(Working Papers)

The micro dynamics of exporting: evidence from French firms

by Ines Buono and Harald Fadinger

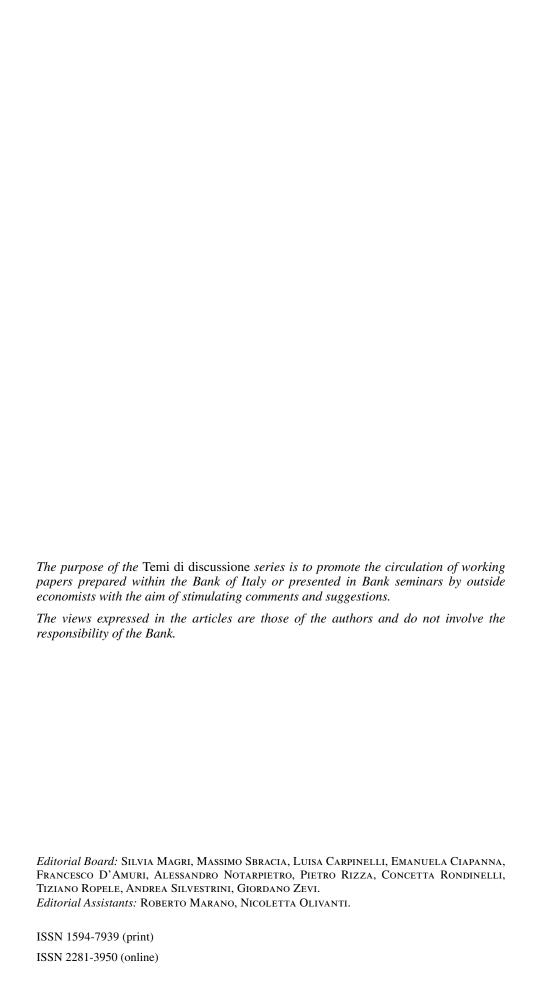


Temi di discussione

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Abstract

The paper investigates the dynamics of export relationships – defined as shipments by a given firm to a given destination in a given year – in a panel of almost 25,000 French exporters over the five-year period 1995-1999. It describes how the relationships evolve over time, presenting a number of stylized facts that are related to various theories of export dynamics, such as a dynamic sunk-cost model and the recent literature on exporting and learning. We find that export relationships are highly dynamic: a large fraction are created or concluded every year, and the values within relationships fluctuate substantially. Most of these dynamics are explained by relationship-specific shocks, not supply or demand shocks. Also, export values are small at first but gradually increase as relationships mature. Finally, while many export relationships are volatile, others are persistent. Previous exports to a destination substantially increase a firm's probability of making current exports there. Taken together, these facts correspond better to the learning than to the sunk-cost hypothesis.

JEL Classification: F10, F12, F14.

Keywords: trade dynamics, extensive and intensive margin, firm heterogeneity.

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1 Introduction¹

This paper investigates the dynamics of export relationships – defined as shipments by a given firm to a given destination in a given year – using a panel of almost 25,000 French exporters over the five-year period 1995-1999. We describe how these export relationships evolve over time and present a number of stylized facts, which we relate to different theories of export dynamics, such as a dynamic Melitz (2003) model and the recent literature on exporting and learning.

Our results show that export relationships are very volatile. In a typical year of our sample, around 25% of all relationships are newly created and around 21% are destroyed (leaving a net creation of around 4%). In addition, export values associated with specific export relationships fluctuate a lot. Individual changes in export values add up to approximately 10% of the total value of French exports. Around 90% of the changes in export values occur within existing trade relationships (intensive margin), while newly created or destroyed relationships (extensive margin) contribute only around 10% to the changes in aggregate export values. Thus, while many relationships are created or destroyed every year, these involve small values.

We show that most of the creation and destruction of export relationships and most of the changes in export values are neither driven by firm-specific (productivity) shocks nor by destination-specific (demand) shocks but by shocks that hit individual export relationships. This casts doubt on the relevance of the canonical sunk-cost models of export dynamics (e.g. Baldwin and Krugman (1989), Dixit (1989), Roberts and Tybout (1997)), which emphasize variation in firm productivity and real exchange rates as the main drivers of fluctuations in firms' export decisions. Instead, it lends support to learning models, in which firms face initial uncertainty either about the demand for their product in a given market (Arkolakis and Papageorgiou (2009), Eaton, Eslava, Krizan, Kugler and Tybout (2008)) or about the reliability of their local partner firms (Araujo and Ornelas (2007), Aeberhardt, Buono and Fadinger (2011), Araujo, Mion and Ornelas (2012)). In such an environment, firms face relationship-specific uncertainty and often make mistakes, which may lead to a break-up of export relationships.

Next, we turn to a description of export values at the beginning of an export relationship and show that these involve small values. If an export relationship survives the initial phase, the value of exports grows fast. Again, this is more in line with a learning model, in which firms are initially

¹The views expressed are those of the authors and do not necessarily reflect those of the Bank of Italy or INSEE. We thank Stefan Berger, who started to work on this project with us. We are especially grateful to Romain Aeberhardt for help with the data and to Antonio Accetturo, Paula Bustos, Gino Gancia, Miklos Koren, Francis Kramarz, Gianmarco Ottaviano, Diego Puga, Massimo Sbracia and Jaume Ventura for comments on an earlier version of the paper. We would also like to thank INSEE for its hospitality.

reluctant to put too much at stake because uncertainty is large, than with a sunk cost model, in which firms enter with large quantities in order to overcome the sunk cost hurdle.

Finally, we show that while many export relationships are volatile, at the same time others are quite persistent. Around 46% of relationships are created or destroyed every year but these relationships are not randomly chosen. In particular, having exported to a specific destination in the previous year increases the probability of exporting to the same destination in the current period by almost 70 percentage points, even controlling for productivity and demand shocks. While state dependence of export decisions is a typical outcome of the sunk-cost model (see Roberts and Tybout (1997)), it is also consistent with learning models, in which firms have to export to a destination in order to learn about local demand or their local partners.

