation between credit supply and productivity documented in section 5 is driven by correlated unobservables, reverse causality, or assortative matching.

## 7 Beyond Measurement: Channels

How does credit supply improve productivity? In this section, we investigate the relations between the credit supply shocks and several productivity-enhancing activities. As described in section 2, INVIND provides information about R&D investment, export, IT-adoption, and self-reported “obstacles to innovation” for a sample of Italian companies in services and manufacturing. Because both questions and respondents vary between waves, each specification of this section relies on a different sample. Furthermore, the sample size is much smaller than in the previous sections, limiting our ability to use our preferred specification.

In section 5.3, we show that credit supply shocks affect productivity immediately. We detect additional productivity growth for at least two years and higher productivity for at least four years. Unfortunately, our empirical framework is not fit to investigate the effect at a longer horizon. Some of the productivity-enhancing strategies studied in this section, such as IT-adoption or better management practices, are likely to affect productivity as soon as they are implemented. Others, such as R&D, might take a few years to produce substantial improvement. Therefore, this section does not only explore the potential mechanisms behind the effect we measure in section 5, but also suggests that credit availability might lead to additional productivity gains in the long run.

### 7.1 IT-intensity of capital stock

The speed of adoption of IT technologies caused large differences in productivity between US and European companies (Bloom *et al.* , 2012). According to Pellegrino & Zingales (2014), failure to take full advantage of the IT revolution is one of the main drivers of Italy’s low productivity growth. Data on personal computers used is available from INVIND for 1999-2001. Purchases of PCs are accounted as investments. Therefore, they enter the computation of capital stock. Slacker credit constraints might allow firms to stay closer to the technological frontier. By making more technological investments, unconstrained firms might have a “better” capital stock. Since researchers do not have detailed information on “quality” or “closeness to the frontier” of inputs, this quality is picked

up by the productivity residual. To test this hypothesis, we measure the “IT-intensity” of firm capital stock as (log) number of PCs per 1,000 euros of capital.

$$IT_{i,t} = \psi_i + \psi_{s,p,t} + \gamma \cdot \phi_{i,t} + \eta_{i,t}$$

Results are presented in column (1) of Table 8. Firms are more likely to increase the IT-intensity of their capital stock when they receive a positive credit supply shock. This finding suggests that financial frictions lower the quality of capital inputs used in production.

## 7.2 Innovation and exporting

Patenting activities have been extensively used as a proxy for firm-level knowledge creation (see [Bernstein \(2015\)](#) and [Kogan \*et al.\* \(2017\)](#) for recent examples). We obtain information for patent applications for a large fraction of Italian companies from PatStat, as described in section 2. In our sample, patent applications became much less common during (and after) the credit crunch. The share of firms applying for at least one patent was approximately 2% between 2002 and 2007. It declined to 1.5% in 2009 and went up to a bit more than 1.6% in the following two years. We observe approximately 5 patent applications per 100 firms per year before the Great Recession, but only around 3.4 in 2009. This pattern, of course, could be driven by lower demand and/or greater uncertainty. To investigate whether availability of credit has a causal impact on patent applications, we estimate the models:

$$PatentApp_{i,t} = \psi_i + \psi_{p,s,t} + \gamma \cdot \phi_{i,t-1} + \eta_{i,t}$$

$$GrowthPatentApp_{i,t} = \psi_i + \psi_{p,s,t} + \gamma \cdot \phi_{i,t-1} + \eta_{i,t}$$

Following the literature on R&D and patents, we impose a lag between the credit shock and patent applications. The growth rate of patent applications is defined as  $GrowthPatentApp_{i,t} =$

$2 \frac{PatentApp_{i,t} - PatentApp_{i,t-2}}{PatentApp_{i,t} + PatentApp_{i,t-2}}$ .<sup>35</sup> Results are reported in columns (2) and (3) of Table 8. Italian firms patent more when they have easier access to bank credit. Appendix D.3 uses the collapse of the interbank market to provide additional evidence on the causal effect of the tightening of credit constraints on innovation.

R&D can increase firm productivity by improving both product quality and process efficiency. Similarly, export can have beneficial effects through two channels: it allows firms to access markets with higher margins and it can improve firm know-how through so called “learning-by-exporting.” The sensitivity of international trade to financial frictions has been studied by several authors (Manova, 2012). We use the INVIND survey to identify firms that export and that have positive R&D expenditures. We focus on the extensive margin and estimate two linear probability models:

$$Pr(R\&D_{i,t} = 1) = \psi_i + \psi_t + \gamma \cdot \phi_{i,t} + \eta_{i,t}$$

$$Pr(Export_{i,t} = 1) = \psi_i + \psi_t + \gamma \cdot \phi_{i,t} + \eta_{i,t}$$

where  $R\&D_{i,t}$  and  $Export_{i,t}$  are dummy variables indicating whether firm  $i$  engages in R&D or exporting in year  $t$ . Results, presented in columns (4) and (5) of Table 8, show that there is a positive and statistically significant relation between credit supply and the propensity for these productivity-enhancing activities. This indicates that companies are more likely to start (and less likely to stop) conducting R&D and exporting when they have easier access to external finance.

Innovative effort is much broader than just formal R&D or IT-adoption. The 2011 survey wave investigates the main constraints to innovative effort. One question asks how important, on a four-item scale, the firm’s difficulties in collecting external funds were in limiting innovation (in 2010). We build the variable  $FinCon_{i,2010}$ , equal to 1 if and only if difficulty in getting external funds is reported to be “somehow important” or “very important” as an obstacle to innovation. Then, we estimate the linear probability model:

$$Pr(FinCon_{i,2010} = 1) = \psi_{s,p} + \gamma \cdot \phi_{i,2010} + \eta_i$$

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<sup>35</sup> In the rest of the paper, we measure growth rates of credit, inputs, output, and TFP using a delta-log approximation. However, patent applications are rare, so we rely on the well-known formula by Davis *et al.* (1996), which can be calculated if patent applications are positive either in year  $t$  or  $t-2$  (or both). We use a two-periods lag rather than one because we have a lagged right-hand-side variable.

Results are presented in column (6) of Table 8, which documents that firms receiving positive credit supply shocks are less likely to consider external funds as a substantial obstacle to innovation. Since the question was asked for only one year of the survey, we cannot use panel variation. Nonetheless, this exercise is an indirect—yet insightful—test of the hypothesis that financial frictions dampen firms’ innovative efforts.

### 7.3 Management practices

Management matters for firm performance, as shown by Bloom *et al.* (2013) for India and by Giorcelli (2016) for Italy. We use credit supply shocks to investigate whether firms improve their management when facing slacker financial constraints. The direction of the relation is not obvious. Scarcity of resources might push firms to improve their internal organization. Conversely, improvement in management practices might require stable financial resources; for instance, to hire professional consulting services or to restructure a production facility. Bhattacharya *et al.* (2013) propose a model in which frictions distort optimal investment in managerial skills.

We obtain firm-level data on management practices from the World Management Survey (WMS).<sup>36</sup> As can be read from the website, WMS “developed an in-depth survey methodology and constructed a robust measure of management practices in order to investigate and explain differences in management practices across firms and countries in different sectors.” Information on data construction can be found in Bloom & Van Reenen (2007). They state that the “practice evaluation tool defines and scores from one (worst practice) to five (best practice) across eighteen key management practices used by industrial firm.” Merging WMS data on Italian companies by name, we obtain a sample of 183 observations. Because we have only one or two survey waves for each firm, we estimate the cross-sectional model:

$$MS_{i,t} = \psi + \gamma \cdot \phi_{i,t} + \eta_{i,t}$$

where  $MS_{i,t}$  is the overall management score for firm  $i$  surveyed in year  $t$ . Results are presented in column (7) of Table 8, which indicates that an increase in credit supply stimulates the adoption

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<sup>36</sup> See <http://worldmanagementsurvey.org/>. We are grateful for the data provided.

of superior management practices. While the small sample size might cast doubt on the robustness of this result, the relation between credit supply shock and management is largely unaffected by the inclusion of a large set of firm-level controls.

## 7.4 Managerial inattention

We propose a novel theory to explain why firms subject to negative credit shocks might decrease their productivity. Dealing with investors and creditors takes a substantial share of executive time. [Bandiera et al. \(2011\)](#) study the use of time by 94 CEOs of top-600 Italian companies. They document that finance is the topic on which the CEO spends the most time talking with others in the firm. Furthermore, of the outsiders with whom CEOs spend the most time, investors and bankers are, respectively, third and fifth. If this is true for Italian top-600 companies, which are all likely to have a professional CFO and other finance-related personnel, the time and effort required to establish and maintain relations with lenders might be even more demanding for the managers and entrepreneurs of smaller private companies which make up the bulk of our sample. Since their time is limited (as in a “temporal” limited span of control<sup>37</sup>), then the more difficult (or time-consuming) it is to find external funds, the less they can work on improving their core business. Entrepreneurs connected to lenders who contract their credit supply might need to spend more time and energy to establish new lending relations. Therefore, they might exert less effort in improving their firm’s productivity. As a colorful piece of anecdotal evidence to support this theory, the aunt of one of the authors was managing the family business during the credit crunch. When asked about the firm’s performance, she used to reply, *“I barely have time to go to the factory, I spend most of my mornings at banks trying to get some money.”* As an indirect test of this mechanism, appendix [A.2](#) and the relative results in [Table A.1](#) show that firms receiving positive credit supply shocks are less likely to try to establish new lending relations. A more direct and complete investigation of this hypothesis is left to future research.

## 8 Conclusion

To grow and thrive, firms need reliable access to external funding. In particular, this paper carefully documents that credit supply is an important determinant of improvement to a firm’s performance, both in the short run and the long run.

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<sup>37</sup>See [Akçigit et al. \(2016\)](#) for an example of the harm caused by the lack of managerial delegation in developing countries.

We therefore study the impact of banks' credit supply on production for a large sample of Italian corporations. We exploit the universe of bank-firm credit relationship over the period 1997-2013 to estimate an additive growth rate model and we separate demand from supply shocks using firm-time and bank-time fixed effects. We improve on the literature by considering two important extensions to this framework. Then, we use the estimated bank-level supply shocks and the stickiness of lending relationship to build a measure of firm-specific shocks to credit supply. We document that firms connected to banks which are expanding their supply of credit acquire more inputs and produce higher output than their competitors. We show that the effect on output is stronger than the effect on inputs, suggesting that productivity is affected by credit availability.

We build a model of production with heterogeneous credit constraints in order to estimate an industry-specific production function and isolate firm idiosyncratic productivity dynamics. Then, we show that credit supply boosts productivity growth and that these effects are sizable, persistent, and robust. Moreover, they are stronger for smaller firms and for companies in sectors relying heavily on bank credit. Furthermore, we exploit the 2007-08 freeze of the interbank market as a natural experiment to support the causal interpretation of our estimates and show that they are not driven by the assortative matching of borrowers and lenders or by reverse causality. Our results imply that financial turmoil can have a persistent effect on aggregate output because it depresses firms' TFP in the short and long run. Furthermore, our findings suggest that financial frictions are harmful beyond their detrimental effects on allocative efficiency.

We show that a negative credit supply shock produces stronger effects than a positive one of the same magnitude. This finding implies that it is not only the quantity of credit supply that matters, but also its stability.

Finally, we show that several productivity-enhancing activities, such as adoption of IT, sound management practices, export orientation, and innovation, are stimulated by credit availability. We also conjecture that a reduction of credit supply might force borrowers (notably, managers and entrepreneurs) to consume time and energy in order to establish connections with additional lenders. Consequently, they might exert less effort in improving business performance. We document that firms' attempts to create new lending relationships are indeed more frequent when they experience negative credit shocks.

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## Figures

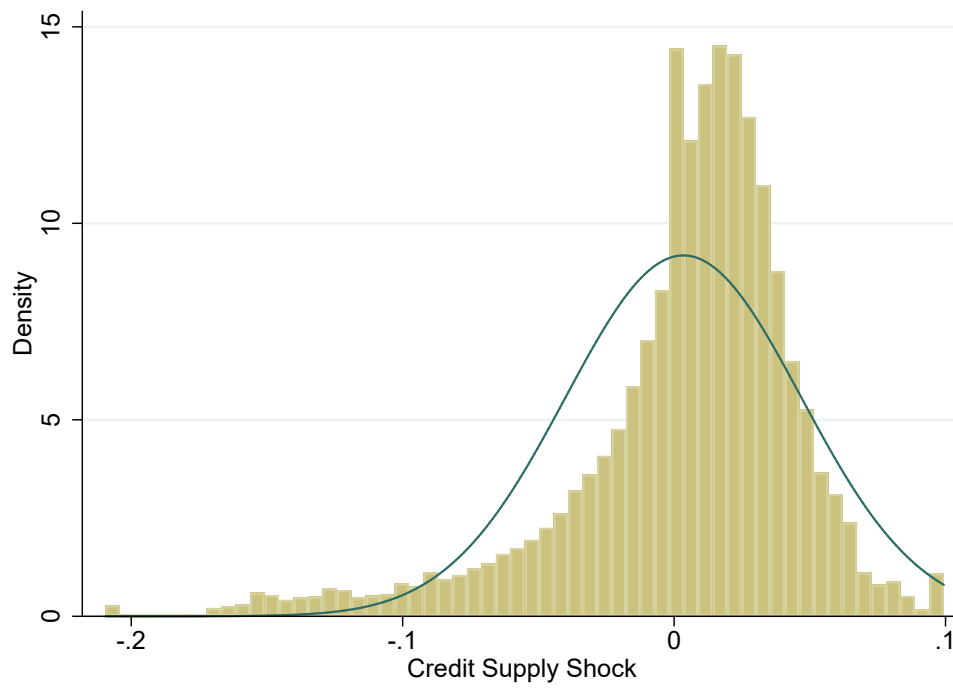


Figure 2: Histogram of credit supply shock. Normal distribution is superimposed. See section [3.1](#) for details.

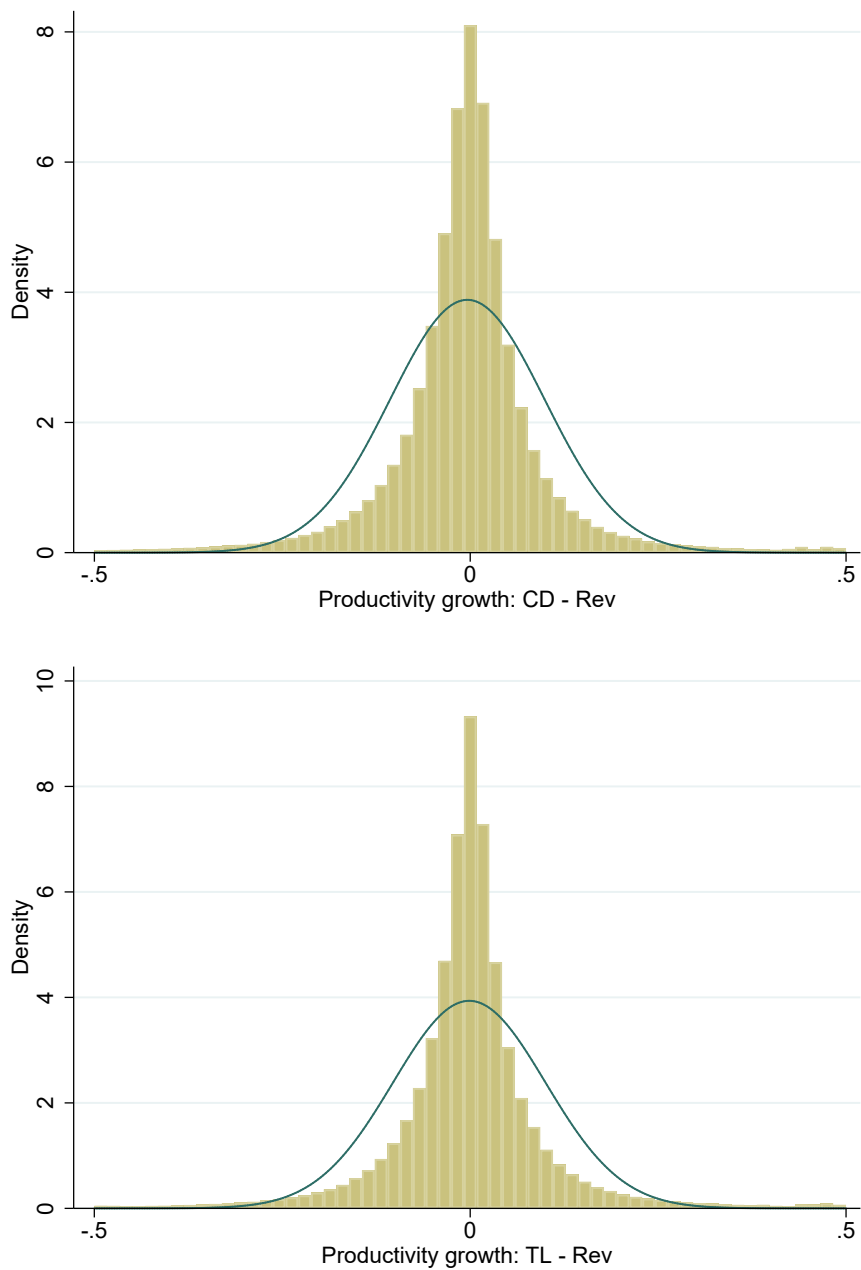


Figure 3: Histogram of productivity growth. Productivity is estimated as a residual from (log) revenues production function. Cobb-Douglas (top panel) or Trans-Log (bottom panel) functional form is assumed. Normal distribution is superimposed. See section 3.2 for details.

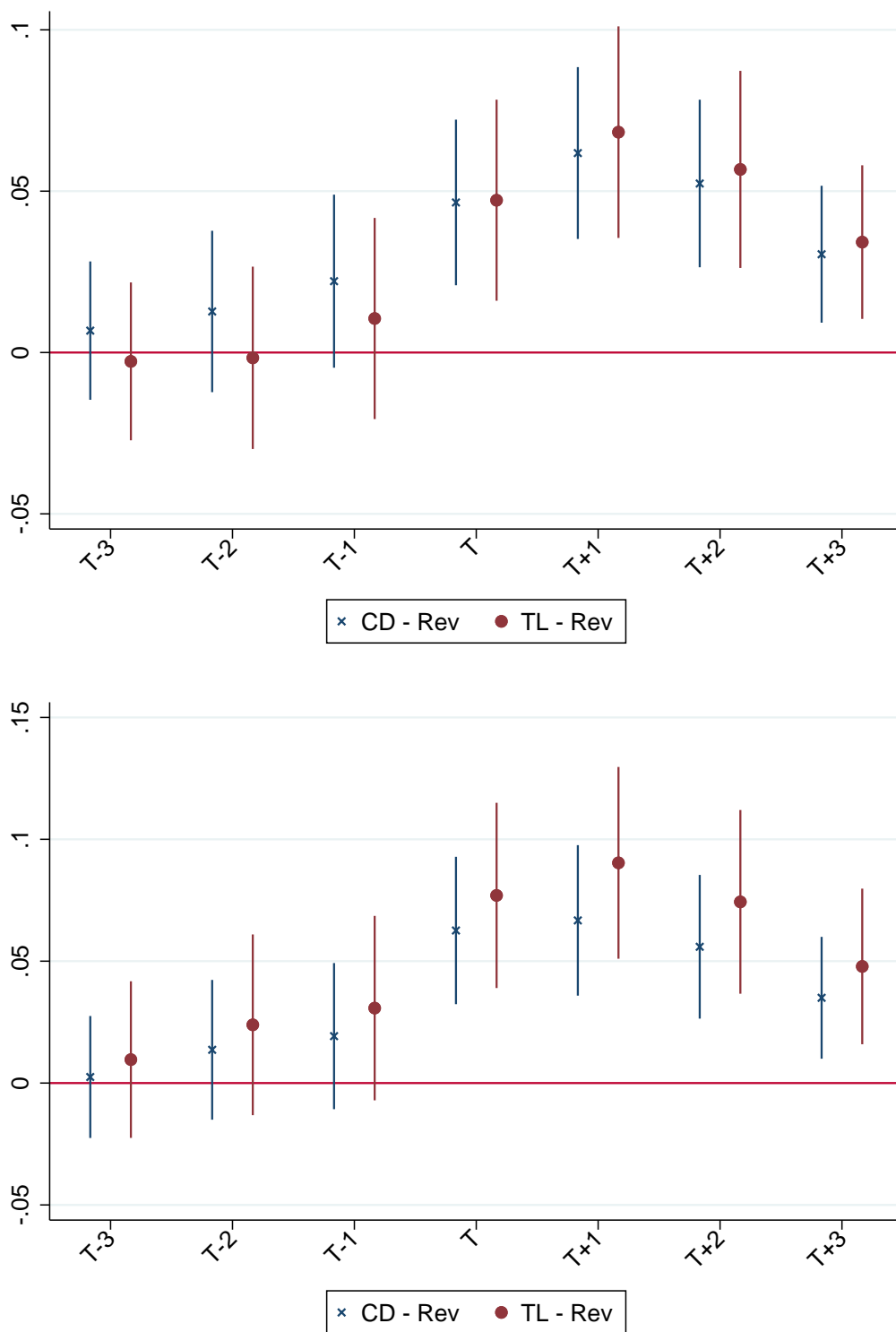


Figure 4: Productivity before and after an unexpected credit supply shock, see section 5.3 for details on the estimated equation. Top panel refers to all industries, while bottom panel refers to manufacturers. 99% confidence intervals are displayed. Productivity is estimated as residual from a (log) revenue production function. Functional form is either Cobb Douglas or Trans-Log. Details on productivity estimation are in section 3.2.



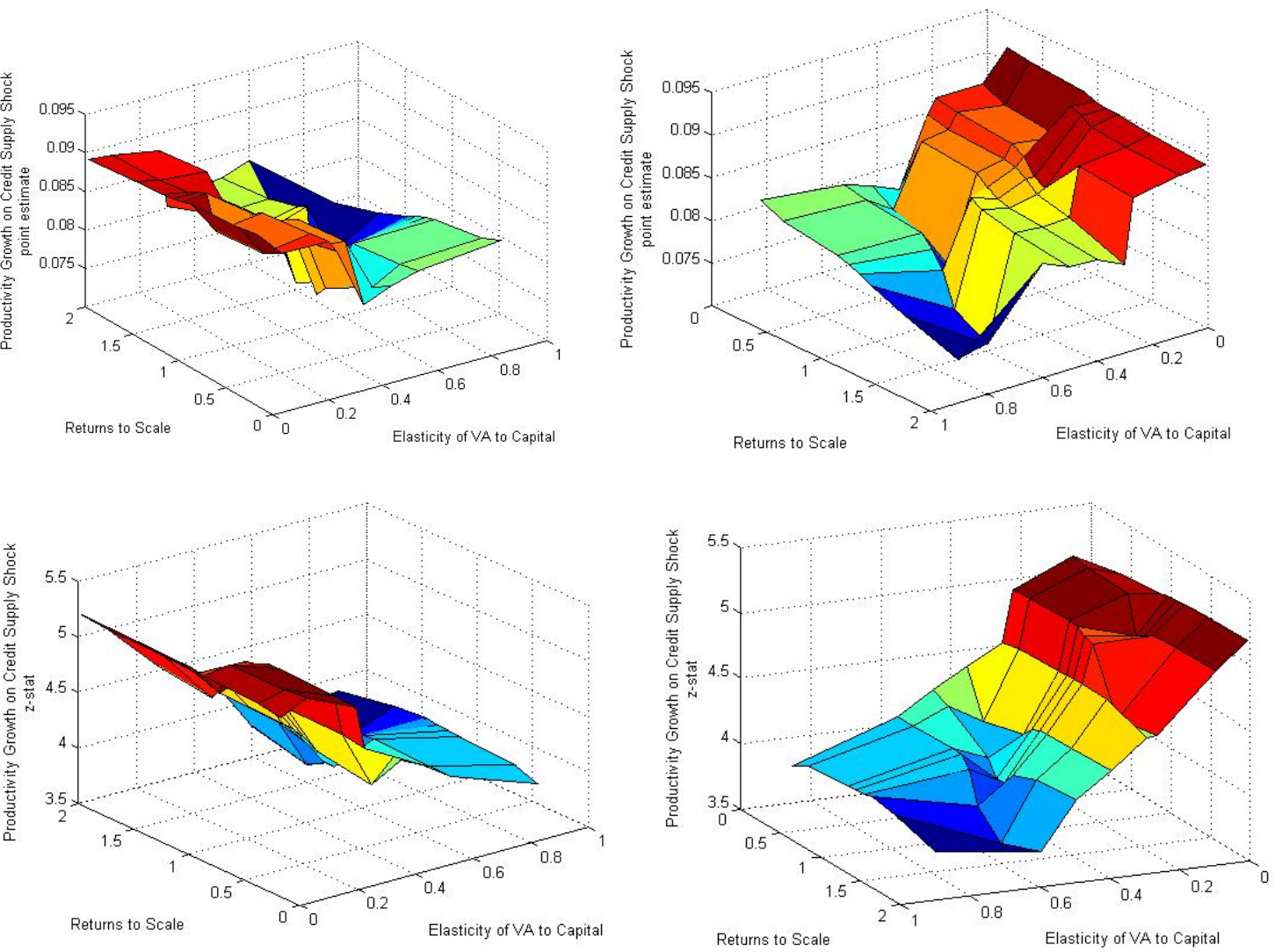


Figure 5: Credit Supply Shock and VA Productivity for different parameters of Cobb Douglas production function. The z-axis reports the estimated parameter  $\gamma_{\rho, \beta_k}$  (top figures) or relative z-stats (bottom figures), from regression  $\Delta\omega_{i,t}(\rho, \beta_k) = \psi_i + \psi_{s,t,p} + \gamma \cdot \phi_{i,t} + \eta_{i,t}$ .  $\rho$  is the return to scale while  $\beta_k$  is the relative elasticity of value added to capital. One observation is one firm for one year between 1998 and 2013 (unbalanced panel). The RHS variable  $\phi_{i,t}$  represents idiosyncratic shock to firm credit supply, and its construction is detailed in section 3.1. The LHS variable is the first difference of Hicks-neutral productivity residual:  $\Delta\omega_{i,t}(\rho, \beta_k) = \Delta va_{i,t} - \Delta\rho(\beta_k k_{i,t} + (1 - \beta_k)l_{i,t})$  where  $va$  is log of value added,  $k$  is the log of capital stock and  $l$  is log of labor (wagebill). Left and right panels show same patterns from two different angles.

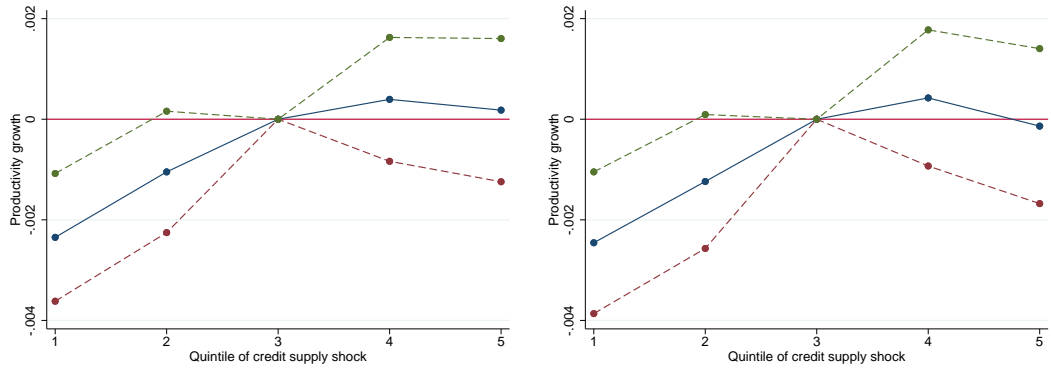


Figure 6: Growth rate of productivity in manufacturing per quintile of credit supply shock. The third quintile (which includes the median credit supply shocks) is normalized to zero. Productivity is residual from a Cobb-Douglas revenues production function (left side) or Trans-Logs revenues production function (right side).

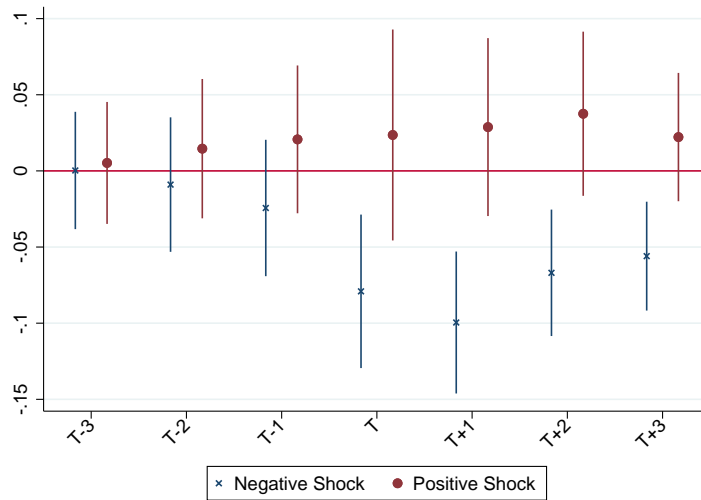


Figure 7: Revenue productivity before and after a credit supply shock - negative vs positive shocks

# Tables

Table 1: Descriptive Statistics - main firm-level variables

VARIABLE	Mean	Std. Dev.	Median	N	Mean	Std. Dev.	Median	N
	All Industries				Manufacturing			
Value Added	5,312	33,819	1,641	656,960	5,409	21,699	1,943	347,990
Net Revenues	27,073	156,638	8,813	656,960	25,351	164,054	8,209	347,990
Wagebill	3,377	19693.9	1,062	656,960	3,466	13452.4	1,299	347,990
Capital Stock	8,636	153,346	1,545	656,960	7,111	40,357	2,058	347,990
Intermediate Inputs	21,888	137,390	6,873	656,960	20,057	150,610	6,119	347,990
Credit Granted	7,924	3,6445	2,737	650,664	8,039	29,760	3,013	345,700
Employees	80	472	28	656,960	79	269	35	347,990

*Notes:* One observation is one firm for one year, between 1998 and 2013 (unbalanced panel). All variables (except for number of employees) are expressed as thousands of 2010 euros using sector-level deflators from national accounts.

*Source:* CADS and Credit Register.

Table 2: Credit, Inputs and Outputs response to Credit Supply Shocks

VARIABLES ( in delta Log)	Credit Received (1)	Value Added (2)	Net Revenues (3)	Capital Stock (4)	Wagebill (5)	Number of Employees (6)	Intermediate Inputs (7)
All Industries							
$\phi_{i,t}$	0.949*** (0.0196)	0.123*** (0.0162)	0.0474*** (0.0109)	0.0619*** (0.0128)	0.0154* (0.00926)	0.00608 (0.00889)	0.0220* (0.0114)
Observations	609,195	656,960	656,960	656,960	656,960	656,960	656,960
$R^2$	0.239	0.224	0.302	0.259	0.324	0.319	0.302
Manufacturing							
$\phi_{i,t}$	0.966*** (0.0253)	0.134*** (0.0201)	0.0436*** (0.0143)	0.0610*** (0.0169)	0.00388 (0.0116)	-0.00892 (0.0108)	0.00716 (0.0152)
Observations	324,926	347,990	347,990	347,990	347,990	347,990	347,990
$R^2$	0.224	0.241	0.309	0.253	0.326	0.317	0.308

*Notes:* Results of estimating model:

$$\Delta x_{i,t} = \psi_i + \psi_{s,t,p} + \gamma \cdot \phi_{i,t} + \eta_{i,t}$$

One observation is one firm for one year between 1998 and 2013 (unbalanced panel).  $\Delta x_{i,t}$  is the delta-log of the variable described on top of each column, in real terms (2010 euros).  $\phi_{i,t}$  is an idiosyncratic shock to firm credit supply, whose construction is detailed in section 3.1. Firm FEs and province $\times$ industry $\times$ year FEs are included. Singleton observations are dropped. A 1% increase in  $\phi_{i,t}$  is the supply shock needed to increase the credit granted to firm  $i$  by 1%. The first column has less observation because some firms might have no credit granted in one year, and therefore delta logs are ill-defined. Standard errors (in parentheses) are clustered at firm level. \*\*\*  $p < 0.01$ , \*\*  $p < 0.05$ , \*  $p < 0.1$

Table 3: Credit Supply Shocks and Productivity Growth

VARIABLES (in delta Log)	Productivity	Productivity	Productivity	Productivity
Functional Form	Cobb-Douglas	Trans-Log	Cobb-Douglas	Trans-Log
Output Measure	Value Added	Value Added	Net Revenues	Net Revenues
	(1)	(2)	(3)	(4)
All industries				
$\phi_{i,t}$	0.0946*** (0.0155)	0.109*** (0.0160)	0.0190*** (0.00477)	0.0259*** (0.00491)
Observations	656,960	656,960	656,960	656,960
$R^2$	0.172	0.185	0.178	0.195
Manufacturing				
$\phi_{i,t}$	0.115*** (0.0178)	0.121*** (0.0186)	0.0303*** (0.00595)	0.0323*** (0.00649)
Observations	347,990	347,990	347,990	347,990
R-squared	0.186	0.200	0.144	0.180

*Notes:* Results of estimating model:

$$\Delta\omega_{i,t} = \psi_i + \psi_{s,t,p} + \gamma \cdot \phi_{i,t} + \eta_{i,t}$$

One observation is one firm for one year between 1998 and 2013 (unbalanced panel). Firm FEs and province $\times$ industry $\times$ year FEs are included. Singleton are dropped. The RHS variable  $\phi_{i,t}$  represents idiosyncratic shock to firm credit supply, and its construction is detailed in section 3.1. A 1% increase in  $\phi_{i,t}$  is the supply shock needed to increase the credit granted to firm  $i$  by 1%. The LHS variable is the first difference of Hicks-neutral productivity residual:  $\Delta\omega_{i,t} = \Delta y_{i,t} - \Delta f(x_{i,t}, \beta)$  where  $y$  is log of net revenues (or log of value added) and  $x$  is a set of inputs. Capital stock, labor, and (for the revenue case only) intermediate inputs are included in  $x$ .  $f(\cdot, \beta)$  is either a first (Cobb-Douglas) or second (Trans-Log) order polynomial in log inputs. Estimation of parameters  $\beta$  is described in section 3.2. Standard errors, in parentheses, are (two-way) clustered at firm and main-lender $\times$ year level. \*\*\* p<0.01, \*\* p<0.05, \* p<0.1