We now turn to a discussion of the related literature. Two closely related papers on firm-destination export dynamics are the descriptive studies of Eaton, Eslava, Kugler and Tybout (2007), who investigate the dynamics of Colombian exporters across destinations, and Lawless (2009), who studies the export patterns of a 5-years sample of Irish exporting firms across destinations. While the findings of these authors are broadly consistent with ours, the focus of those contributions is somewhat different. Eaton et al. (2007)'s descriptive analysis is centred on the observation that most new entrants in a given destination export very small values and only few survive in the long run. Those who do survive, however, grow very fast and contribute a fair amount to aggregate Columbian export growth in the longer run. Lawless (2009), on the other hand, is interested in exporters' simultaneous entry into and exit from a given destination, the gradual fashion in which exporters expand the number of destinations to which they export and the small contribution of new relationships to aggregate export growth. None of these papers use formal econometric techniques to support their findings and they also do not relate their results to the theories of export dynamics, as we do in the present paper.²

This paper is also related to the literature on export dynamics and sunk fixed costs (Baldwin and Krugman (1989), Dixit (1989)). Starting with the contributions of Roberts and Tybout (1997) and Bernard and Jensen (2004), a line of empirical work has investigated the dynamics of firms' export status. These papers use firm-level data sets that provide information on firms' aggregate export

²Many other papers, such as Bernard, Eaton, Jensen and Kortum (2003), Bernard, Jensen, Redding and Schott (2007) and Bernard, Redding and Schott (2010) among the others, reveal stylized facts on firms that export, even if they do not focus on export dynamics. Moreover, other works use the same data of this paper to explore other characteristics of firms that export. Eaton, Kortum and Kramarz (2004) are the first to analyze the destination component of firms' exports in a single year, dissecting the behavior of the intensive and the extensive margins of trade. More recently, Berman, Berthou and Héricourt (2011) explore how firms' sales interact across markets, Buono and Lalanne (2011), Berthou and Fontagné (2012) and Buono and Lalanne (2012) analyze how extensive and intensive margins react to change in trade costs, while Bricongne, Fontagn, Gaulier, Taglioni and Vicard (2010) study the effect of the crisis in 2008-2009 on trade margins.

values but do not include data on export values or status by destination. Thus, they do not allow a study of the dynamics of individual export relationships. The main conclusion of those studies is that firms' export status is very persistent and that past export status is an important predictor of current export status, a finding that is interpreted as a piece of evidence in favour of the sunk-cost model. Das, Roberts and Tybout (2007) structurally estimate a model with heterogeneous firms and sunk costs to export using a panel of Columbian exporters. They find the sunk fixed costs to export to be as high as 400,000 US dollars for these firms. Critically evaluating these results, Ruhl and Willis (2008) have shown that the standard model of firm heterogeneity with sunk costs predicts export values which are too large upon entry and hazard rates that increase over time, which is at odds with the empirical evidence.

A more recent line of research is motivated by the empirical observations of Eaton et al. (2007) that entry into export markets usually occurs with small values and that hazards decline with the age of the export relationship. To explain these facts, Eaton et al. (2008) and Arkolakis and Papageorgiou (2009) develop models of Bayesian learning. In this setting, firms are initially uncertain about their local demand in the export market and therefore start small. If they discover that demand is large, however, export values grow fast.³

Finally, Araujo and Ornelas (2007), Aeberhardt et al. (2011) and Araujo et al. (2012)) build models where exporters have to match with a local distributor in each market. Initially, the importer's type is unknown and has to be learned through experience. Some distributors run away with exported goods if they can. As a consequence, export values are initially small and increase as exporters become more confident about the reliability of their partners.

The paper is organized as follows: in Section two we sketch several alternative theories of export dynamics and discuss their implications. Section three describes the data set, Section four is dedicated to the empirics of export dynamics. Finally, Section five concludes.

2 A Shock-Augmented Melitz (2003) Model

In this section we sketch a simple multi-destination extension of Melitz (2003), augmented for three types of shocks: destination-specific demand shocks, firm-specific productivity shocks and relationship (firm-destination)-specific shocks. We derive the model's implications for the levels and growth rates of export values and for export probabilities in order to test to what extent such a model is consistent

³Other papers that emphasize learning about local demand are Segura-Cayuela and Vilarrubia (2008) and Albornoz, Pardo, Corcos and Ornelas (2012), who focus on learning from other exporters (export destinations).

with the patterns observed in the French firm-level trade data.

2.1The Baseline Model

Let firms be indexed by i = 1, ..., I, destinations by c = 1, ..., C and time by t = 1, ..., T. Let A_{ct} be total expenditure of destination c consumers at time t, let τ_{ct} be variable trade costs with destination c at time t (reflecting tariffs, transport costs and real exchange rates with France) and let λ_{ict} be the relationship-specific part of demand, reflecting demand for a particular variety in a specific destination at time t.

Consumers in every destination c have Dixit-Stiglitz preferences⁴ over differentiated varieties, which give rise to the following demand function for individual varieties:

$$x_{ict} = \frac{p_{ict}^{-\varepsilon} \lambda_{ict} A_{ct}}{P_{ct}^{1-\varepsilon}},\tag{1}$$

where p_{ict} is the price of firm i's product in destination c at time $t, \varepsilon > 1$ is the elasticity of demand and $P_{ct} \equiv (\sum_{ict \in I_{ct}} p_{ict}^{1-\varepsilon})^{\frac{1}{1-\varepsilon}}$ is the price index in destination c at time t. Firms are monopolistically competitive and are heterogeneous in productivity ϕ_{it} , which is drawn each period from a distribution $G(\phi)$ with support on $(0,\infty)$. Their costs to produce x_{ict} units for destination c at time t is described by the following cost function:

$$TC(\phi_{it}, x_{ict}) = (\tau_{ct}/\phi_{it})x_{ict} + f_{ct}, \tag{2}$$

where $\tau_{ct} \geq 1$ is an iceberg variable trade cost and f_{ct} is the per-period fixed cost of exporting to destination c at time t.⁵ For each destination, firms maximize per-period profits from exporting subject to demand (1) and their cost function (2). The solution to the profit maximization problem implies that optimal prices are a fixed mark-up over marginal costs:

$$p_{ict} = \frac{\varepsilon}{\varepsilon - 1} \frac{\tau_{ct}}{\phi_{it}} \tag{3}$$