Table 4: Credit Supply Shocks and Productivity Growth: Robustness - Cobb-Douglas Revenue Productivity

VARIABLES (delta Logs)	Productivity	Productivity	Productivity	Productivity	Productivity	Productivity	Productivity	Productivity	Productivity	Productivity
Model	Baseline	Firm Controls	Important Borrowers	Pooled Estimator	Alternative FEs structure	Match Controls	Split Sample	4 Digits Sector	Endogenous Exit	Cost Share
	(1)	(2)	(3)	(4)	(5)	(6)	(7)	(8)	(9)	(10)
All Industries										
$\phi_{i,t}$	0.0190*** (0.00477)	0.0248*** (0.00534)	0.0182*** (0.00540)	0.0131*** (0.00327)	0.0171*** (0.00471)	0.0234*** (0.00604)	0.0197*** (0.00503)	0.0278*** (0.00585)	0.0166*** (0.00465)	0.0256*** (0.00736)
Observations	656,960	483,665	521,741	656,960	656,960	656,960	656,960	587,873	656,960	545,162
R-squared	0.178	0.184	0.192	0.006	0.096	0.178	0.178	0.272	0.177	0.185
Manufacturing										
$\phi_{i,t}$	0.0303*** (0.00595)	0.0362*** (0.00657)	0.0330*** (0.00698)	0.0188*** (0.00443)	0.0321*** (0.00600)	0.0331*** (0.00739)	0.0292*** (0.00633)	0.0401*** (0.00731)	0.0295*** (0.00639)	0.0537*** (0.0104)
Observations	347,990	262,308	280,346	347,990	347,990	347,990	347,990	309,887	347,990	291,071
R-squared	0.144	0.153	0.150	0.004	0.071	0.144	0.144	0.259	0.166	0.161

Notes: Results of estimating model:

$$\Delta\omega_{i,t} = \psi_i + \psi_{s,t,p} + \gamma \cdot \phi_{i,t} + \eta_{i,t}$$

One observation is one firm for one year between 1998 and 2013 (unbalanced panel). Firm FEs and province×industry×year FEs are included. Singletons are dropped.  $\phi_{i,t}$  is an idiosyncratic shock to credit supply, whose construction is detailed in section 3.1. A 1% increase in  $\phi_{i,t}$  is the supply shock needed to increase the credit granted to firm  $i$  by 1%. The LHS variable is the first difference of productivity residual:  $\Delta\omega_{i,t} = \Delta y_{i,t} - \beta_k \cdot \Delta k_{i,t} - \beta_l \cdot \Delta l_{i,t} - \beta_m \cdot \Delta m_{i,t}$  where  $y$  is log of net revenues,  $k$  is log of capital stock,  $l$  is labor (measured by log of wagebill) and  $m$  is log of intermediate inputs. Estimation of parameters  $\beta$  is described in section 3.2. Column (2) add a set of lagged controls to baseline specification: polynomial in size (assets) and the ratios of value added, liquidity, cash flow and bank debt to assets. It excludes observation with missing or zero values for any control variable. Column (3) excludes any firm that, at any point in time, was the recipient of more than 1% of the total credit of any financial intermediary. Column (4) excludes firm FEs. Column (5) includes additively firm FEs, province FEs, year FEs, and industry FEs. Column (6) uses an alternative measure of credit supply shocks which control for match-specific covariates, see section 3.1. Column (7) uses an alternative credit supply shocks estimated with a “split sample” procedure, in order to control for finite sample biases. Column (8) uses a 4-digits (rather than 2) industry definition. It is estimated over less observations because of more singletons. Column (9) estimate productivity allowing for endogenous firm exit, as in [Olley & Pakes \(1996\)](#). Column (10) estimates productivity using the cost share method. It contains less observation because services and materials are entered additively, and not all firms report both intermediates. Standard errors, in parentheses, are (two-way) clustered at firm and main-lender×year level. \*\*\* p<0.01, \*\* p<0.05, \* p<0.1

Table 5: Credit Supply Shocks and Productivity Growth - Heterogeneity - Cobb-Douglas Revenue Productivity

VARIABLES (in delta Log)	Productivity	Productivity	Productivity	Productivity	Productivity
Heterogeneity Dimension	Large (Assets) (1)	Large (Workforce) (2)	Few Lenders (3)	High Sectoral Leverage (4)	High Sectoral Cash Flow (5)
All Industries					
$\phi_{i,t}$	0.0219*** (0.00503)	0.0197*** (0.00557)	-0.0143 (0.0185)	0.00872 (0.00786)	0.0206*** (0.00513)
Interaction	-0.00413 (0.00997)	0.00170 (0.00907)	0.0348* (0.0183)	0.0168* (0.00942)	-0.00340 (0.00940)
Observations	637,989	637,989	656,700	656,960	656,960
R-squared	0.176	0.176	0.178	0.178	0.178
Manufacturing					
$\phi_{i,t}$	0.0365*** (0.00718)	0.0354*** (0.00776)	-0.0128 (0.0244)	0.0105 (0.0105)	0.0352*** (0.00725)
Interaction	-0.0258** (0.0125)	-0.0223* (0.0123)	0.0445* (0.0245)	0.0309** (0.0127)	-0.0138 (0.0131)
Observations	339,747	339,747	347,916	347,990	347,990
R-squared	0.141	0.141	0.144	0.144	0.144

Notes: Model is  $\Delta\omega_{i,t} = \psi_i + \psi_{s,t,p} + \psi_d \cdot D_{i,t} + \gamma \cdot \phi_{i,t} + \gamma_{het} \cdot \phi_{i,t} \cdot D_{i,t} + \eta_{i,t}$ . One observation is one firm for one year between 1998 and 2013 (unbalanced panel). Firm FEs and province $\times$ industry $\times$ year FEs are included. Singleton are dropped. The RHS variable  $\phi_{i,t}$  represents idiosyncratic shock to firm credit supply, and its construction is detailed in section 3.1. A 1% increase in  $\phi_{i,t}$  is the supply shock needed to increase the credit granted to firm  $i$  by 1%. The LHS variable is the first difference of productivity residual:  $\Delta\omega_{i,t} = \Delta y_{i,t} - \beta_k \cdot \Delta k_{i,t} - \beta_l \cdot \Delta l_{i,t} - \beta_m \cdot \Delta m_{i,t}$  where  $y$  is log of net revenues,  $k$  is log of capital stock,  $l$  is labor (measured by log of wagebill) and  $m$  is log of intermediate inputs. Estimation of parameters  $\beta$  is described in section 3.2. Each specification add a categorical dummy  $D_{i,t}$  and the interaction term between the category and  $\phi_{i,t}$ . Categorical dummy  $D_{i,t}$  is equal to one iff: for column (1), firm is in the upper quartile for size, according to previous year assets, for column (2), firm is in the upper quartile for size, according to previous year number of employees, for column (3), firm is in the bottom half according to previous year number of lenders, for column (4), firm is in the top half according to sectoral mean leverage (debt over assets), and for column (5), firm is in the top half according to sectoral mean cash flow over book value of capital. Standard errors, in parentheses, are (two-way) clustered at firm and main-lender $\times$ year level. \*\*\* p<0.01, \*\* p<0.05, \* p<0.1.

Table 6: Exposure to Interbank Market and Productivity Growth

VARIABLES (in delta Log)	Productivity	Productivity	Productivity	Productivity
Functional Form	Cobb-Douglas	Trans-Log	Cobb-Douglas	Trans-Log
Output Measure	Value Added	Value Added	Net Revenues	Net Revenues
	(1)	(2)	(3)	(4)
All Industries				
$ITBK_{i,2006}$	-0.0477** (0.0239)	-0.0574** (0.0257)	-0.0172** (0.00757)	-0.0222*** (0.00781)
Observations	110,746	110,746	110,746	110,746
$R^2$	0.112	0.117	0.101	0.122
Manufacturing				
$ITBK_{i,2006}$	-0.0802** (0.0329)	-0.106*** (0.0361)	-0.00837 (0.00960)	-0.0178* (0.0106)
Observations	58,191	58,191	58,187	58,187
$R^2$	0.134	0.143	0.086	0.113

Notes: Model is  $\Delta\omega_{i,t} = \psi_{s,t,p} + \gamma \cdot ITBK_{i,2006} + \eta_{i,t}$ . One observation is one firm for one year between 2007 and 2009 (unbalanced panel). Province $\times$ industry $\times$ year FEs are included. Singleton are dropped. The RHS variable  $ITBK_{i,2006}$  is the weighted average of firm's  $i$  lenders' liability on the interbank market over assets in 2006. The LHS variable is the first difference of Hicks-neutral productivity residual:  $\Delta\omega_{i,t} = \Delta y_{i,t} - \Delta f(x_{i,t}, \beta)$  where  $y$  is log of net revenues (or log of value added) and  $x$  is a set of inputs. Capital stock, labor, and (for the revenue case only) intermediate inputs are included in  $x$ .  $f(\cdot, \beta)$  is either a first (Cobb-Douglas) or second (Trans-Log) order polynomial in log inputs. Estimation of parameters  $\beta$  is described in section 3.2. Standard errors (in parentheses) are clustered at firm level. \*\*\*  $p < 0.01$ , \*\*  $p < 0.05$ , \*  $p < 0.1$ .

Table 7: Exposure to Interbank Market and Productivity Growth - Placebos

VARIABLES (in delta Log)	Productivity	Productivity	Productivity	Productivity	Productivity	Productivity	Productivity	Productivity
Functional Form	Cobb-Douglas	Trans-Log	Cobb-Douglas	Trans-Log	Cobb-Douglas	Trans-Log	Cobb-Douglas	Trans-Log
Output Measure	Value Added	Value Added	Net Revenues	Net Revenues	Value Added	Value Added	Net Revenues	Net Revenues
Years	[2004, 2006]	[2004, 2006]	[2004, 2006]	[2004, 2006]	[2003, 2005]	[2003, 2005]	[2003, 2005]	[2003, 2005]
	(1)	(2)	(3)	(4)	(5)	(6)	(7)	(8)
All Industries								
$ITBK_{i,2006}$	-0.0148 (0.0225)	-0.0133 (0.0246)	-0.00560 (0.00798)	-0.00702 (0.00793)				
$ITBK_{i,2002}$					-0.0195 (0.0247)	-0.0379 (0.0266)	0.00476 (0.00873)	-0.00486 (0.00832)
Observations	115,042	115,042	115,042	115,042	97,059	97,059	97,059	97,059
$R^2$	0.074	0.079	0.090	0.112	0.078	0.083	0.100	0.111
Manufacturing								
$ITBK_{i,2006}$	0.00230 (0.0307)	0.0194 (0.0340)	-0.0134 (0.0100)	-0.0169 (0.0119)				
$ITBK_{i,2002}$					0.0106 (0.0337)	-0.00820 (0.0372)	0.00118 (0.0106)	0.00908 (0.0122)
Observations	58,949	58,949	58,949	58,949	51,226	51,226	51,226	51,226
$R^2$	0.086	0.090	0.081	0.108	0.078	0.083	0.100	0.111

Notes: Model is  $\Delta\omega_{i,t} = \psi_{s,t,p} + \gamma \cdot ITBK_{i,\tau} + \eta_{i,t}$ . One observation is one firm for one year between 2004 and 2006 or between 2003 and 2005 (unbalanced panel). Province×industry×year FEs are included. Singleton are dropped. The RHS variable  $ITBK_{i,2006}$  is the weighted average of firm's  $i$  lenders' liability on the interbank market over assets in 2006. The LHS variable is the first difference of Hicks-neutral productivity residual:  $\Delta\omega_{i,t} = \Delta y_{i,t} - \Delta f(x_{i,t}, \beta)$  where  $y$  is log of net revenues (or log of value added) and  $x$  is a set of inputs. Capital stock, labor, and (for the revenue case only) intermediate inputs are included in  $x$ .  $f(\cdot, \beta)$  is either a first (Cobb-Douglas) or second (Trans-Log) order polynomial in log inputs. Estimation of parameters  $\beta$  is described in section 3.2. Standard errors (in parentheses) are clustered at firm level. \*\*\* p<0.01, \*\* p<0.05, \* p<0.1.



Table 8: Credit Supply Shock and Productivity-Enhancing Activities

VARIABLES	PCs per unit of Capital (1)	No. of Patents Applications (2)	Patent Growth (3)	<i>R&amp;D</i> (4)	<i>Export</i> (5)	<i>FinCon</i> <sub><i>i</i>,2010</sub> (6)	Management Score (7)
$\phi_{i,t}$	0.808*** (0.289)			0.238* (0.128)	0.152* (0.085)	-1.629*** (0.594)	2.166* (1.116)
$\phi_{i,t-1}$		0.0418** (0.0197)	1.759** (0.883)				
Model	Panel	Panel	Panel	Panel	Panel	Cross Section	Cross Section
Observations	3,632	517,165	13,522	5,991	13,249	506	183
$R^2$	0.968	0.757	0.562	0.872	0.843	0.421	0.020

*Notes:* Columns (1) to (5): model  $Y_{i,t} = \psi_i + \psi_{s,t,p} + \gamma \cdot \phi_{i,t} + \eta_{i,t}$ . One observation is one firm for one year between 1999 and 2001 (unbalanced panel). Firm FEs and province×industry×year FEs are included. Singleton are dropped. The RHS variable  $\phi_{i,t}$  represents idiosyncratic shock to firm credit supply, and (see section 3.1). A 1% increase in  $\phi_{i,t}$  is the supply shock needed to increase the credit granted to firm  $i$  by 1%. RHS variable is lagged in column (2) and (3). LHS of column (1) is the log of number of PCs per 1,000 euros of capital. The LHS variable in column (2) is the number of patent application made from company  $i$  in year  $t$ . Column (3) is the growth rate of the number of patent applications made by company  $i$  in year  $t$  versus  $t-2$ . Columns (4) and (5): model is  $D_{i,t} = \psi_i + \psi_t + \gamma \cdot \phi_{i,t} + \eta_{i,t}$ .  $D_{i,t}$  is a dummy variable taking value 1 iff firm  $i$  in year  $t$  has positive investment in R&D, column (4), or a dummy variable taking value 1 iff firm  $i$  in year  $t$  has positive export revenues, column (5). Column (6): linear probability model is  $FinCon_{i,2010} = \psi_{s,p} + \gamma \cdot \phi_{i,2010} + \eta_i$ . One observation is one firm (cross section). Province×Industry fixed effects are included.  $FinCon_{i,2010}$  is a dummy taking value one iff firm  $i$  reports “difficulties to get external funds” as an important or somehow important obstacle to innovation. Number of PCs, export activity, R&D investments, and self-reported obstacle to innovation are taken from INVIND. Column (7): model is  $MS_{i,t} = \psi + \gamma \cdot \phi_{i,t} + \eta_{i,t}$ . One observation is one firm observed for one or two years (cross section).  $MS_{i,t}$  is firm  $i$  overall management score provided by the World Management Survey (Bloom & Van Reenen, 2007). It takes value from 1-5. Standard errors (in parentheses) are clustered at firm level. \*\*\* p<0.01, \*\* p<0.05, \* p<0.1. See section 8 for more details.