Thus, export values of firm i to destination c at time t are given by:

$$p_{ict}x_{ict} = \left(\frac{\varepsilon - 1}{\varepsilon}\right)^{\varepsilon - 1} \phi_{it}^{\varepsilon - 1} \left(\frac{P_{ct}}{\tau_{ct}}\right)^{\varepsilon - 1} A_{ct}\lambda_{ict} \tag{4}$$

 $^{{}^4}U_c = \sum_{ict \in I_{ct}} (\lambda_{ict}^{-\varepsilon} \chi_{ict}^{\frac{\varepsilon-1}{\varepsilon}})^{\frac{\varepsilon}{\varepsilon-1}}.$ ⁵It is standard to introduce variable costs as iceberg costs: for each unit of good to arrive in destination c, the firm has ship $\tau_{ct} \geq 1$ units of the good.

Hence, up to a constant, we can write log export values as:⁶ $\log(p_{ict}x_{ict}) \approx d_{it} + d_{ct} + u_{ict}$, where d_{it} represents firm-time-specific factors, d_{ct} represents destination-time-specific factors and u_{ict} stands for relationship-specific factors. Denoting the difference operator by Δ , export growth rates can be written as

$$g_{ict} = \Delta \log(p_{ict}x_{ict}) = \Delta d_{it} + \Delta d_{ct} + \Delta u_{ict}. \tag{5}$$

Thus, export growth rates are explained by three components: firm-specific productivity shocks (Δd_{it}), destination-specific demand shocks (Δd_{ct}) and relationship-specific shocks (Δu_{ict}).

We now derive the conditions for firms' export decisions to a given destination. Profits from exporting to destination c at time t are given by:

$$\Pi_{ict} = \frac{p_{ict}x_{ict}}{\varepsilon} - f_{ct} = 1/\varepsilon \left(\frac{\varepsilon - 1}{\varepsilon}\right)^{\varepsilon - 1} \phi_{it}^{\varepsilon - 1} \left(\frac{P_{ct}}{\tau_{ct}}\right)^{\varepsilon - 1} A_{ct}\lambda_{ict} - f_{ct}.$$
 (6)

Therefore, firm i exports to destination c at time t if and only if $\frac{\Pi_{ict}}{f_{ct}} > 1$. Defining the latent variable $Z = \log\left(\frac{\Pi_{ict}}{f_{ct}}\right)$, a firm exports to a given destination if and only if Z > 0 and zero otherwise, where up to a constant Z can be written as the sum of firm-time, destination-time and relationship-specific factors: $Z \approx d_{it} + d_{ct} + u_{ict}$. Thus, the probability that firm i exports to destination c at time t is given by $Prob(Z > 0) = Prob(Y_{ict} = 1)$, where $Y_{ict} \in \{0, 1\}$ is an indicator variable for firm i's export status to destination c at time t. Hence:

$$Y_{ict} = 1 \iff d_{it} + d_{ct} + u_{ict} > 0. \tag{7}$$

Summing up, we have the following predictions on export values/growth rates and export status:

- Changes in export values/ growth rates should be driven primarily by destination-specific demand shocks and by firm-specific productivity shocks because relationship-specific shocks are a residual that is unexplained in the model.
- 2. There should be no relation between export values and the age of the export relationship conditional on productivity and destination-specific demand. Firms enter a destination whenever they can overcome the fixed cost hurdle (when productivity and/or local demand is high) and choose the optimal quantity in each period.
- 3. The current export status should depend primarily on contemporaneous productivity and destinationspecific demand, since changes in relationship-specific demand, which may cause entry or exit,

⁶Here,
$$d_{it} \equiv (\varepsilon - 1) \log(\phi_{it}), d_{ct} \equiv \log \left(A_{ct} \left(\frac{P_{ct}}{\tau_{ct}} \right)^{\varepsilon - 1} \right)$$
 and $u_{ict} \equiv \log(\lambda_{ict})$.

remain unexplained in the model.

4. The model implies that there is no state dependence of export status, in the sense that having exported to a destination in the previous period does not affect the probability of exporting to the same destination in the current period conditional on firm-time and destination-time specific factors.

2.2 Introducing Sunk Costs

We now add sunk fixed costs to export to the model. Let $\Omega_{ict} = (\phi_{it}, A_{ct}, P_{ct}, \tau_{ct}, \lambda_{ict})$ be the information set of firm i regarding export destination c at time t and let $\Pi_{ict}(\Omega_{ict})$ be per-period profits as in equation (6). Assume now that in addition to the per-period fixed exporting cost there is also a sunk fixed cost of exporting, \tilde{f} , which has to be paid upon entry to any given destination c. Moreover, assume that \tilde{f} is identical for all countries and firms. Finally, destination-specific demand and firm-specific productivity both follow AR(1) processes: $\log(A_{ct}) = \log(\bar{A}_c) + \rho_A \log(A_{ct-1}) + \nu_{ct}$ and $\log(\phi_{it}) = \log(\bar{\phi}_i) + \rho_{\phi} \log(\phi_{it-1}) + \epsilon_{it}$, where ν_{ct} and ϵ_{it} are normally distributed i.i.d shocks. In such a model, the price and quantity decisions conditional on exporting remain static as before. However, the presence of the sunk cost together with the Markov-processes for productivity and demand convert the firm's entry decision into a dynamic optimization problem, since the firm has to forecast future values of productivity and destination-specific demand when deciding about its export status in any given period. Thus, firms have to choose an infinite sequence of entry decisions $\{Y_{ict}, Y_{ict+1}...\}$. The Bellman equation for firms' entry problem is given by:

$$V_{ict}(\Omega_{ict}) = \max_{Y_{ict}} \{ Y_{ict}[\Pi_{ict}(\Omega_{ict}) - \tilde{f}(1 - Y_{ict-1})] + \delta E_t[V_{ict+1}(\Omega_{ict+1})|Y_{ict}] \}, \tag{8}$$

where δ is the discount factor and E_t denotes expectations conditioned on information set Ω_{ict} . It is optimal for firm i to enter destination c in period t, i.e. to choose $Y_{ict} = 1$ whenever

$$\Pi_{ict}(\Omega_{ict}) + \delta E_t[V_{ict+1}(\Omega_{ict+1})|Y_{ict} = 1] - \delta E_t[V_{ict+1}(\Omega_{ict+1})|Y_{ict} = 0] - \tilde{f}(1 - Y_{ict-1}) \ge 0.$$
 (9)

Note that \tilde{f} enters the expression both through past export status and the expression for the expectations as can be seen from equation (8). Similarly, Ω_{ict} enters the equation through $\Pi_{ict}(\Omega_{ict})$ and the expression for expected future profits because, given the Markov property of the information set, the current state set helps to forecast future values of Ω_{ict} . This implies, of course, that more persistent shocks have larger effects on Y_{ict} , while small shocks may not move export status at all.

We can write condition (9) in reduced form by proxying for the term $\Pi_{ict}(\Omega_{ict}) + \delta E_t[V_{ict+1}(\Omega_{ict+1})|Y_{ict} = 1] - \delta E_t[V_{ict+1}(\Omega_{ict+1})|Y_{ict} = 0]$ with a combination of firm-time (δ_{it}) , destination-time (δ_{ct}) and relationship-specific (u_{ict}) factors. In this way, we obtain a reduced-form equation for the probability of exporting:

$$Y_{ict} = 1 \iff \beta_1 Y_{ict-1} + \delta_{it} + \delta_{ct} + u_{ict} > 0 \tag{10}$$

Thus, in the presence of sunk costs, export decisions are state-dependent, in the sense that past export status matters for the current probability of exporting. Firms are more likely to export to a given destination once they have paid the sunk cost than when they have not, even when conditioning on productivity and destination-specific demand. This is because there are values of productivity and demand such that the net present value of exporting is positive if and only if the sunk cost has already been paid (see equation (9), where the term $\tilde{f}(1-Y_{ict-1})$ becomes zero whenever $Y_{ict-1}=1$). Note also that given the linearity of the reduced-form model the coefficient of past export status measures the size of the sunk cost, since $\tilde{f}Y_{ict-1}=\beta_1Y_{ict-1}$. Moreover, \tilde{f} also enters in the (non-linear) expressions for the expectations: this is captured as part of the fixed effects and the error term, μ_{ict} .

Since the expected value from exporting is positively related to the expected net present value of per-period export profits – which are positive when per-period export revenues are larger than the per-period fixed cost f_{ct} – a given proportional shock to destination-specific demand (A_{ct}) is less likely to induce firms to exit from a larger market (with higher \bar{A}_c). The reason is that in such a market, even when demand is low per-period profits are still positive. This implies that the previous export status should have a larger effect on the probability of exporting in larger markets. Similarly, a given proportional shock to idiosyncratic productivity should make exit less likely for those firms that have on average higher productivity (higher $\bar{\phi}_i$). Thus, in a regression of current export status on its past value and interactions of past export status with market size and productivity we expect the coefficients on the interaction terms to be positive:⁸

$$Y_{ict} = 1 \iff \beta_1 Y_{ict-1} + \beta_2 Y_{ict-1} A_{ct} + \beta_3 Y_{ict-1} \phi_{it} + \delta_{it} + \delta_{ct} + u_{ict} > 0$$

$$\tag{11}$$

Summarizing, the sunk-cost model has the following testable implications:

⁷Thus, correlation between observables and the error term seems likely. This may bias the estimate of the effect of past export status. Other studies, such as Roberts and Tybout (1997) or Bernard and Jensen (2004), face a similar problem since they also proxy for the non-linear expectation term with a linear expression of observables. They just assume that the error is uncorrelated with observables.

⁸Note that this model – at least in reduced form – is indistinguishable from a model where sunk fixed costs are destination- and/or firm-specific, being larger for larger destinations or more productive firms, i.e. $\tilde{f} = \tilde{f}_{ic} = \tilde{f}_c + \tilde{f}_i$. Such a model would also imply positive coefficients on the interaction terms.

- 1. Firms should enter with large export values in order to recoup the sunk fixed cost.
- Once a firm has entered an export destination, export values/growth rates should be explained mainly by current productivity and demand shocks and should be independent of relationship age.
- 3. Export decisions should be state-dependent.
- 4. State dependence of export decisions should be larger in bigger markets and for more productive firms.

2.3 Introducing Learning

The sunk-cost model provides an explanation for state dependence of exporting decisions. However, it does not provide any micro-foundation for relationship-specific shocks. We now sketch two models that micro-found such shocks: a first one, where there is uncertainty about product-specific demand in a given destination that has to be learned through experience and a second one, in which exporters are uncertain about the reliability of their local partners.

2.3.1 Learning about Local Demand

Here, we briefly sketch Arkolakis and Papageorgiou (2009)'s model of Bayesian learning about local demand. Initially, firms are uncertain about the demand for their product in a particular destination and they have to learn it through experience. The set-up is as follows. First, firms draw productivity according to a Markov process. Next, they choose the export quantity and pay a per period exporting fixed cost. Subsequently, firms receive a noisy signal about local demand. Given the signal, firms update their beliefs and decide whether to stay or exit.