# A Additional Materials on Estimation of Credit Supply Shocks

## A.1 Extensions of the additive growth rate model

### A.1.1 Substitution Patterns

Various forms of the empirical model (1) are widely used in the literature on borrower-lender relations and real effect of financial shocks. However, it does not come without loss of generality. In particular, since companies might have multiple lending relationships, we can expect supply shocks of other connected bank to be included in (1). For instance, letting  $b$  and  $b'$  be the lenders of firm  $i$ , a more complete model of credit is

$$\frac{C_{i,b,t}}{C_{i,b,t-1}} = \frac{C(J_t, D_{i,t}, S_{b,t}, S_{b',t}, X_{i,b,t})}{C(J_{t-1}, D_{i,t-1}, S_{b,t-1}, S_{b',t-1}, X_{i,b,t-1})}$$

which, leads to

$$\Delta c_{i,b,t} = j_t + d_{i,t} + \phi_{b,t} + \alpha \cdot \phi_{b',t} + \epsilon_{i,b,t} \quad (13)$$

Building and estimating a credit demand and supply model with multiple lending relations and more realistic substitution patterns is beyond the scope of this project. Nonetheless, we perform several empirical exercise to assess the consequences of the exclusion of other banks' supply from (1). To do so, we firstly estimate the credit supply shock from the restricted model (1): let  $\hat{\phi}_{b,t}^0$  be the resulting estimate. For each bank-firm pair, we define  $b'$  as the main substitute for  $b$  as the main lender of firm  $i$  during period  $t - 1$ . In case  $b$  is the main lender of firm  $i$ , then  $b'$  is the second main lender of  $i$  in period  $t - 1$ . Then, we include the first-stage estimate of credit shock of bank  $b'$  as an additional control in (1). That is, we estimate:

$$\Delta c_{i,b,t} = j_t + d_{i,t} + \phi_{b,t} + \alpha_t \cdot \hat{\phi}_{b',t}^0 + \epsilon_{i,b,t} \quad (14)$$

Defining  $\hat{\phi}_{b,t}^1$  the estimate of  $\phi_{b,t}$  from (14), the correlation between  $\hat{\phi}_{b,t}^1$  and  $\hat{\phi}_{b,t}^0$  is  $\approx 0.99$  for all years  $t$ , suggesting that the exclusion of supply shocks of other potential borrowers from (3.1) is extremely unlikely to affect significantly our results. We conclude that ignoring substitution and complementarity does not significantly affect the impact of credit supply shocks on productivity.<sup>38</sup>

### A.1.2 Loan and Relation Characteristics

We may relax Assumption 2, by imposing **Assumption 2b**:

$$\epsilon_{i,b,t} = \alpha \cdot o_{i,b,t-1} + \acute{\epsilon}_{i,b,t}$$

and

$$\acute{\epsilon}_{i,b,t} \perp \mathcal{D}_i, \mathcal{S}_b$$

---

<sup>38</sup>Specification (14) considers only the effect of the main alternative lender. However, firms' financing decision might be affected by idiosyncratic shocks to all other connected lenders. Therefore, for each bank-firm pair, we compute  $\hat{\phi}_{-b,i,t}^0$  as the average of  $\hat{\phi}_{b',t}^0$  of all banks lending to  $i$  but  $b$ . Then, we consider the model:

$$\Delta c_{i,b,t} = j_t + d_{i,t} + \phi_{b,t} + \alpha_t \cdot \hat{\phi}_{-b,i,t}^0 + \epsilon_{i,b,t} \quad (15)$$

Supply shocks estimated from (15) are, again, extremely highly correlated with the ones estimated from the main specification (3.1) (correlation coefficient is around 98 per cent for all years). Replicating our main specification using credit supply shocks estimated in (14) and (15) confirms our results.

with  $o_{i,b,t-1}$  are observable characteristics of the lending relations between firm  $i$  and financial intermediary  $b$ . The vector  $o_{i,b,t-1}$  includes: size of the loan relative to borrower's total credit received, size of the loan relative to lender's total credit granted, interest rate,<sup>39</sup> length of the lending relations, type of credit instrument used, presence of past non-performing loans.

Assumption 2b allows to estimate bank and firm factors from

$$\Delta c_{i,b,t} = c_t + d_{i,t} + \phi_{b,t} + \alpha \cdot x_{i,b,t-1} + \epsilon_{i,b,t} + approx_{i,b,t} \quad (16)$$

As match-specific controls, we include interest rates, length of lending relations, share of  $C_{i,b,t-1}$  in the portfolio of the lender and in the portfolio of the borrower, type of credit instrument used, share of non-performing loans, share of credit covered by collateral. Supply shocks estimated from equation (16) and (3.1) show correlations above 94% for most years, which mitigate concern that unobservable elements of  $\epsilon_{i,b,t}$  are significantly biasing estimate of  $\phi_{b,t}$ . Section 5.1 shows that the main results of the paper are unaffected by using the alternative credit supply shocks derived from decomposition (16).

Summing-up, we develop new tests to estimate whether substitution and complementarity patterns between lenders, and bank-firm match-specific shocks affect our results on the impact of credit supply on productivity. In our data, this does not seem to be the case. Yet, notice that this may not be the case in other more specialized or concentrated markets, such as the one of syndicated loans. Given the widespread use of additive growth rate model, we suggest that our tests represent an important sanity check to be performed by researchers.

## A.2 Credit Supply Shocks and Credit Applications to new Lenders

The credit register contains information about firms' application for loans or credit lines with new lenders.<sup>40</sup> We expect a borrower to be less likely to apply for credit with new lenders when the lenders with whom it is already connected are expanding credit supply. On the other side, if the additive growth rate model is severely misspecified, then what we define as credit supply might be contaminated by demand-side factors. Consequently, we would expect a positive correlation between these demand-side factors and loans applications with all lenders (and, therefore, with new lenders as well). Therefore, we estimate the model:

$$\begin{aligned} App_{i,t} &= \psi_i + \psi_{p,s,t} + \gamma \cdot \phi_{i,t} + \eta_{i,t} \\ Pr(App_{i,t} > 0) &= \psi_i + \psi_{p,s,t} + \gamma \cdot \phi_{i,t} + \eta_{i,t} \\ App_{i,t} | App_{i,t} > 0 &= \psi_i + \psi_{p,s,t} + \gamma \cdot \phi_{i,t} + \eta_{i,t} \end{aligned}$$

where  $App_{i,t}$  is the number of previously unconnected lenders receiving a credit application from  $i$  in period  $t$ . Results are reported in Table A.1, which shows that firms receiving positive credit supply shocks decrease their search for alternative lenders both on the intensive and on the extensive margin. They confirm the soundness of the procedure to disentangle supply-side variation from demand-side factors.

<sup>39</sup>We need to impute interest rates for roughly a third of the observations.

<sup>40</sup>The very reason of having a national credit register is to provide lenders with information about potential borrowers' credit history. When a lender submits a query to the credit register to seek information about a firm which is not among its current borrowers, the researcher infer that the firm applied for a loan or a credit line with that bank.

### A.3 Which factors shape credit supply?

Jensen & Johannesen (2016) show that Danish banks entering the 2007-2008 financial crisis with weaker balance sheets declined sharply their lending to retail customers, depressing their consumption. They use a proxy for liquidity in 2007 (loans over deposits) to measure balance sheet strength. To provide additional evidence that our measure of credit supply captures bank-level shocks, for each year  $t$  between 2007 and 2009 (credit crunch) we estimate the linear model:

$$\phi_{b,t} = \phi_t + \gamma \cdot \left( \frac{Loans}{Deposits} \right)_{b,2007} + \eta_{b,t}$$

Results are presented in Column (1) of Table A.10: banks with lower liquidity at the beginning of the credit crunch, decreased more their credit supply. We follow Jensen & Johannesen (2016) and rely on 2007 balance sheet to construct the RHS variable. Results are robust to use 2006 values. Column (2) of the same table uses the capital adequacy ratio (the ratio of bank capital to risk-weighted assets) as an alternative measure of financial strength: better capitalized banks decreased less their credit supply during the credit crunch.

During the Europe sovereign debt crisis, the spread on bonds issued by Italian government (and other southern European countries) increased sharply. Investors responded by acquiring more of these assets, “crowding out” credit to private non-financial corporation (Broner *et al.*, 2014). Therefore, for each year between 2010 and 2013, we estimate the model:

$$\phi_{b,t} = \phi_t + \gamma \cdot \Delta \left( \frac{Sovereign}{Assets} \right)_{b,t} + \eta_{b,t}$$

where  $\left( \frac{Sovereign}{Assets} \right)_{b,t}$  is the share of sovereign bonds on the total assets of bank  $b$ . Results are presented in column (3) of of Table A.10: banks increasing more their exposure to sovereign debt decreased their credit supply to corporate borrowers.

After an M&A episode, acquired banks generally reduce (in the short-run) their supply of credit to pre-existing borrowers (Buono & Formai, 2013). Following this intuition, we estimate the model:

$$\phi_{i,t} = \phi_i + \phi_{t,p,s} + \gamma \cdot MA_{i,t} + \eta_{i,t}$$

where  $\phi_{i,t}$  is the credit supply shock experienced by firm  $i$  and  $MA_{i,t}$  is the share of  $i$ 's previous period lenders which are being acquired by another financial institution. Results are presented in column (4) of Table A.10: credit supply is negatively affected by lenders' M&A episodes.

The collection of results presented in this section, being consistent with the relevant literature, provide additional support to our measure of credit supply shocks.

## B Production with Heterogeneous Credit Constraints

### B.1 Main model

Firm  $i$ , operates in sector  $s$  and province  $p$ . For simplicity, we omit the subscript  $s$ , although all parameters, are industry-specific. In each year  $t$ , firm combines capital, labor and materials to generate revenues:

$$Y_{i,t} = \exp\{\omega_{i,t}\} F(L_{i,t}, K_{i,t}, M_{i,t}, \beta) \tag{17}$$

or value added

$$VA_{i,t} = \exp\{\omega_{i,t}\}F(L_{i,t}, K_{i,t}, \beta)$$

As it is common in the literature (Olley & Pakes, 1996) we assume that productivity can be decomposed into a structural component and an error term:

$$\omega_{i,t} = \tilde{\omega}_{i,t} + \epsilon_{i,t}^Y$$

$\tilde{\omega}_{i,t}$  is correlated over time and it is known by the firm before starting production. Therefore, it affects inputs acquisition and other firm decisions.  $\epsilon_{i,t}^Y$  is an *iid* shock, which takes place after input decisions have been made and does not convey information about future productivity. It is often referred as measurement error of output, or “pure luck”.

### Capital accumulation

Capital stock is accumulated according to the usual law of motion:

$$K_{i,t} = I_{i,t} + (1 - \delta_t)K_{i,t-1}$$

### Prices

Firms are price-takers on the input markets. Prices of materials are assumed to be shaped by national prices of inputs and by local inflation shocks (measured by local CPI):

$$P_{p,t}^M = P_t^M \cdot P(cpi_{p,t})$$

we do not observe firm level or local level prices of intermediate inputs, so we need to assume a form for  $P(\cdot)$ .<sup>41</sup>

### Variable Profits and Utility Function

We focus the discussion on revenue productivity. The value added case is a straightforward simplification.<sup>42</sup> Variable profits are:

$$\begin{aligned} \pi(K_{i,t}, L_{i,t}, \tilde{\omega}_{i,t}, w_t, P_{p,t}^M, \epsilon_{i,t}^Y) &= \\ &= \exp\{\omega_{i,t} + \epsilon_{i,t}^Y\}F(L_{i,t}, K_{i,t}, M_{i,t}, \beta) - w_t L_{i,t} - M_{i,t} P_{p,t}^M \end{aligned}$$

The firm maximizes owner’s utility from the dividend stream  $Div_{i,t}$ :

$$u(Div_{i,t}, \epsilon_t^U) + E \left[ \sum_{\tau>t} \left( \frac{1}{1+R} \right)^{\tau-t} u(Div_{i,\tau}, \epsilon_\tau^U) \right]$$

### Credit Supply Shocks

At the beginning of the period, firm  $i$  is connected to a subset of the banks operating in the country,  $\mathcal{B}_{i,t-1}$ . Each bank “experience” a credit supply shock  $\phi_{b,t}$ . Firm  $i$  receive a credit supply shock equal to  $\phi_{i,t} = \sum_{b \in \mathcal{B}_{i,t-1}} \phi_{b,t} \cdot wc_{b,i,t-1}$  where weights are proportional to the share of credit

<sup>41</sup>See section 2.

<sup>42</sup>For a discussion see Akerberg *et al.* (2015) and De Loecker & Scott (2016).

received from firm  $i$  from each lender in previous period.

We assume  $B_{i,t}$  and  $\{wc_{b,i,t-1}\}$  evolve exogenously, while the quantity of debt is endogenously chosen. In section 6 we exploit a natural experiment to control for assortative matching between borrowers and lenders and we show that our empirical results hold.

### Budget and Credit constraints

Firm faces a budget constraints:

$$\begin{aligned} Div_{i,t} + K_{i,t} + B_{i,t-1} (1 + r_{i,t}) + Adj \left( L_{i,t}, L_{i,t-1}, K_{i,t}, K_{i,t-1}, J_t, \epsilon_{i,t}^{adj} \right) = \\ = \pi_{i,t} + B_{i,t} + (1 - \delta_t) K_{i,t-1} \end{aligned}$$

where  $B_{i,t}$  is the quantity of euros borrowed,  $Adj(\cdot)$  are adjustment costs for labor and capital,  $J_t$  is the set of all industry-wide state variables.

credit constraint

$$B_{i,t} \leq K_{i,t-1} \cdot \Gamma(B_{i,t-1}, \phi_{i,t}, \tilde{\omega}_{i,t}, J_t)$$

we also allow single inputs to have specific financially-generated constraints.

$$\begin{aligned} M_{i,t} &\leq K_{i,t-1} \cdot \Gamma^M(B_{i,t-1}, \phi_{i,t}, \tilde{\omega}_{i,t}, J_t) \\ K_{i,t} &\leq K_{i,t-1} \cdot \Gamma^K(B_{i,t-1}, \phi_{i,t}, \tilde{\omega}_{i,t}, J_t) \\ L_{i,t} &\leq K_{i,t-1} \cdot \Gamma^L(B_{i,t-1}, \phi_{i,t}, \tilde{\omega}_{i,t}, J_t) \end{aligned}$$

furthermore, we assume the function  $\Gamma^M$  is increasing in its second and third arguments.

The presence of general credit constraints does not imply that intermediate inputs are constrained. However, we want to allow for this possibility, since firms might need to pay in advance part of the material inputs and availability of credit (especially credit lines) might limit their ability to do so. Whether or not an input is effectively contained by availability of external funds depends on the relative cash cycle. For instance, capital investments might be more sensitive to credit availability than labor because they have to be paid fully in advance. However, while firms make financial and real decision in continuous time (or every day), our model discretize time in yearly periods, as it is commonly do by the literature because balance-sheets data are available at yearly frequency. Allowing for input-specific constraints is a way to partly reconcile the model with reality.

### Timing

At the beginning of the period firms observe  $\tilde{\omega}_{i,t}$ ,  $\epsilon_{i,t}^K$ ,  $\epsilon_{i,t}^{adj}$ ,  $\phi_{i,t}$  and all elements of  $J_t$  (like  $\epsilon_t^u$  and  $cpi_{p,t}$ ). Then, it choose jointly  $B_{i,t}$  and all inputs  $K_{i,t}$ ,  $L_{i,t}$  and  $M_{i,t}$ . Firms do not observe the non transmitted error  $\epsilon_{i,t}^Y$  until the end of the period. When the random shock  $\epsilon_{i,t}^Y$  is realized,  $Div_{i,t}$  is set as residual from the budget constraint and it is consumed.

### Law of Motion

The non transmitted shock  $\epsilon_{i,t}^Y$  is pure luck and, therefore, it is *i.i.d* and independent of any other component of the model. Without loss of generality, we set  $E[\exp\{\epsilon_{i,t}^Y\}] = 1$ . The law of motion of all other shocks  $\epsilon_{i,t}$  is left unrestricted. However, notice that the absence of any shock directly affecting intermediate inputs is essential for identification.

Following the tradition of the control function, we impose a law of motion for productivity. That is,

$$E[\tilde{\omega}_{i,t}|\mathcal{I}_{t-1}] = E[\tilde{\omega}_{i,t}|\tilde{\omega}_{i,t-1}, \phi_{i,t-1}, J_{t-1}] \quad (18)$$

where  $\mathcal{I}_{t-1}$  is firm's information set at time  $t-1$ . Assumption 18 relaxes the classical Markovian structure by allowing credit supply to affect productivity dynamics.

Furthermore, defining:

$$\zeta_{i,t} := \tilde{\omega}_{i,t} - E[\tilde{\omega}_{i,t}|\mathcal{I}_{t-1}] \quad (19)$$

and

$$\zeta_{i,t}^\phi := \phi_{i,t} - E[\phi_{i,t}|\phi_{i,t-1}] \quad (20)$$

we assume  $\zeta^{\phi_{i,t}}$  is independent of all  $\epsilon$ 's.

### Demand for intermediate inputs

The optimal quantity of intermediate input is

$$M^*(K_{i,t}, L_{i,t}, K_{i,t-1}\tilde{\omega}_{i,t}, J_t, B_{i,t-1}, P_{p,t}^M) = \min\{M^{unc}(K_{i,t}, L_{i,t}, \tilde{\omega}_{i,t}, P_{p,t}^M); K_{i,t-1} \cdot \Gamma^M(B_{i,t-1}, \phi_{i,t}, \tilde{\omega}_{i,t}, J_t)\}$$

where  $M^{unc}$  solves

$$\frac{\partial F(L_{i,t}, K_{i,t}, M, \beta)}{\partial M} \exp\{\tilde{\omega}_{i,t}\} = P_{p,t}^M$$

under assumptions above,<sup>43</sup>  $M^*$  is increasing in productivity for each level of the other factors. Therefore,  $\exists$  an unknown function  $M^{-1}$  such that:

$$\tilde{\omega}_{i,t} = M^{-1}(K_{i,t}, L_{i,t}, K_{i,t-1}, \phi_{i,t}, cpi_{p,t}, J_t)$$

which allows to write

$$Y_{i,t} = \exp\{\epsilon_{i,t}^Y\} F(L_{i,t}, K_{i,t}, M_{i,t}, \beta) \exp\{M^{-1}(K_{i,t}, L_{i,t}, K_{i,t-1}, \phi_{i,t}, cpi_{p,t}, J_t)\}$$

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<sup>43</sup>Formally, we assume that  $M$  is chosen within a set  $\mathcal{A}_M$  such that  $F(K, L, \cdot)$  is increasing in its last argument for each value of  $K$  and  $L$ . Then,  $M^{unc}$  is increasing in  $\tilde{\omega}$  by Topkis theorem. This is trivially true for the Cobb Douglas case, as long as  $K > 0$  and  $L > 0$ .

therefore, for some unknown function  $\tilde{\Psi}$

$$Y_{i,t} = \exp\{\epsilon_{i,t}^Y\} \tilde{\Psi}(L_{i,t}, K_{i,t}, M_{i,t}, K_{i,t-1}, \phi_{i,t}, cpi_{p,t}, J_t) \quad (21)$$

### B.1.1 Estimation of the Production Function

We aim at estimating the shape of the production function  $F(\cdot)$ , in order to back out productivity residual and investigate the joint dynamics of credit supply and productivity. Firstly, we write the main equations in logarithmic terms. Variables in logs are indicated by lowercase letters.