The signals about idiosyncratic demand have the following form: Observed demand is $\lambda_{ict} = \exp(\alpha_{ic} + \epsilon_{ict})$, where α_{ic} is the true demand parameter that is drawn upon entry from a Normal distribution and $\epsilon_{ict} \sim N(0, \sigma^2)$ is i.i.d. noise.

In each period, firms optimally choose quantities in a static way given current expected demand, before the signal is observed. They maximize expected per-period profits

$$E_t[\Pi_{ict}(\Omega_{ict})] = E_t(p_{ict})x_{ict} - \frac{\tau_{ct}}{\phi_{it}}x_{ict} - f_{ct}, \tag{12}$$

subject to the expected inverse demand function $E_t(p_{ict}) = x_{ict}^{\frac{-1}{\varepsilon}} P_{ct}^{\frac{\varepsilon-1}{\varepsilon}} E_t(\lambda_{ict}^{\frac{1}{\varepsilon}}) A_{ct}^{\frac{1}{\varepsilon}}$. Here the expectations are over the distribution of λ_{ict} conditional on having received n signals with mean $\alpha_{ic}(n)$ and $\Omega_{ict} = 0$

 $(\phi_{it}, A_{ct}, \tau_{ct}, P_{ct}, n, \alpha_{ic}(n))$. Then export values conditional on having received n signals are given by

$$p_{ict}x_{ict}(n,\alpha_{ic}(n)) = \left(\frac{\varepsilon - 1}{\varepsilon}\right)^{\varepsilon - 1} \phi_{it}^{\varepsilon - 1} A_{ct} \left(\frac{P_{ct}}{\tau_{ct}}\right)^{\varepsilon - 1} \lambda_{ict}^{\frac{1}{\varepsilon}} [E_t(\lambda_{ict}^{\frac{1}{\varepsilon}})]^{\varepsilon - 1}.$$
 (13)

Due to Bayesian learning $E_t(\lambda_{ict}^{\frac{1}{\varepsilon}})$ follows a Markov process.⁹ This implies that learning provides an explanation for relationship-specific shocks. If $E_0(\lambda_{ic0}^{\frac{1}{\varepsilon}})$ is small, initially, export values are relatively small due to uncertainty about local demand. Once firms have received several positive signals, they become confident that demand for their product is high and export values increase over time. Differently, negative signals cause a decrease in export values (and eventually exit). Since the learning gain is decreasing over time, export growth rates are initially large and decreasing over time.

Learning also makes the entry and exit decision dynamic: firms have to take into account that they only learn about local demand as long as they keep exporting. Thus, the export decision is the solution to the following Bellman equation:

$$V_{ict}(\Omega_{ict}) = \max_{Y_{ict}} \{ E_t[\Pi_{ict}(\Omega_{ict})] Y_{ict} + \delta E_t[V_{ict+1}(\Omega_{ict+1})|Y_{ict}] \}$$
(14)

As a result of learning about local demand, firms are willing to make initial losses in order to learn their idiosyncratic demand. Thus, hazard rates are initially high, but after receiving a number of positive signals, they decline fast. Moreover, export decisions are state-dependent: since the idiosyncratic part of demand follows a Markov process, having exported in the previous period increases the probability of exporting in the current period, even when controlling for productivity and destination-specific demand, because it correlates with high values of idiosyncratic demand. Finally, in larger destinations, where average demand is higher, or for more productive firms the probability of a given export relationship surviving from one period to the next is greater even when conditioning on market size and productivity. This is because in these cases a smaller value of idiosyncratic demand is sufficient to make exporting worthwhile.

Thus, the model has the following predictions:

- 1. Changes in export values/growth rates should be driven by relationship-specific shocks in addition to productivity and demand shocks.
- 2. Export values should be initially small and increasing with age conditional on survival.

⁹In particular, one can show that the optimal forecast is given by $E_t[\log(\lambda_{ict})] = (1 - k_t)E_{t-1}[\log(\lambda_{ict-1})] + k_t\epsilon_{ict}$ where $k_t = \frac{\hat{\sigma}_t^2}{\hat{\sigma}_t^2 + \sigma^2}$ is the Kalman gain and $\hat{\sigma}_t^2$ is the posterior error covariance in period t.

- 3. Growth rates of export values should be decreasing with age conditional on survival.
- 4. Changes in export decisions should be driven by relationship-specific shocks in addition to productivity and demand shocks.
- 5. Export decisions should be state-dependent due to learning.
- 6. State dependence of export decisions should be larger in bigger markets and for more productive firms.

2.3.2 Learning about Local Partners

Finally, we sketch a model of learning about local partners. Araujo and Ornelas (2007), Aeberhardt et al. (2011) and Araujo et al. (2012) develop a model where exporters have to match with local partners (importers) in each destination in order to export. There is incomplete information about the reliability of importers – some importers may violate contracts if local legal institutions are bad and hold up the exporters – and exporters have to learn the type of their partners through experience. Thus, initially firms export small quantities in order not to expose themselves to large losses if the importer does not respect the contract. While a contract violation leads to a termination of the export relationship, each time the contract is respected, the exporter becomes more confident that their partner is trustworthy and exports increase. This leads to export values that increase with the duration of the relationship and hazard rates that decline over time, as relationships involving unreliable partners are weeded out. Finally, export status is state dependent: a given firm is more likely to export to a given destination if she has exported there in the previous period because partners can be found only with a given probability and exporters are reluctant to give up a partner as long as they do not observe a contract violation. Moreover, state dependence depends on exporter and destination characteristics. In particular, state dependence is larger for more productive exporters and in larger export destinations because those relationships are more valuable for importers and so it is easier to sustain cooperation. Finally – and this prediction is specific to this type of model –, state dependence is larger in destinations with better legal institutions because unreliable importers who try to violate a contract are prevented from doing so by the legal system.

As in the baseline model sketched above, exporters are monopolists for their particular variety and consumers have Dixit-Stiglitz preferences. Thus, price and quantity decisions are static and export values are given by:

$$p_{ict}x_{ict} = \left(\frac{\varepsilon - 1}{\varepsilon}\right)^{\varepsilon - 1} \phi_{it}^{\varepsilon - 1} \left(\frac{P_{ct}}{\tau_{ct}}\right)^{\varepsilon - 1} A_{ct}\lambda_{ict},\tag{15}$$

where $\lambda_{ict} = \alpha[1 - \theta_{ict}(1 - \gamma)]^{\varepsilon-1}$. Here, α is the fraction of profits that the contract assigns to the exporter, θ_{ict} is the subjective probability of the exporter that her partner is unreliable and γ is a measure of the effectiveness of the local legal system. Thus, export values are decreasing in θ_{ict} . Again, due to Bayesian learning about the type of the partner, λ_{ict} and therefore export values follow a Markov process.¹⁰

Summarizing, this model delivers very similar predictions to the model of learning about local demand. In particular, the list of predictions from Section (2.3.1) is also valid for the model of learning about local partners.¹¹

2.4 Implications

Here, we briefly sum up the testable implications of the different models sketched above. First, all models predict the presence of volatility in the extensive and the intensive margin of trade at the firm-country level. However, while according to the shock-augmented Melitz model both changes in export values (or export growth rates) and changes in export decisions are mainly the result of productivity and demand shocks, according to learning models relationship-specific shocks, for which different micro-foundations are provided, should be an important driver of those changes. Both models predict the existence of state dependence of export decisions. In the first case there is an option value of continuing a relationship because of the sunk entry cost and in the second case the option value of continuing a relationship is due to firms trying not to lose important information collected earlier. Moreover, in both models state dependence should be market- and firm-specific, i.e. it should be higher for bigger markets and more productive firms. However, while in the learning models firms start exporting by shipping small quantities, which eventually increase as relations get older and more reliable, the shock-augmented Melitz model predicts high export values upon entry that are independent of the age of the relationship. In next section we take these predictions to data to empirically explore firm export behaviour and to differentiate between the different models of export dynamics.

¹⁰Specifically, the subjective probability of the exporter that her partner is impatient conditional on the contract being respected is $\theta_{ict}(r) = \frac{\gamma \theta_{ict-1}}{\gamma \theta_{ict-1} + 1 - \theta_{ict-1}} < \theta_{ict-1}$, and the subjective probability conditional on a contract violation is $\theta_{ict}(v) = 1$.

¹¹An exception is the prediction on the growth rate of export values, which should be increasing over time in this model.

3 Data

The main data source for our analysis is the Douanes data base, available at the French Statistical Agency (INSEE), which contains all French customs data. For each firm, it allows us to precisely observe its exports to any destination in a given year. Each firm is assigned to a sector using the 2-digit NES classification system. Thus, excluding agriculture and services, we have firms in 15 manufacturing industries.¹²

Douanes data report 97% of the value of national trade. According to the requirements of Eurostat, Douanes data should contain all flows which are above 1,000 euros for extra-EU trade. The reporting threshold for intra-EU trade, instead, changed several times in the sample period. It went from 250,000 FF to 650,500 FF in 2001 and then was changed to 100,000 euros in 2002. Export values below this threshold are usually reported but reporting is incomplete because it is not compulsory under French law. To the extent that reporting below the threshold is a random draw from the population, this should not affect our results. In unreported robustness checks, which are available upon request, we have excluded EU destinations from our sample. All our results remained unaffected.

We combine the Douanes data with the Bénéfices Réels Normaux (BRN) data base, also available at INSEE, which provides detailed balance-sheet information. This data base allows us to construct labor productivity by firm measured as value added per worker. We take all and only those firms which exist continuously and export in at least one year in the time-span we are analyzing. We abstract from non-exporters in our time-span, as they do not add any information for our analysis.¹³ Thus, our sample is representative for the set of all French potential exporters during the sample period.¹⁴

To capture all the volatility of export relationships we decide to use all possible destinations in the dataset for which we also have data on covariates, which we use in the last part of the analysis, for a total of 146 destinations.¹⁵ This choice forced us to limit the numbers of years of our analysis to 5 (we choose 1995-1999). In fact, considering the number of firms that export at least once to at least one destination in that period, the total amount of relationships may reach 18 millions (=5*146*24,536), the maximum number of observations we could handle given the computational constraints we were

¹²A finer disaggregation is possible but not pursued in this paper as we are more interested in aggregate patterns than in sector-specific differences.

¹³In fact their behavior is perfectly explained by firm-time fixed effects.

¹⁴Note that not including non-exporters has no impact on our results on export dynamics as long as (sunk) fixed costs to export are destination-specific, as our theoretical model suggests. We do not analyze sunk export costs that are independent of the export destination because this would require to use data on non-exporters.

¹⁵Country names and codes are reported in the appendix.

subject to.^{16,17} A different approach is to restrict the number of destinations and instead to increase the number of years. To analyze the drivers of creation and destruction of export relationships as well as variation at the intensive margin at the destination level, the five year sample period is long enough.¹⁸ However, for a survival analysis we would need a longer dataset, since the maximum duration in our sample is five years.

While we leave for future research robustness checks with different samples (longer time-span, sample which includes also firms that entry/exit the domestic market, and so on), our results are robust to the exclusions of some countries (those in the European union) and to the exclusions of very small and very big firms (we replicate results excluding, alternatively the 1st and the 99th and the 10th and the 90th percentile of export values.).¹⁹

Finally, there are several other studies, which confirm some of our findings. The first relies on the results reported in Aeberhardt et al. (2011) who use the Douanes database with 75 destinations, 13 years and around 7,000 firms, where the lower number of firms is partially due to attrition. Starting from the empirical evidence on export dynamics, among which the relevance of relationship-specific shocks highlighted in this paper, the authors test the implications of a learning model of exporting. Using a long data-set allows them to pursue a survival analysis which sheds light on the determinants of the length of the active relationships. Further external validation is provided by studies on other countries, like Lawless (2009) on a sample of Irish firms in the five-year time-span from 2001-2004, and Eaton et al. (2007) with a sample of Colombian firms, between 1996 and 2005.