Revenues are:

$$y_{i,t} = \tilde{\omega}_{i,t} + \epsilon_{i,t}^Y + f(l_{i,t}, k_{i,t}, m_{i,t}, \beta) \quad (22)$$

where  $f(\cdot)$  is known up to the parameter  $\beta$ , which we aim to estimate. Revenues, can also be written as

$$y_{i,t} = \Psi(l_{i,t}, k_{i,t}, m_{i,t}, k_{i,t-1}, \phi_{i,t}, cpi_{p,t}, J_t) + \epsilon_{i,t}^Y \quad (23)$$

for some unknown function  $\Psi$ . Following [Akerberg \*et al.\* \(2015\)](#), we estimate the model in two stages. In the first stage we purge the output from the noise  $\epsilon_{i,t}^Y$ . We estimate  $\Psi$  as  $E_t[y_{i,t}|l_{i,t}, k_{i,t}, m_{i,t}, k_{i,t-1}, \phi_{i,t}, cpi_{p,t}]$ .<sup>44</sup>

From equation (18) we can write

$$\tilde{\omega}_{i,t} = g_t(\tilde{\omega}_{i,t-1}, \phi_{i,t-1}) + \zeta_{i,t}$$

with  $g_t$  unknown.

By definition, we have

$$E[\zeta_{i,t} + \epsilon_{i,t}^Y | \mathcal{I}_{t-1}] = 0$$

therefore,

$$E[y_{i,t} - f(l_{i,t}, k_{i,t}, m_{i,t}, \beta) - g_t(\Psi_{i,t-1} - f(l_{i,t-1}, k_{i,t-1}, m_{i,t-1}, \beta), \phi_{i,t-1}) | \mathcal{I}_{t-1}] = 0$$

leading to the moment condition

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<sup>44</sup>  $E[y_{i,t}|x_{i,t}, J_t] = \phi(x_{i,t}, J_t)$  for some unknown function  $\phi$ , which we approximate as a third order polynomial in  $x_{i,t}$  plus year fixed effects. We follow this approximation procedure through the paper.



$$E \left[ \begin{array}{c|c} \Psi_{i,t} - f(l_{i,t}, k_{i,t}, m_{i,t}, \beta) - & l_{i,t-1} \\ +g_t(\Psi_{i,t-1} - f(l_{i,t-1}, k_{i,t-1}, m_{i,t-1}, \beta), \phi_{i,t-1}) & inv_{i,t-1} \\ & \Psi_{i,t-1} \\ & m_{i,t-1} \\ & \dots \end{array} \right] = 0 \quad (24)$$

Moments (24) allow joint estimation of the structural parameter  $\beta$  and of the unknown function  $g_t$ .<sup>45</sup>

We parametrize  $f(\cdot)$  as either linear (Cobb-Douglas) or quadratic (Trans-Log) in logs. These two functions can be seen as a first and second order log-linear approximation of any smooth production function  $F(\cdot)$ . Since our results are extremely similar between Cobb-Douglas and Trans-Log, we do not believe it is useful to add higher order terms. Production functions are industry-specific. We drop sectors for which less than 300 firm-year observations are available, because of difficulties in estimating production function with few observations.

The control function approach allows to estimate production function parameters by controlling for simultaneity bias in the choice of inputs.<sup>46</sup> Furthermore, the inclusion of local price shocks  $cpi_{p,t}$  in the control function overcomes the non-identification results of [Gandhi et al. \(2011\)](#). In our baseline specification, we do not include endogenous exit decision in the model. Section 5.1 shows that such an inclusion does not significantly affect the main results of this paper.

If credit supply affects productivity, then it is correlated with  $\zeta_{i,t}$ . Moreover,  $\phi_{i,t}$  is correlated over time: in fact, regression of  $\phi_{i,t}$  on  $\phi_{i,t-1}$  gives a coefficient of  $\approx 0.5$  if no fixed effect is included and  $\approx 0.2$  if firm fixed effects are included. Furthermore, it affects input acquisition, as documented by section 4. Therefore, if one excludes credit supply shocks from the model, past inputs are correlated with the productivity innovation, and there are no valid instruments to identify the parameter of interests.

## Results

Table A.2 provides some descriptives for the Cobb-Douglas production function estimates.<sup>47</sup> The mean<sup>48</sup> elasticity of value added to capital (to labor) is  $\approx .17$  ( $\approx .64$ ) for the whole economy and  $\approx .19$  ( $\approx .62$ ) for manufacturing. The mean elasticity of net revenues to capital (to labor) is  $\approx .07$  ( $\approx .14$ ) for the whole economy and  $\approx .04$  ( $\approx .13$ ) for manufacturing. The mean elasticity of net revenues to intermediate inputs is  $\approx .81$  for both manufacturing and all industries.<sup>49</sup>

<sup>45</sup>We follow [De Loecker & Warzynski \(2012\)](#) and we perform this second stage in two steps. For each guess a parameter value  $\beta_{guess}$ , we can compute a corresponding  $\omega_{i,t}(\beta_{guess})$ . Then, by regressing  $\omega_{i,t}(\beta_{guess})$  on a polynomial in  $\omega_{i,t-1}(\beta_{guess})$  and  $\phi_{i,t-1}$  plus year fixed effects we get a sample analog of  $\zeta(\beta_{guess})$ . We estimate  $\beta$  by minimizing then sample analog of  $E[\zeta(\beta_{guess}) \cdot instruments_{i,t-1}]$ .

<sup>46</sup>That is, since more productive firms are likely to acquire more inputs, a simple regression of output on inputs does not recover the structural parameters of interest.

<sup>47</sup>As it is shown in section 5 the relations between productivity and credit does not change if one consider CD or Trans-Log production function, therefore we do not find it useful to analyses in detail the more complicated case.

<sup>48</sup>We take averages weighted for number of observation in the sample of the main specification.

<sup>49</sup>Under the assumption of single good producers, we can translate these revenue production function estimate into a quantity production function. The relations depend on the competitive structure of the product market, see, for

## B.2 An alternative empirical model of production

The credit supply shock  $\phi_{i,t}$  is a weighted average of supply shocks  $\phi_{b,t}$  of the banks connected to firm  $i$  in the previous period. A potential critique of the main model is that it includes credit supply *changes* as an element of the *level* of the constraint faced each period by the firm. Indeed, one might prefer to include  $S_{b,t}$  (i.e., all the bank-level factors affecting  $b$ 's ability and willingness to provide credit to its borrowers) into the credit constraint function  $\Gamma$ . An additional potential problem is that, to perform the inversion of the error, the researcher need to observe the exact value of  $\phi_{i,t}$ . The credit shifter estimated as in section 3.1 might be considered a proxy of the real variation in credit constraints. For instance, the actual credit supply faced by a firm can be affected by new banks it connects to during the year. In this section, we provide an alternative model of production with credit constraints that address all these issues, at the cost of relying on a first-order log-linearization of the estimating equations.<sup>50</sup> This model provides an alternative firm-specific estimate of productivity growth: the impact of credit supply shocks on it is qualitatively and quantitatively similar.

Production, utility and budget constraints are as in section (3.2). Credit constraints are:

$$B_{i,t} \leq K_{i,t-1} \cdot \Gamma (B_{i,t-1}, S_{b(i),t}, \tilde{\omega}_{i,t}, J_t)$$

and

$$\begin{aligned} M_{i,t} &\leq K_{i,t-1} \cdot \Gamma^M (B_{i,t-1}, S_{b(i),t}, \tilde{\omega}_{i,t}, J_t) \\ K_{i,t} &\leq K_{i,t-1} \cdot \Gamma^K (B_{i,t-1}, S_{b(i),t}, \tilde{\omega}_{i,t}, J_t) \\ L_{i,t} &\leq K_{i,t-1} \cdot \Gamma^L (B_{i,t-1}, S_{b(i),t}, \tilde{\omega}_{i,t}, J_t) \end{aligned}$$

where  $b(i)$  is the set of banks connected to  $i$  at the beginning of the period, and  $S_{b(i),t}$  are bank-level factors determining credit supply. Log output is:

$$Y_{i,t} = \tilde{\Psi} (L_{i,t}, K_{i,t}, M_{i,t}, K_{i,t}, K_{i,t-1}, S_{b(i),t}, cpi_{p,t}, J_t) + \epsilon_{i,t}^Y \quad (25)$$

and the law of motion of productivity is

$$E[\exp\{\tilde{\omega}_{i,t}\} | \mathcal{I}_{t-1}] = E[\exp\{\tilde{\omega}_{i,t}\} | \omega_{t-1}, S_{b(i),t-1}, J_{t-1}] = G_t (\tilde{\omega}_{i,t-1}, S_{b(i),t-1}) \quad (26)$$

log-linearizing equations (25) and (26) and taking first differences yields:

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instance, De Loecker (2011). If firms are price takers on the output market, then the quantity elasticities are equal to revenue elasticities. However, if firms compete under monopolistic competition and consumers have CES utility, then, for each input  $x$  the relations between quantity and revenue elasticity is  $\beta^{x,quantity} = \beta^x \cdot \frac{\sigma}{\sigma-1}$  where  $\sigma$  is the elasticity of demand. We compute sector level estimate of  $\sigma$  following Pozzi & Schivardi (2016) in order to calculate the mean quantity-elasticities for manufacturing, which are, respectively  $\approx .05$  for capital,  $\approx .17$  for labor and  $\approx 1.06$  for intermediate inputs.

<sup>50</sup>For a critique of the use of log-linear approximation, see Carroll (2001).

$$\begin{aligned}\Delta y_{i,t} &= \psi_t + \psi_l \Delta l_{i,t} + \psi_k \Delta k_{i,t} + \psi_m \Delta m_{i,t} \\ &+ \psi_{km1} \Delta k_{i,t} + \psi_\phi \phi_{b(i),t} + \psi_{cpi} \Delta cpi_{p,t} + approx_{i,t}^y + \Delta \epsilon_{i,t}^Y\end{aligned}$$

and

$$\Delta \tilde{\omega}_{i,t} = g_t + g_\omega \cdot \Delta \tilde{\omega}_{i,t-1} + g_\phi \cdot \phi_{b(i),t-1} + v_{i,t} + approx_{i,t}^\omega \quad (27)$$

where  $\phi_{b(i),t} = c_3 \Delta s_{b(i),t}$  as defined in section 3.1, and  $\zeta_{i,t}$  is an expectation error and  $approx_{i,t}^\omega$  approximation error. The econometrician observe  $\widehat{\phi}_{i,t}$  which is a noisy proxy of  $\phi_{b(i),t}$ :

$$\widehat{\phi}_{i,t} = \phi_{b(i),t} + \epsilon_{i,t}^\phi \quad (28)$$

the measurement error  $\epsilon_{i,t}^\phi$  is assumed to be *i.i.d* and uncorrelated over time. We can rewrite the equations as

$$\begin{aligned}\Delta y_{i,t} &= \Delta \psi_{i,t} + approx_{i,t}^y + \Delta \epsilon_{i,t}^Y = \\ &= \psi_t + \psi_l \cdot \Delta l_{i,t} + \psi_k \cdot \Delta k_{i,t} + \psi_m \cdot \Delta m_{i,t} + \psi_{km1} \cdot \Delta k_{i,t-1} + \\ &\quad + \psi_\phi \cdot \widehat{\phi}_{i,t} + \psi_{cpi} \cdot \Delta cpi_{p,t} + approx_{i,t}^y - \epsilon_{i,t}^\phi + \Delta \epsilon_{i,t}^Y\end{aligned}$$

In the first stage, we produce an estimate for  $\Delta \psi_{i,t}$  from equation . Since  $\epsilon_{i,t}^\phi$  is correlated with input acquisition at period  $t$ , we use past values of inputs as instrument, together with contemporaneous value of  $\phi_{i,t}$  and  $cpi$ .

Then, we consider the log-lin approximation of the expected value of productivity

$$\Delta \tilde{\omega}_{i,t} = g_t + g_\omega \cdot \Delta \tilde{\omega}_{i,t-1} + g_\phi \cdot \phi_{i,t-1} - \epsilon_{i,t-1}^\psi + \Delta \zeta_{i,t} + \Delta approx_{i,t}^\omega \quad (29)$$

to estimate the model we need to add an assumption on the approximation error:

$$E [\Delta approx_{i,t}^\omega | \mathcal{I}_{t-2}] = 0$$

implying moment conditions

$$E [-\epsilon_{i,t-1}^\psi + \Delta \zeta_{i,t} + \Delta approx_{i,t}^\omega | \mathcal{I}_{t-2}] = 0$$

or, equivalently

$$E \left[ \begin{array}{c|c} \Delta\psi_{i,t} - \Delta f(l_{i,t}, k_{i,t}, m_{i,t}, \beta) & l_{i,t-2} \\ -g_t - g_\omega \cdot (\Delta\psi_{i,t-1} - \Delta f(l_{i,t-1}, k_{i,t-1}, m_{i,t-1}, \beta)) - g_\phi \cdot \phi_{i,t-1} & inv_{i,t-2} \\ & \Psi_{i,t-2} \\ & m_{i,t-2} \\ & \dots \end{array} \right] = 0$$

which allow to estimate the parameter  $\beta$  and recover productivity residual  $\omega_{i,t}$ . Table A.3 shows that the effect of credit supply shock on productivity is extremely similar if production function is estimated with this alternative procedure rather than baseline of Table 3.

## C Additional Materials on Credit Supply and Productivity Growth

### C.1 TFP, TFPR and TFPQ - cont'd

We follow De Loecker (2011) and consider a firm producing quantity  $Q_{i,t}$  of a single differentiated good, at price  $P_{i,t}$ , and facing a CES demand function. Let its production function be a Cobb-Douglas. Quantity produced (supply) is

$$Q_{i,t} = \exp\{\omega_{i,t}^q + f(l_{i,t}, k_{i,t}, m_{i,t}, \beta^q)\} = \exp\{\omega_{i,t}^q + \beta_l^q \cdot l_{i,t} + \beta_k^q \cdot k_{i,t} + \beta_m^q \cdot m_{i,t}\}$$

Quantity sold (demand) is:

$$Q_{i,t} = \left( \frac{P_{i,t}}{P_t} \right)^{-\sigma} \exp\{\theta_{i,t}\}$$

where  $P_t$  is national deflator and  $\theta_{i,t}$  reflects demand conditions, both endogenous (e.g. quality of the product offered) and exogenous (e.g. local economic shocks) with respect to firm's activity. We follow Pozzi & Schivardi (2016) and refer to  $\theta$  as "market appeal". Then, the deflated revenues are:

$$Y_{i,t} = \frac{P_{i,t} \cdot Q_{i,t}}{P_t} = Q_{i,t}^{\frac{\sigma-1}{\sigma}} \cdot \exp\{\theta_{i,t}\}$$

therefore, taking logs:

$$\begin{aligned} y_{i,t} &= \frac{1}{\sigma} \cdot \theta_{i,t} + \frac{\sigma-1}{\sigma} \cdot \omega_{i,t}^q + \frac{\sigma-1}{\sigma} \cdot f(l_{i,t}, k_{i,t}, m_{i,t}, \beta^q) = \\ &= \frac{1}{\sigma} \cdot \theta_{i,t} + \frac{\sigma-1}{\sigma} \cdot \omega_{i,t}^q + \beta_l \cdot l_{i,t} + \beta_k \cdot k_{i,t} + \beta_m \cdot m_{i,t} \end{aligned}$$

with  $\beta_x = \frac{\sigma-1}{\sigma} \cdot \beta_x^q$ . The growth rate of productivity is:

$$\Delta\omega_{i,t} = \frac{1}{\sigma} \cdot \Delta\theta_{i,t} + \frac{\sigma-1}{\sigma} \cdot \Delta\omega_{i,t}^q$$

which clarifies that the an increase in any revenue-based measure of productivity can be generated either by an increase in technical efficiency or by an increase in market appeal of firm’s  $i$  product. Productivity-enhancing activities can affect both terms. For instance a process innovation is more likely to increase  $\omega_{i,t}^q$ , while a product innovation should mainly affect  $\theta_{i,t}$ , see [Hall \(2011\)](#) and [Peters et al. \(2017b\)](#).