We also use several control variables that come from other sources. Data on average real GDP and real GDP per worker for the sample period are from the Penn World Tables (Mark 6.2) and data on distance from Paris are taken from Rose (2004).

Table 1 presents descriptive statistics for firm and country variables. Export values, productivity and the number of export destinations by firm follow left-skewed distributions with a long right tail. The median export value is around 29,732 euros (exp (10.3)). The median number of destinations a firm exports to is 5, with a mean of 12 and a maximum of 143.

¹⁶Note that we require any given firm to be an exporter both in Douanes and in BRN and we restrict the sample to firms that survived for the entire time-span. While the first requirement was necessary to have a cleaner database, the second was a simplifying choice. Note that this choice does not bias our results since attrition occurs mainly over a longer time-span and for small domestic firms rather than for exporters and here we are considering only exporters.

¹⁷Our data source is the same as the one used by Eaton et al. (2004) and Eaton, Kortum and Kramarz (2011). They report 34,035 exporters for the year 1986 that sell to 113 destinations outside France. We have fewer exporters in our data set for several reasons. First, we require exporters to exist continuously during the sample period. Second, we require firms to be both in the Douanes and in the BRN database and to have information on value added and employment. Finally, we focus on manufacturing sectors.

¹⁸Also Lawless (2009) uses a 5 years sample to provide evidence on entry/exits of firms into different destinations.

¹⁹Results are available upon request.

4 Dynamics of Export Relationships and Export Values

In this section we describe the export dynamics of French firms. We define an export relationship as observing a positive export value by a given firm to a given destination in a given year. 20 In this case, the indicator Y_{ict} is equal to 1. When a relationship is created or destroyed, the value of exports changes through the extensive margin. Conversely, when the value of the exports changes within an existing relationship then trade is adjusting through its intensive margin. We first show how volatile export relationships and export values are. We then compute the contribution of productivity-, demand-, and relationship-specific shocks to changes in the status of export relationships and to changes in export values.

4.1 Volatility of Export Status, Export Relationships and Export Values

We start by describing fluctuations in export status, i.e. participation in export activity, which is the margin of adjustment analysed by Bernard and Jensen (2004).²¹ In Table 2, we report for each year the number of exporters in the sample, the number of firms which cease to export and those which begin to do so. From one year to the next, almost 9% of exporters cease to export; conversely, a slightly higher 12% are new exporters. In a typical year of our sample, there is a net increase in the number of exporters, which – aggregating entries and exits into export activity – turns out to be relatively small (3%).

More dynamics can be uncovered when we dig deeper and investigate the volatility of export relationships.²² Entry into and exit from specific export destinations are very frequent phenomena. In column (2) of Table 3 we report for each year the number of active relationships in the sample.

Columns (3) and (5) report the number of destroyed and created relationships year by year. We find that each year around 25% of all firm-destination relationships are newly created, while around 21% of relationships are destroyed, with the difference being positive net creation of export relationships. This suggests that there are a lot of trade micro-dynamics that remain hidden when we aggregate statistics to a firm's overall export status. Finally, it is worth noticing that around 50% of the destroyed

²⁰We do not observe exports of different products to a given destination. To the extent that firms export multiple products to a given destination, we do not observe shocks that affect only the number of products exported to a given destination. As a result, we will tend to underestimate the role of extensive margin adjustment for these exporters, since movements at the extensive margin will occur when a firm stops exporting all its products to a given destination.

²¹Bernard and Jensen (2004) use a data set that only provides information on whether a firm is an exporter or not. In our case we also know to which destination a firm exports, and therefore we can separately analyse the export-status and export-relationships of each firm.

²²Note that export status volatility is a lower bound for relationship volatility. Both are identical if and only if all firms simultaneously enter into or exit from all destinations. Otherwise relationship volatility is larger than export status volatility.

relationships are re-created in at least one subsequent year and around 70% of created relationships are destroyed in at least one subsequent year in the sample. We can conclude that export relationships are very volatile.

We next analyse volatility in terms of export values. Are export values as volatile as export relations? We address this issue by separating the adjustment in export values that occurs within newly created/destroyed relations (extensive margin) from the one that occurs within existing ones (intensive margin).

We denote with Q_t the value of aggregate French exports (given by the sum of export values of all existing relations in a year, q_{ict}), and index firms by i, countries by c and years by t. Thus,

$$Q_t = \sum_{i \in I} \sum_{c \in C} q_{ict}. \tag{16}$$

We consider growth in export values using midpoint growth rates:²³

$$G_t = \sum_{c \in C} \sum_{i \in I} g_{ict} s_{ict}, \tag{17}$$

where s_{ict} is the average export share of firm i in country c in total French exports, $s_{ict} = \frac{q_{ict} + q_{ict-1}}{Q_t + Q_{t-1}}$, and g_{ict} is the midpoint growth rate of export value of firm i in country c, $g_{ict} = \frac{q_{ict} - q_{ict-1}}{1/2(q_{ict} + q_{ict-1})}$.