The main empirical specification (equation 9) of this paper can be re-written as:

$$\frac{1}{\sigma} \cdot \Delta\theta_{i,t} + \frac{\sigma - 1}{\sigma} \cdot \Delta\omega_{i,t}^q = \psi_i + \psi_{p,s,t} + \gamma \cdot \phi_{i,t} + \eta_{i,t} \quad (30)$$

Equation (30) highlights that an empirical investigation based on data on revenues rather than quantities presents both challenges and opportunities. The main challenge is to provide evidence that the results are not driven by correlation between output demand (or other local competitive conditions)<sup>51</sup> and credit supply factors: evidence provided in section 5.1 and 6 are reassuring on this regard. At the same time, we have the opportunity to take into account other sources of productivity increase, besides technical efficiency ([Hall, 2011](#)). These encompass improvements in quality of the product offered and access to new markets or new niches that may result in an increase in markups. Measures of pure technical efficiency may ignore changes in product quality, which are found to explain the vast majority of the heterogeneity in firm size ([Hottman et al. , 2016](#)). Moreover, notice that it is difficult to properly define quantity productivity in service industries, where products are intrinsically non-homogeneous. How to measure, for instance, the “quantity” produced by a law firm?

An additional concern is that under a more general (inverse) demand function,  $P_{i,t} = D(Q_{i,t}, \theta_{i,t}, P_t)$ , credit supply might alter pricing incentives and create an increase in measured revenue productivity even without a change in technical efficiency  $\omega_{i,t}^q$  or market appeal  $\theta_{i,t}$ . In fact, productivity growth can be also written as:

$$\Delta\omega_{i,t} = \Delta p_{i,t} + \Delta\omega_{i,t}^q$$

However, this is a not a very worrisome concern. In fact, it is shown in the paper that positive credit supply shocks increase input acquisition. Therefore, even if productivity does not respond to credit shocks, quantity produced also goes up. As long as demand is decreasing in prices (implying inverse demand is decreasing in quantity), a firm has to set lower prices in order to sell the additional quantity produced. Then, a positive credit supply shocks decreases prices and, consequently, revenue productivity, for a given level of technical efficiency and product appeal. We show, instead, a positive effect of credit on productivity.<sup>52</sup>

Does  $\theta$  or  $\omega^q$  respond more strongly to a credit supply shocks? Let us consider again the case of monopolistic competition and CES demand. Let us assume that equation (30) can be decomposed

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<sup>51</sup>For instance, if there is an overlap between the output market of the borrower and the lending market of the lender, then healthier lenders are also connected to firms receiving positive demand shocks.

<sup>52</sup>It is possible that more complex interaction between financial constraints and pricing incentives might arise because of the presence of demand dynamics (e.g. demand today depends on prices set yesterday). It is not possible to exclude that this might be the case under some assumptions. However, [Chevalier & Scharfstein \(1996\)](#) show that more financially constrained firms set higher prices (at least in the supermarket industry, during recessions) because they are more likely to exit the market and, therefore, “care less” about future demand, which support the causal interpretation of our results.

in two parts

$$\begin{aligned}\Delta\theta_{i,t} &= a_i + a_{p,s,t} + \gamma^\theta \cdot \phi_{i,t} + \epsilon_{i,t}^\theta \\ \Delta\omega_{i,t}^q &= b_i + b_{p,s,t} + \gamma^q \cdot \phi_{i,t} + \epsilon_{i,t}^q\end{aligned}$$

therefore

$$\frac{1}{\sigma} \cdot \Delta\theta_{i,t} + \frac{\sigma-1}{\sigma} \cdot \Delta\omega_{i,t}^q = \psi_i + \psi_{p,s,t} + \left( \frac{1}{\sigma} \cdot \gamma^\theta + \frac{\sigma-1}{\sigma} \cdot \gamma^q \right) \cdot \phi_{i,t} + \eta_{i,t}$$

under the exclusion restrictions of uncorrelation of  $\phi_{i,t}$  with respect to  $\eta_{i,t} = \epsilon_{i,t}^q + \epsilon_{i,t}^\theta$  (conditional on fixed effects), then, the parameter recovered by estimating the main equation (9) is

$$\gamma = \left( \frac{1}{\sigma} \cdot \gamma^\theta + \frac{\sigma-1}{\sigma} \cdot \gamma^q \right)$$

The derivative of  $\gamma$  with respect to  $\sigma$  is

$$\frac{\partial\gamma}{\partial\sigma} = \frac{1}{\sigma^2} \cdot (\gamma^q - \gamma^\theta)$$

implying

$$\frac{\partial\gamma}{\partial\sigma} > 0 \Leftrightarrow \gamma^q > \gamma^\theta$$

That is, the effect of credit supply shocks on revenue productivity is increasing in the elasticity of demand if the effect of credit supply shocks on technical efficiency is stronger than the effect on market appeal, and viceversa. Consequently, under the (strong) assumption that  $\gamma^q$  and  $\gamma^\theta$  are both constant across all industries, we can use sectoral variation in  $\sigma$  to test whether  $\gamma^q > \gamma^\theta$  or viceversa. To do so, we estimate the heterogeneity model:

$$\Delta\omega_{i,t} = \psi_i + \psi_{p,s,t} + (\gamma + \gamma^\sigma \cdot HE_s) \cdot \phi_{i,t} + \eta_{i,t} \quad (31)$$

where  $HE_s$  is a dummy equal to one iff industry  $s$  has an elasticity of demand above the median. As in appendix B.1, we follow Pozzi & Schivardi (2016) and estimate elasticity of demand from INVIND self-reported elasticities.<sup>53</sup>

Results are reported in Table A.4. The effect of credit supply shocks on revenues productivity is significantly stronger in industries with higher elasticity of demand. Demand elasticity can be correlated with many technological or economic factors. Therefore, this empirical finding should be interpreted with extreme caution. Nonetheless, Table A.4 suggests that the effect of credit supply

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<sup>53</sup>We assume each two digit industry has a single elasticity and take the median value among all the responses. We drop responses implying negative values of  $\sigma$ . We use both 2007 and 1996 survey waves.

on technical efficiency is likely larger than the effect on market appeal, at least in manufacturing.

## C.2 Unobservable Selection and Coefficient Stability (Oster, 2016)

Oster (2016) develops a framework to evaluate coefficient stability and changes in  $R^2$  when including observable controls. This framework, which builds on work by Altonji *et al.* (2005), is tailored to study how much the coefficient of a linear regression is robust to the presence of unobservable variables. It formalizes a commonly used intuitive approach: if the researcher includes relevant controls in a linear regression and the coefficient associated with the variable of interest does not vary, then it is “unlikely” that omitted variables are significantly affecting the results.

In order to implement this approach in our setting, let us define  $R_{un}$  and  $\gamma_{un}$  as the R-squared and the coefficient of interest of the unrestricted regression (full set of fixed effects) and  $R_{con}$  and  $\gamma_{con}$  as their restricted counterpart (from regression with only province and sector and year fixed effect, but no interaction). They can be found in columns (1) and (5) of table 4 (for the revenue Cobb-Douglas case), see section 5.1. The formula at the end of section 3.2 of the 2016 working paper version of Oster (2016) defines as “approximated bias adjusted treatment effect” the coefficient

$$\gamma(\delta, R_{max}) = \gamma_{un} - \delta \cdot (\gamma_{con} - \gamma_{un}) \cdot \frac{R_{max} - R_{un}}{R_{un} - R_{con}}$$

where  $\delta, R_{max}$  are two parameters to be chosen by the researcher.  $R_{max}$  is the maximum R-squared that a regression including all the observable and unobservable variables can attain. We set  $R_{max}$  equal to 1, that is the most conservative value.  $\delta$  is a parameter governing the relative importance of unobservable variables with respect to the observable controls. It is common to set  $\delta = 1$ , that is, to assume that observable and unobservable have the same correlation with the variable of interest. However, we choose  $\delta = 2$  in order to be very conservative. As suggested in section 3.4 of Oster (2016), we build bounding set for  $\gamma$  using  $\gamma_{uc}$  and  $\gamma(\delta = 2, R_{max} = 1)$  as extreme points. Results, which are presented in Table A.6, show that these bounding sets never contain 0. Therefore, our results on the effect of credit shocks on productivity growth (section 5) are “robust” to the presence of unobservable shocks.

## C.3 Measurement Error

Most of the production function literature assume that inputs are measured without error.<sup>54</sup> However, the complete absence of any measurement error is an utopia. Therefore, the reader might be concerned that the mismeasurement of inputs with respect to output is an important driver of our results. Section 5.1 deals with robustness of the findings with respect to misspecification of the production functions. A further concern is that that we find a residual effect of the credit supply shocks on productivity because we are not able to fully control for inputs. In fact, we can re-write equation (9) as:

$$\Delta y_{i,t} = \psi_i + \psi_{p,s,t} + \Delta f(k_{i,t}, l_{i,t}, m_{i,t}, \beta) + \gamma \cdot \phi_{i,t} + \eta_{i,t}$$

where the  $\beta$  parameters are computed on a first stage. Given that  $(k_{i,t}, l_{i,t}, m_{i,t})$  are correlated, measurement error in the inputs might lead inconsistent estimates for  $\gamma$ . Table 2, where inputs are on

<sup>54</sup>There are few notable exceptions, such as Collard-Wexler & De Loecker (2016).

the left hand side, mitigates these concern. Measurement error on the dependent variable<sup>55</sup> worsen estimates precision, but does not lead to inconsistent estimates. Therefore, the finding that output respond more than inputs (except capital), which is the statistical finding informing the productivity results, cannot be generated by classical mismeasurment.

The combination of factor hoarding and adjustment costs might generate more pernicious forms of misurement errors and create spurious correlation between credit supply and productivity. For instance, as a consequence of a tightening in the credit constraint, a firm might immediately scale down production by acquiring less intermediate inputs and, let’s say, disinvest part of the capital goods. However, because of employment protection legislation, firing workers might take some time even though these are factually out of production. Therefore, the researcher would observe a wagebill or headcount overestimating the real workforce. Similarly, we observe only capital stock and not its utilization. If using capital is costly, for instance because of endogenous deterioration, firms might respond to negative credit supply shocks partially by changing utilization rate rather than investments. While these concerns are well grounded, and our empirical analysis would be more complete if we could observed capital utilization and hours worked, they cannot be a main driver of our results. In fact, these stories are based on delayed adjustments and they could create short-term productivity loss from negative shocks. Conversely, section 5.3 shows that effect of credit supply shocks last for, at least, few years.

## C.4 Small vs large lenders

Is the effect of credit supply on productivity driven by small or large banks? To answer to this question, we compute, for each borrower firms, the average size of its lenders:

$$LenderSize_{i,t} = \sum_b Assets_{b,t-1} \cdot \frac{C_{b,i,t-1}}{\sum_{b'} C_{b',i,t-1}}$$

where  $Assets_{b,t-1}$  is the total asset size of bank  $b$  and  $C_{b,i,t-1}$  is the credit granted by  $b$  to firm  $i$ . Then, we re-estimate the main equation (9) excluding firms in the top (“large banks”) or bottom (“small banks”) quartile of the (year-specific) distribution of  $LenderSize_{i,t}$ . Results are reported in Table A.7. The effect of credit supply on productivity is robust to the exclusion of firms borrowing from large or small lenders. This finding implies that our main results are not driven by banks of a specific size.

## D Additional Materials on Interbank Shock

### D.1 Credit Granted and Credit Supply

This section investigates whether the exposure to the interbank market was a significant negative credit supply shock, as we argue in section 6. For each firm  $i$  active in industry  $s$  and province  $p$  over the years  $t \in [2007, 2009]$ , we estimate the equations:

$$\begin{aligned} \Delta credit_{i,t} &= \psi_{p,s,t} + \gamma \cdot INTBK_{i,2006} + \eta_{i,t} \\ \phi_{i,t} &= \psi_{p,s,t} + \gamma \cdot INTBK_{i,2006} + \eta_{i,t} \end{aligned}$$

---

<sup>55</sup>The difference of two classical measurement errors is still a classical measurement error.



Results are shown in Table A.5, which documents that firms more exposed to the collapse of the interbank market decrease more the credit received with respect to others operating in the same industry and location. An increase of dependence from the interbank market of 1%, lead to a decrease of the growth rate of credit granted between a quarter and a fifth of a percentage point, see columns (2) and (4). Furthermore, columns (1) and (3) show that the measure of credit supply shocks  $\phi_{i,t}$  does respond negatively to the interbank shocks.

## D.2 Interbank and Sensitivity to Business Cycle

A further concern is that, although firms more exposed to interbank market had equal average productivity growth before the credit crunch, they were more sensitive to business cycle fluctuation and, therefore, they suffered more during the recession following the financial turmoil. For each firm in the sample, we estimate its sensitivity to business cycle from equation:

$$\Delta y_{i,t} = \alpha_i + \beta_i \cdot GDPgr_t + \epsilon_{i,t}$$

where  $GDPgr_t$  is the growth rate of Italian GDP in year  $t$  and  $y_{i,t}$  is one of two outcomes: (logs) value added or (logs) revenues. The model is estimated using all available years before 2006. Then, we study the correlation between the three measures of sensitivity to fluctuation and the interbank exposure in 2006. Table A.8 shows that firms more exposed to the collapse of the interbank market were not significantly more sensitive to downturns before 2006.

## D.3 Interbank and Patents

Did the credit crunch affected innovation in Italy? In section 7 we show that credit supply shocks have a significant impact on firm-level innovation. We also provide some descriptive statistics suggesting that the 2007-08 credit crunch severely damaged patenting activities. Here, we investigate whether firms hit harder by the credit crunch patented less afterward. Thus, we estimate the linear model:

$$PatentGrowth_i = \psi_p + \psi_s + \gamma INTBK_{i,2006} + \eta_i$$

where

$$PatentGrowth_i = 2 \cdot \frac{Patent_{post,i} - Patent_{pre,i}}{Patent_{post,i} + Patent_{pre,i}}$$

and  $Patent_{post,i}$  and  $Patent_{pre,i}$  is the total number of patent applications done by firm  $i$  in the post and in the pre periods. The interbank market started collapsing in the late 2007, although general economic activity started declining only afterward. We thus use 2007 as the start of the “post” period and 2006 as the end of the “pre” period. We use either 2009 (the very worst year for Italian economy) or 2010 (a short period of recovery) as the end of the “post” period. We use either 2001 or 2002 as start of the “pre” period. Notice that the LHS variable can be computed only for firms which patent at least once either in the “pre” or “post” period (or both). Results are presented in Table A.9: firms connected to lenders more exposed to the collapse of the interbank market contracted more their patent applications. The boundaries of the periods do not matter.

## E Additional Materials on Mechanisms

### E.1 Notes on R&D and Export

In section 7 we estimate equations:

$$\begin{aligned}Pr(R\&D_{i,t} = 1) &= \psi_i + \psi_t + \gamma \cdot \phi_{i,t} + \eta_{i,t} \\Pr(Export_{i,t} = 1) &= \psi_i + \psi_t + \gamma \cdot \phi_{i,t} + \eta_{i,t}\end{aligned}$$

If we include the full set of province×industry×year fixed effects to control for local business cycle and industry-specific shocks the sample size halves (because of singletons). While the estimated coefficients stay in the same ballpark, the standard errors raise enough that we are not able to reject the null of no effect on R&D at the conventional level.

# Additional Figures

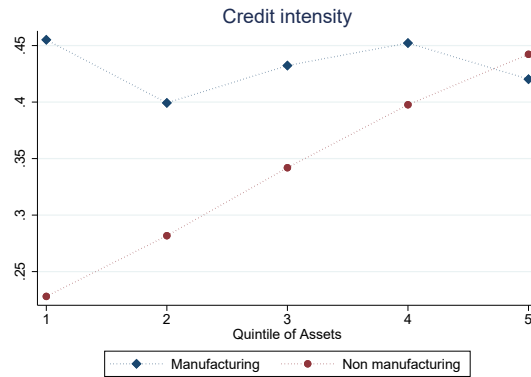


Figure A.1: Credit intensity per quintile of asset size. Credit intensity is the ratio of credit granted over net revenues and it is winsorized at top 2%

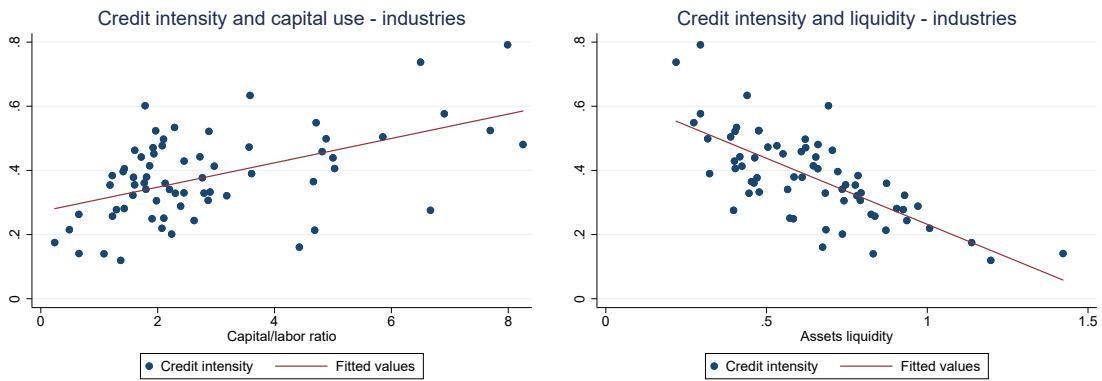


Figure A.2: Industry (2-digits) average credit intensity and capital to labor ratio (left panel) or liquidity (right panel). Credit intensity is the ratio of credit granted over net revenues and it is winsorized at top 2% before taking averages. Capital to labor ratio is the ratio of capital stock over total wagebill. Liquidity is the ratio of liquid assets over book value of capital.

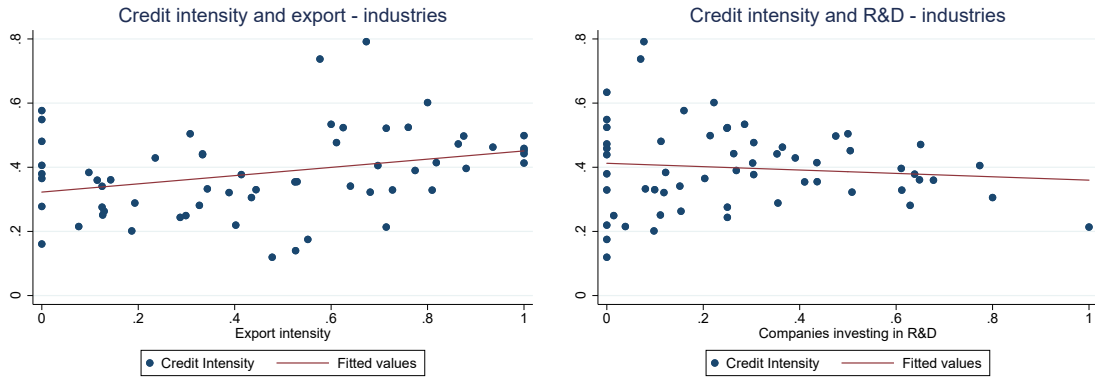


Figure A.3: Industry (2-digits) average credit intensity and share of companies engaging in export (left panel) or R&D (right panel). Credit intensity is the ratio of credit granted over net revenues and it is winsorized at top 2% before taking averages. Data on export and R&D are taken from INVIND survey (see sections 2 and 7) and represent noisy estimate of the effective export and R&D intensity. The slope of the fitted line in left panel is significantly larger from zero, while the one in the right panel is statistically indistinguishable from zero.

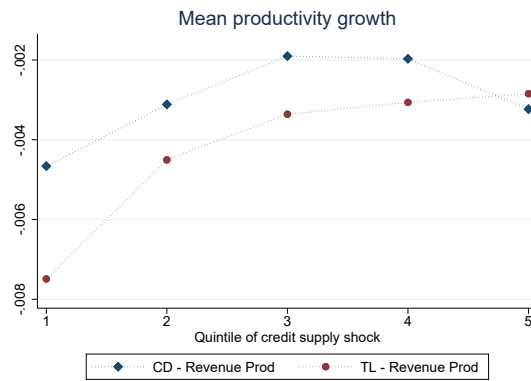


Figure A.4: Average productivity growth per quintile of credit supply shock

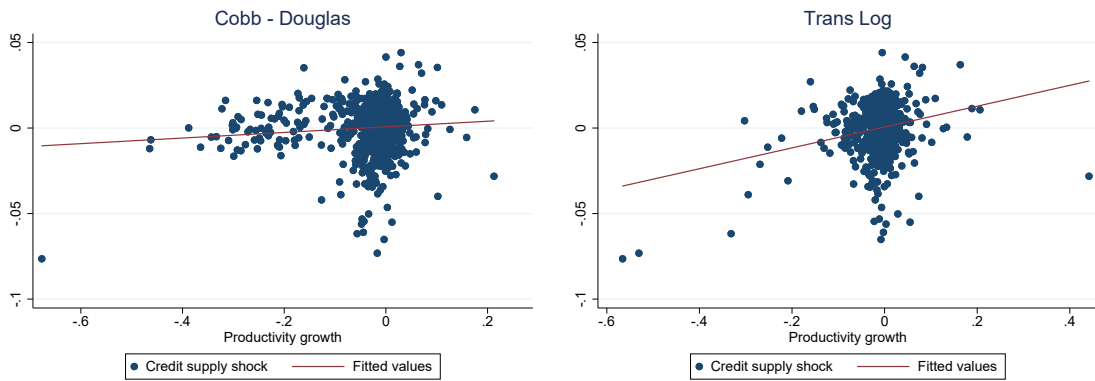


Figure A.5: Industry×year average revenue productivity growth and credit supply shocks. Fitted lines in both panels have a slope significantly larger than zero (1% confidence). We drop two observations with extremely negative value of average the credit supply shock

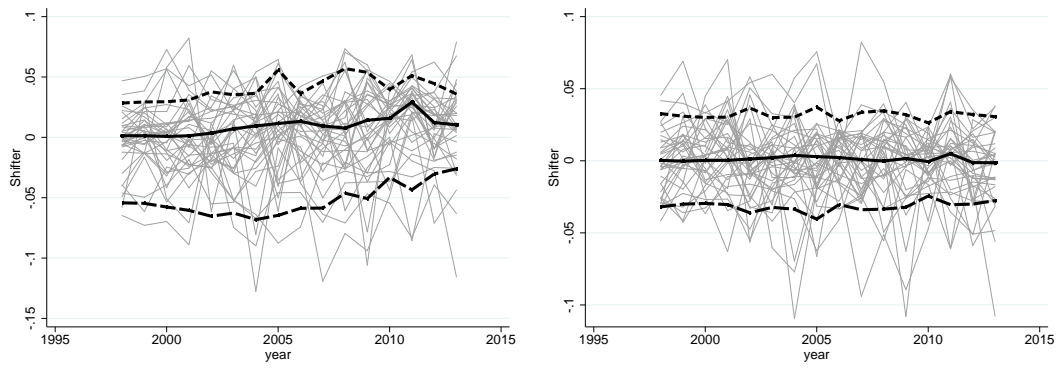


Figure A.6: Figures display evolution of Credit Supply Shock experienced by a 1.5% random sample. Right panel shows residualized values after taking out FEs

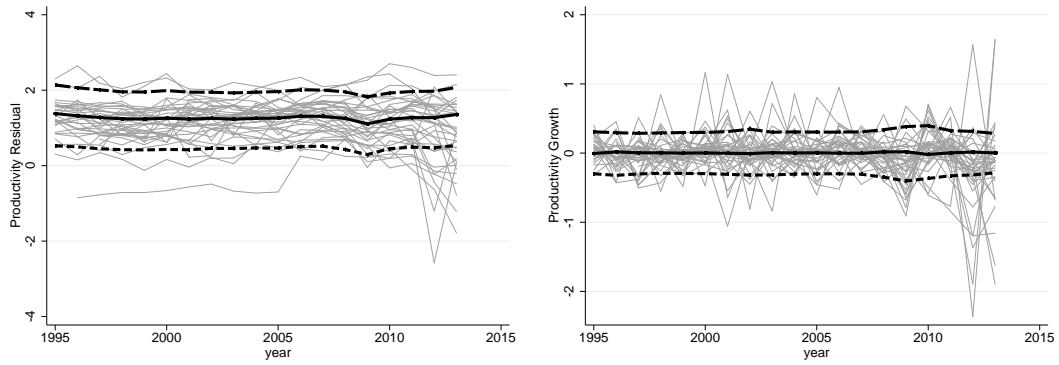


Figure A.7: Figures display evolution of Productivity (Cobb-Douglas, Value Added) for 1.5% random sample. Right panel shows residualized values after taking out FEs

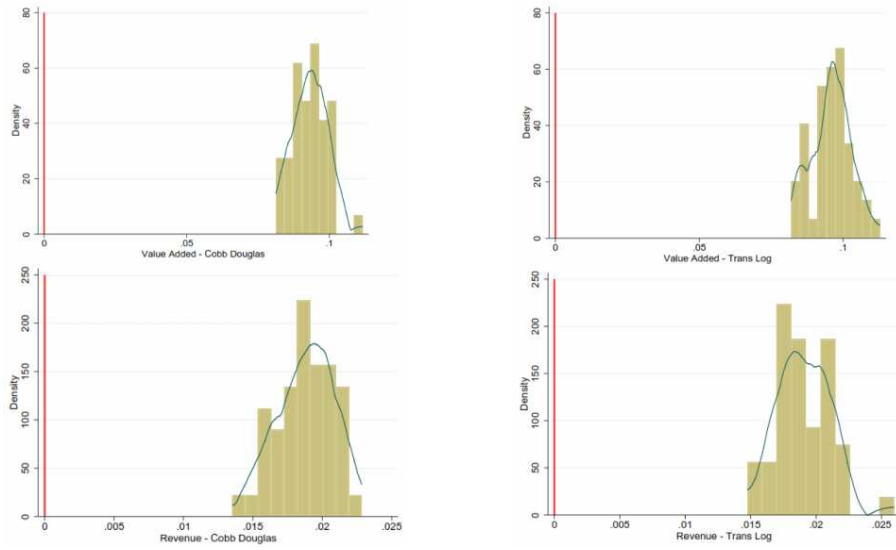


Figure A.8: Distribution of  $\gamma$  from equation  $\Delta\omega_{i,t} = \psi_i + \psi_{p,s,t} + \gamma \cdot \phi_{i,t} + \eta_{i,t}$ . See section 5 for details. Distribution is computed from 50 (firm-level) bootstrapped sample. Industry level production function and firm level productivity growth is re-estimated for each bootstrapped sample. Estimates are all above zero (red vertical line) for all samples.

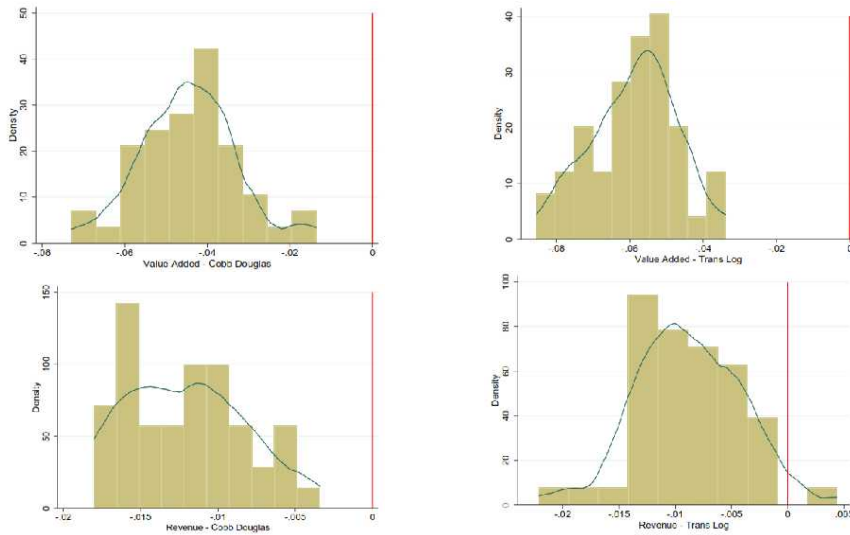


Figure A.9: Distribution of  $\gamma$  from equation  $\Delta\omega_{i,t} = \psi_{p,s,t} + \gamma \cdot ITBK_{i,2006} + \eta_{i,t}$ . See section 6 for details. Distribution is computed from 50 (firm-level) bootstrapped sample. Industry level production function and firm level productivity growth is re-estimated for each bootstrapped sample. Estimates are all below zero for all samples except one (one of the estimates related to the revenue- trans log productivity case).

## Additional Tables

Table A.1: Credit Supply Shock and Loan Applications to New Lenders

VARIABLES	N. of PI (1)	N. of PI   N>0 (2)	Pr(N. of PI >0) (3)
All Industries			
$\phi_{i,t}$	-0.537*** (0.0796)	-0.458*** (0.113)	-0.0780*** (0.0173)
Observations	656,960	456,888	656,960
$R^2$	0.500	0.477	0.348
Manufacturing			
$\phi_{i,t}$	-0.424*** (0.113)	-0.355** (0.161)	-0.0583** (0.0242)
Observations	347,990	246,453	347,990
$R^2$	0.491	0.462	0.349

*Notes:* One observation is one firm for one year between 1998 and 2013 (unbalanced panel). Firm FEs and province $\times$ industry $\times$ year FEs are included. Singleton are dropped. The RHS variable  $\phi_{i,t}$  represents idiosyncratic shock to firm credit supply, and its construction is detailed in section 3.1. A 1% increase in  $\phi_{i,t}$  is the supply shock needed to increase the credit granted to firm  $i$  by 1%. The LHS variables are built from the number of banks which request information about firm  $i$  in year  $t$  and they proxy for the number of of firm  $i$ ' applications with previously unconnected lenders. Column (1) conflates both intensive and extensive margin, while (2) is a linear probability model for the probability of making any application. Column (3) considers the extensive margin only. Standard errors (in parentheses) are clustered at firm level. \*\*\*  $p < 0.01$ , \*\*  $p < 0.05$ , \*  $p < 0.1$

Table A.2: Descriptive Statistics - Cobb Douglas Parameters

VARIABLE	Mean	Std. Dev.	Mean (Quantity)	Std. Dev. (Quantity)
<b>All Industries</b>				
Value Added				
$\beta_l$	.64	.16	.85	.25
$\beta_k$	.17	.14	.22	.19
Revenues				
$\beta_l$	.14	.17	.18	.22
$\beta_k$	.06	.06	.08	.08
$\beta_m$	.81	.11	1.08	.25
<b>Manufacturing</b>				
Value Added				
$\beta_l$	.62	.15	.81	.22
$\beta_k$	.19	.16	.24	.21
Revenues				
$\beta_l$	.13	.05	.17	.07
$\beta_k$	.04	.02	.05	.03
$\beta_m$	.81	.1	1.06	.12

Notes:  $\beta_x$  is the estimated elasticity of output with respect to input  $x$ . Estimation of the parameters is performed at sector level, details are provided in section 3.2. Standard deviation represent sectoral variations and not estimation error. “Quantity” parameters are calculated by multiplying the estimate of sales-generating production function by the correction term  $\frac{\sigma}{\sigma-1}$ , where  $\sigma$  is the elasticity of demand. The correction is exact if firms are monopolistic competitors, see De Loecker (2011).  $\sigma$  is estimated from self-reported elasticity of demand, as in Pozzi & Schivardi (2016).

Table A.3: Credit Supply Shock and Productivity Growth (alternative model)

VARIABLES (in delta Log)	Productivity (1)	Productivity (2)	Productivity (3)	Productivity (4)
$\phi_{i,t}$	0.117*** (0.0178)	0.0347*** (0.00589)	0.100*** (0.0145)	0.0231*** (0.00412)
Observations	347,990	347,990	656,960	656,960
$R^2$	0.194	0.146	0.182	0.172
Functional Form	Cobb-Douglas	Cobb-Douglas	Cobb-Douglas	Cobb-Douglas
Output Measure	Value Added	Net Revenues	Value Added	Net Revenues
Sample	Manufacturing	Manufacturing	All	All

Notes: One observation is one firm for one year (panel). Firm FEs and province×industry×year FEs are included. The RHS variable  $\phi_{i,t}$  represents idiosyncratic shock to firm credit supply, and its construction is detailed in section A.1. Productivity is estimated following the model in section B.2. Standard errors (in parentheses) are clustered at firm level. \*\*\* p<0.01, \*\* p<0.05, \* p<0.1



Table A.4: Credit Supply Shock and Productivity Growth - Heterogeneity by Demand Elasticity

VARIABLES (in delta Log)	Productivity (1)	Productivity (2)	Productivity (3)	Productivity (4)
All Industries				
$\phi_{i,t}$	0.0854*** (0.0169)	0.101*** (0.0175)	0.0146*** (0.00445)	0.0199*** (0.00470)
$\phi_{i,t} \cdot HS_s$	0.0399 (0.0321)	0.0345 (0.0342)	0.0156 (0.0134)	0.0221* (0.0121)
Observations	649,662	649,662	649,662	649,662
R-squared	0.170	0.182	0.176	0.188
Manufacturing				
$\phi_{i,t}$	0.0774*** (0.0255)	0.0906*** (0.0260)	0.0201** (0.00823)	0.0171* (0.00938)
$\phi_{i,t} \cdot HS_s$	0.0787** (0.0354)	0.0623* (0.0374)	0.0212* (0.0121)	0.0316** (0.0132)
Observations	347,990	347,990	347,990	347,990
R-squared	0.186	0.200	0.144	0.180

*Notes:*  $\Delta\omega_{i,t} = \psi_i + \psi_{s,t,p} + (\gamma + \gamma^\sigma \cdot HE_s) \cdot \phi_{i,t} + \eta_{i,t}$ . One observation is one firm for one year between 1998 and 2013 (unbalanced panel). Firm FEs and province $\times$ industry $\times$ year FEs are included. Singleton are dropped. The RHS variable  $\phi_{i,t}$  represents idiosyncratic shock to firm credit supply, and its construction is detailed in section 3.1. A 1% increase in  $\phi_{i,t}$  is the supply shock needed to increase the credit granted to firm  $i$  by 1%. The LHS variable is the first difference of Hicks-neutral productivity residual:  $\Delta\omega_{i,t} = \Delta y_{i,t} - \Delta f(x_{i,t}, \beta)$  where  $y$  is log of net revenues (or log of value added) and  $x$  is a set of inputs. Capital stock, labor, and (for the revenue case only) intermediate inputs are included in  $x$ .  $f(\cdot, \beta)$  is either a first (Cobb-Douglas) or second (Trans-Log) order polynomial in log inputs. Estimation of parameters  $\beta$  is described in section 3.2.  $HE_s$  is a dummy variable equal to one iff firm  $i$  is in an industry with elasticity of demand above the sample median. Sectoral elasticities are calculated from INVIND (1996 and 2007 waves) self-reported elasticity of demand (we take median of all responses within a 2-digits industry). Standard errors (in parentheses) are clustered at firm level. \*\*\*  $p < 0.01$ , \*\*  $p < 0.05$ , \*  $p < 0.1$

Table A.5: Exposure to Interbank Market, Credit Supply Shocks and Credit Granted

VARIABLES	Credit Supply	Credit Granted (delta log)	Credit Supply	Credit Granted (delta log)
	(1)	(2)	(3)	(4)
$ITBK_{i,2006}$	-0.137*** (0.00624)	-0.203*** (0.0383)	-0.160*** (0.00900)	-0.253*** (0.0509)
Observations	110070	108267	57986	57349
$R^2$	0.187	0.093	0.194	0.089

Notes: Models are  $\phi_{i,t} = \psi_{s,t,p} + \gamma \cdot ITBK_{i,2006} + \eta_{i,t}$  and  $\Delta cred_{i,t} = \psi_{s,t,p} + \gamma \cdot ITBK_{i,2006} + \eta_{i,t}$ . One observation is one firm for one year between 2007 and 2009 (unbalanced panel). Province  $\times$  industry  $\times$  year FEs are included. Singleton are dropped. The RHS variable  $ITBK_{i,2006}$  is the weighted average of firm's  $i$  lenders' liability on the interbank market over assets in 2006. The first LHS variable is the credit supply shocks  $\phi_{i,t}$ , construction is detailed in section 3.2. The second LHS is the first difference of the log of the credit granted to firm  $i$  by all financial intermediaries at the end of year  $t$ . Standard errors (in parentheses) are clustered at firm level. \*\*\*  $p < 0.01$ , \*\*  $p < 0.05$ , \*  $p < 0.1$ .

Table A.6: Credit Supply Shocks and Productivity Growth - bounding sets (Oster, 2016)

VARIABLES (in delta Log)	Productivity Cobb-Douglas Value Added	Productivity Trans-Log Value Added	Productivity Cobb-Douglas Net Revenues	Productivity Trans-Log Net Revenues
	(1)	(2)	(3)	(4)
All industries				
$\phi_{i,t}$	[0.043 ; 0.095]	[0.057 ; 0.11]	[0.019 ; 0.066]	[0.026 ; 0.071]
Observations	656,960	656,960	656,960	656,960
Manufacturing				
$\phi_{i,t}$	[0.069 ; 0.115]	[0.097 ; 0.121]	[0.014 ; 0.030]	[0.032 ; 0.126]
Observations	347,990	347,990	347,990	347,990

Notes:  $\Delta \omega_{i,t} = \psi_i + \psi_{s,t,p} + \gamma \cdot \phi_{i,t} + \eta_{i,t}$ . One observation is one firm for one year between 1998 and 2013 (unbalanced panel). Bounding sets are built following Oster (2016), see appendix C.2 for details.

Table A.7: Credit Supply Shock and Productivity Growth - Exclusion of small or large lenders

VARIABLES (in delta Log)	Productivity	Productivity	Productivity	Productivity
Exclude	Small Banks	Large Banks	Small Banks	Large Banks
	(1)	(2)	(3)	(4)
	Manufacturing		All Industries	
$\phi_{i,t}$	0.0176** (0.00712)	0.0335*** (0.00738)	0.0174*** (0.00544)	0.0188*** (0.00552)
Observations	261,375	260,308	492,427	489,502
R-squared	0.164	0.172	0.197	0.211

Notes:  $\Delta\omega_{i,t} = \psi_i + \psi_{s,t,p} + \eta_{i,t}$ . One observation is one firm for one year between 1998 and 2013 (unbalanced panel). Firm FEs and province $\times$ industry $\times$ year FEs are included. Singleton are dropped. Appendix C.4 details the classification of firms is “large bank” and “small bank” groups. The RHS variable  $\phi_{i,t}$  represents idiosyncratic shock to firm credit supply, and its construction is detailed in section 3.1. A 1% increase in  $\phi_{i,t}$  is the supply shock needed to increase the credit granted to firm  $i$  by 1%. The LHS variable is the first difference of productivity residual:  $\Delta\omega_{i,t} = \Delta y_{i,t} - \beta_k \cdot \Delta k_{i,t} - \beta_l \cdot \Delta l_{i,t} - \beta_m \cdot \Delta m_{i,t}$  where  $y$  is log of net revenues,  $k$  is log of capital stock,  $l$  is labor (measured by log of wagebill) and  $m$  is log of intermediate inputs. Estimation of parameters  $\beta$  is described in section 3.2. Standard errors (in parentheses) are clustered at firm level. \*\*\* p<0.01, \*\* p<0.05, \* p<0.1

Table A.8: Exposure to Interbank in 2006 and pre-2006 sensitivity to business cycle fluctuations.

VARIABLES	Sensitivity to business cycle	Sensitivity to business cycle	Sensitivity to business cycle	Sensitivity to business cycle
	(1)	(2)	(3)	(4)
	All Industries		Manufacturing	
$ITBK_{i,2006}$	0.0767 (0.108)	0.0604 (0.0377)	0.195 (0.225)	-0.0292 (0.0464)
Observations	34,004	34,004	17,759	17,759
$R^2$	0.030	0.104	0.026	0.091
Output measure	Value Added (delta log)	Revenues (delta log)	Value Added (delta log)	Revenues (delta log)

Notes: Model is  $\beta_i = \psi + \gamma \cdot ITBK_{i,2006} + \eta_i$ . The RHS variable  $ITBK_{i,2006}$  is the weighted average of firm’s  $i$  lenders’ liability on the interbank market over assets in 2006. The LHS variable  $\beta_i$  is the estimated parameter from the model  $\Delta y_{i,t} = \alpha_i + \beta_i \cdot GDPgr_t + \epsilon_{i,t}$  where  $\Delta y_{i,t}$  is the delta log of revenues or value added produced by firm  $i$  in year  $t < 2006$  and  $GDPgr_t$  is the growth rate of Italy’s GDP.

Table A.9: Exposure to Interbank in 2006 and growth of patent applications

VARIABLES	Patent Growth	Patent Growth	Patent Growth
“Pre”	[2003;2006]	[2002;2006]	[2002;2006]
“Post”	[2007;2009]	[2007;2010]	[2007;2009]
	(1)	(2)	(3)
	-1.463*	-1.703**	-1.738**
	(0.783)	(0.834)	(0.747)
Observations	2,903	3,217	3,040
R-squared	0.058	0.054	0.057

Notes: Results of estimating model:

$$PatentGrowth_i = \psi_p + \psi_s + \gamma INTBK_{i,2006} + \eta_i$$

where  $PatentGrowth_i = 2 \cdot \frac{Patent_{post,i} - Patent_{pre,i}}{Patent_{post,i} + Patent_{pre,i}}$ .  $Patent_{post,i}$  ( $Patent_{pre,i}$ ) is the total number of patent applications made by firm  $i$  in the “pre” (“post”) period. The three columns are different because of the definition of “pre” and “post” periods.

Table A.10: Bank characteristics, M&amp;A and credit supply shocks

VARIABLES	Credit Supply	Credit Supply	Credit Supply	Credit supply
(delta log)				
Years	[2007,2009]	[2007,2009]	[2010,2013]	All years
Unit of observation	Bank	Bank	Bank	Firm
	(1)	(2)	(3)	(4)
$\left(\frac{loans}{deposits}\right)_{b,2007}$	-0.0189*** (0.00533)			
$\left(\frac{capital}{RWA}\right)_{b,2007}$		0.00133** (0.000588)		
$\Delta \left(\frac{sovereign}{Assets}\right)_{b,t}$			-0.155*** (0.0400)	
$M\&A_{i,t}$				-0.0117*** (0.000657)
Observations	1,635	1,635	2,034	652,692
R-squared	0.086	0.076	0.010	0.521

Notes: Columns (1)-(2): model is  $\phi_{b,t} = \psi_t + \gamma \cdot X_{b,\tau} + \eta_{b,t}$  where  $\phi_{b,t}$  is a bank-level measure of changes in credit supply (see section 3.1 for details) and  $X_{b,2007}$  is a bank-level characteristic at time 2007. Column (3): model is  $\phi_{b,t} = \psi_t + \gamma \cdot \Delta \left(\frac{sovereign}{Assets}\right)_{b,t} + \eta_{b,t}$  where  $\left(\frac{sovereign}{Assets}\right)_{b,t}$  is the share of sovereign debt on bank  $b$  assets. Column (4): model is  $\phi_{i,t} = \psi_i + \psi_{p,s,t} + \gamma \cdot MA_{i,t} + \eta_{i,t}$  where  $\phi_{i,t}$  is a firm-level measure of shock to credit supply (see section 3.1 for details) and  $MA_{i,t}$  is the share of firm  $i$  previous period lenders which undergo a merger and acquisition (as a target).

Table A.11: Credit Supply Shocks and Productivity Growth: Robustness - Cobb-Douglas Value Added Productivity

VARIABLES (delta Logs)	Productivity	Productivity	Productivity	Productivity	Productivity	Productivity	Productivity	Productivity	Productivity
Model	Baseline	Firm Controls	Important Borrowers	Pooled Estimator	Alternative FEs structure	Match Controls	Split Sample	4 Digits Sector	Endogenous Exit
	(1)	(2)	(3)	(4)	(5)	(6)	(7)	(8)	(9)
All Industries									
$\phi_{i,t}$	0.0946*** (0.0155)	0.106*** (0.0175)	0.0865*** (0.0172)	0.0436*** (0.00963)	0.0968*** (0.0151)	0.101*** (0.0188)	0.0932*** (0.0164)	0.0988*** (0.0183)	0.0898*** (0.0156)
Observations	656,960	483,665	521,741	656,960	656,960	656,960	656,960	587,873	656,960
R-squared	0.172	0.191	0.185	0.021	0.104	0.172	0.172	0.267	0.175
Manufacturing									
$\phi_{i,t}$	0.115*** (0.0178)	0.122*** (0.0211)	0.126*** (0.0196)	0.0405*** (0.0120)	0.116*** (0.0180)	0.114*** (0.0216)	0.114*** (0.0188)	0.127*** (0.0208)	0.111*** (0.0180)
Observations	347,990	262,308	280,346	347,990	347,990	347,990	347,990	309,887	347,990
R-squared	0.186	0.209	0.198	0.032	0.110	0.186	0.186	0.278	0.191

*Notes:* Model is  $\Delta\omega_{i,t} = \psi_i + \psi_{s,t,p} + \gamma \cdot \phi_{i,t} + \eta_{i,t}$ . One observation is one firm for one year between 1998 and 2013 (unbalanced panel). Firm FEs and province $\times$ industry $\times$ year FEs are included. Singleton are dropped. The RHS variable  $\phi_{i,t}$  represents idiosyncratic shock to firm credit supply, and its construction is detailed in section 3.1. A 1% increase in  $\phi_{i,t}$  is the supply shock needed to increase the credit granted to firm  $i$  by 1%. The LHS variable is the first difference of productivity residual:  $\Delta\omega_{i,t} = \Delta va_{i,t} - \beta_k \cdot \Delta k_{i,t} - \beta_l \cdot \Delta l_{i,t}$  where  $va$  is log of net value added,  $k$  is log of capital stock and  $l$  is labor (measured by log of wagebill). Estimation of parameters  $\beta$  is described in section 3.2. Column (2) add a set of lagged controls to baseline specification: polynomial in size (assets) and the ratios of value added, liquidity, cash flow and bank debt to assets. It excludes observation with missing or zero values for any control variable. Column (3) excludes any firm that, at any point in time, received more than 1% of the credit by any financial intermediary. Column (4) use pooled estimator (rather than “within”) by dropping firm FEs. Column (5) includes firm FEs, province FEs, year FEs and industry FEs, but do not include province $\times$ year $\times$ industry FEs. Column (6) uses an alternative measure of credit supply shocks which control for match-specific covariates, see section 3.1. Column (7) uses, as an instrument, an alternative credit supply shocks estimated with a “split sample” procedure, in order to control for finite sample biases. Column (8) uses a 4-digits (rather than 2) industry definition both for the estimation of productivity parameters and for the FEs structure. It contains less observations because of the singleton dropping. Column (9) estimate productivity allowing for endogenous firm exit, as in [Olley & Pakes \(1996\)](#). Standard errors, in parentheses, are (two-way) clustered at firm and main-bank $\times$ year level. \*\*\* p<0.01, \*\* p<0.05, \* p<0.1

Table A.12: Credit Supply Shocks and Productivity Growth: Robustness - Translog Revenue Productivity

VARIABLES (delta Logs)	Productivity	Productivity	Productivity	Productivity	Productivity	Productivity	Productivity	Productivity	Productivity
Model	Baseline	Firm Controls	Important Borrowers	Pooled Estimator	Alternative FEs structure	Match Controls	Split Sample	4 Digits Sector	Endogenous Exit
	(1)	(2)	(3)	(4)	(5)	(6)	(7)	(8)	(9)
All Industries									
$\phi_{i,t}$	0.0259*** (0.00491)	0.0268*** (0.00563)	0.0246*** (0.00571)	0.0210*** (0.00367)	0.0244*** (0.00492)	0.0297*** (0.00621)	0.0261*** (0.00521)	0.0274*** (0.00573)	0.0242*** (0.00452)
Observations	656,960	483,665	521,741	656,960	656,960	656,960	656,960	586,012	656,960
R-squared	0.195	0.202	0.207	0.007	0.100	0.195	0.195	0.267	0.182
Manufacturing									
$\phi_{i,t}$	0.0323*** (0.00649)	0.0363*** (0.00710)	0.0343*** (0.00765)	0.0304*** (0.00483)	0.0274*** (0.00659)	0.0335*** (0.00809)	0.0315*** (0.00693)	0.0361*** (0.00820)	0.0299*** (0.00644)
Observations	347,990	262,308	280,346	347,990	347,990	347,990	347,990	309,252	347,990
R-squared	0.180	0.191	0.185	0.012	0.093	0.180	0.180	0.277	0.164

*Notes:* Model is  $\Delta\omega_{i,t} = \psi_i + \psi_{s,t,p} + \gamma \cdot \phi_{i,t} + \eta_{i,t}$ . One observation is one firm for one year between 1998 and 2013 (unbalanced panel). Firm FEs and province×industry×year FEs are included. Singleton are dropped. The RHS variable  $\phi_{i,t}$  represents idiosyncratic shock to firm credit supply, and its construction is detailed in section 3.1. A 1% increase in  $\phi_{i,t}$  is the supply shock needed to increase the credit granted to firm  $i$  by 1%. The LHS variable is the first difference of productivity residual:  $\Delta\omega_{i,t} = \Delta y_{i,t} - f(k_{i,t}, l_{i,t}, m_{i,t}, \beta)$  where  $y$  is log of net revenues,  $k$  is log of capital stock,  $l$  is labor (measured by log of wagebill),  $m$  is log of intermediate inputs, and  $f(\cdot, \beta)$  is a second order polynomial. Estimation of parameters  $\beta$  is described in section 3.2. Column (2) add a set of lagged controls to baseline specification: polynomial in size (assets) and the ratios of value added, liquidity, cash flow and bank debt to assets. It excludes observation with missing or zero values for any control variable. Column (3) excludes any firm that, at any point in time, received more than 1% of the credit by any financial intermediary. Column (4) use pooled estimator (rather than “within”) by dropping firm FEs. Column (5) includes firm FEs, province FEs, year FEs and industry FEs, but do not include province×year×industry FEs. Column (6) uses an alternative measure of credit supply shocks which control for match-specific covariates, see section 3.1. Column (7) uses, as an instrument, an alternative credit supply shocks estimated with a “split sample” procedure, in order to control for finite sample biases. Column (8) uses a 4-digits (rather than 2) industry definition both for the estimation of productivity parameters and for the FEs structure. It contains less observations because of the singleton dropping. Column (9) estimate productivity allowing for endogenous firm exit, as in [Olley & Pakes \(1996\)](#). Standard errors (in parentheses) are clustered at firm level. \*\*\* p<0.01, \*\* p<0.05, \* p<0.1

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