To see to what extent adjustments in export values are due to the extensive margin and to the intensive margin we classify all export relationships into four subsets: entry – newly formed relationships (those for which $q_{ict-1} = 0$ and $q_{ict} > 0$), exit – destroyed relationships (for which $q_{ict-1} > 0$ and $q_{ict} = 0$), increase – continuing relations for which export values increase ($0 < q_{ict-1} < q_{ict}$), and decrease – continuing relations for which export values decrease ($q_{ict-1} > q_{ict} > 0$). We can thus write:

$$G_t = \sum_{ic \in entry_t} g_{ict} s_{ict} + \sum_{ic \in exit_t} g_{ict} s_{ict} + \sum_{ic \in increase_t} g_{ict} s_{ict} + \sum_{ic \in decrease_t} g_{ict} s_{ict}.$$
(18)

To get a better sense of the magnitudes and the relative contributions of each of the four terms we take absolute values of midpoint growth rates of all firm-destination relationships and aggregate them to obtain the gross export growth rate, \hat{G}_t :

$$\widehat{G}_t = \sum_{ic \in entry_t} |g_{ict}| s_{ict} + \sum_{ic \in exit_t} |g_{ict}| s_{ict} + \sum_{ic \in increase_t} |g_{ict}| s_{ict} + \sum_{ic \in decrease_t} |g_{ict}| s_{ict}$$
(19)

²³This overcomes the problem that we would have with ordinary growth rates, which are not defined for created/destroyed relations. Note that the midpoint growth rate lies in the interval [-2,2] and takes the value -2 in the case of exit and 2 in the case of entry.

Table 4 reports the gross (midpoint) growth rate, the contribution of each of the four components of decomposition (19), as well as the aggregate net growth rate for different years. The net midpoint growth rate of exports is roughly 1%, while the gross midpoint growth rate is almost 10%. This difference indicates that export values are very volatile as well, although less so than export relationships. The contributions of newly-created and destroyed relationships to the gross growth rate are respectively 7.3% and 3.6% in 1996. The intensive margin explains the rest, with increasing values within existing relations explaining 48% and decreasing values within existing relations explaining 41%. This pattern is very similar across different years.²⁴ Moreover, as shown in Figure 1, the contribution of intensive and extensive margins to export dynamics seem to be quite similar across different sectors, suggesting that the volatility of export values is mainly due to changes along the intensive margin in all sectors. We conclude that while creation and destruction of export relationships is very frequent, those relationships involve shipments of small values, thereby contributing relatively little to the aggregate volatility of export values.

4.2 Explaining Volatility of Relationships and Export Values

We now investigate which kinds of shocks are responsible for the choice of French firms to enter and exit from various destinations (creation an destruction of export relationships) and to adjust their export growth rates. Are changes mainly due to firm-specific supply and destination-specific demand shocks or are they due to those shocks that hit a specific export relationship (firm-destination-specific shocks)? Should we find that the last type of shock is important, our analysis would support those theories that provide a micro-foundation for this kind of shock (i.e. learning models).

According to our shock-augmented Melitz model, the export decision of a given firm i to a given destination c at time t can be expressed as

$$Y_{ict} = \delta_{it} + \delta_{ct} + u_{ict}. \tag{20}$$

Using a linear probability model with firm-time dummies, δ_{it} , and destination-time dummies, δ_{ct} , we can decompose the variance of Y_{ict} into the variance of δ_{it} (firm-time-specific component), the variance of δ_{ct} (destination-time-specific component), and the residual variance (firm-destination-time-specific component) using standard analysis of variance (ANOVA) methods. Standard errors

²⁴Following our methodology, (Bricongne et al. (2010)) apply this decomposition using a sample of French data at the product level that covers years from 2002 to 2007. Results are in line with the ones reported here. For instance, they find that the contributions of newly-created and destroyed relationships, which they define at the firm-destination-product level, are respectively 6.5% and 5.9% in their sample period.

are clustered at the firm-destination level. In Table 5 we present the results of this decomposition by 2-digit NES sector.^{25,26} For a typical sector, around 15% of the variance of export decisions is explained by the firm-time-specific component (productivity), and another 15% by the destination-time-specific component (effective market size), while the remaining 70% are residual variance. Thus, the relationship-specific component is very important in explaining variation in export decisions. There is also quite a lot of variation across sectors in terms of the relative importance of the supply and demand component. For instance in the sector "Apparel, Textile and Leather Products" 19% of the variance is explained by the variance of supply against 9% explained by the variance of demand; conversely more than 21% of the export status of firms in the sector "Drugs, Soaps and Cleaners" is explained by the variance of supply, against 14% explained by the variance of demand.²⁷

We next look at changes in the export status as the dependent variable. As suggested by our simple shock-augmented Melitz model (6), changes in export status should be driven by changes in export profits which may be due to firm-specific shocks, destination-specific shocks or relationship-specific shocks (i.e. the export status changes if $\Pi_{ict-1}/f_{ct-1} < 1$ and $\Pi_{ict}/f_{ct} > 1$ or the other way round). We define C_{ict} as an indicator variable, which is equal to 1 if a firm enters or exits from a destination $(\Delta Y_{ict} \in \{-1,1\})$ and 0 if a firm does not change its export status to a given destination $(\Delta Y_{ict} = 0)$. Again, we use a linear probability model to decompose the variance of C_{ict} :

$$C_{ict} = \delta_{it} + \delta_{ct} + u_{ict}. (21)$$

The explanatory variables include firm-time dummies, δ_{it} , and destination-time dummies, δ_{ct} , while u_{ict} is an error term. Again, we perform the analysis by 2-digit NES sector. Results are reported in Table 6. It is evident that the bulk of the variation in entry and exit decisions (around 92,6% on average) remains unexplained by the sum of supply and demand shocks. Hence, most of the creation and destruction of export relations is due to relationship-specific shocks. Note that this pattern is very stable across different sectors. Moreover, the portion of variability explained by demand shocks (around 3 to 6 %) is rather similar to the one related to supply shocks (2 to 4%).²⁸

²⁵Using the full sample for the ANOVA is not possible due to technical limitations. Moreover, analysis by sector has the advantage of making demand shocks sector-specific. Note that the number of observations in the extensive margin analysis is much higher than that in the intensive margin one, since in the former we need to include all possible destinations for each firm and each year, thus obtaining 17,911,280 observations (=24,536*146*5), while the latter only includes observations with positive export values.

²⁶Sample selection is not a problem for the extensive margin analysis, since we use the full set of potential exporters and destinations. For the intensive margin analysis we use firm-time and country-time fixed effects. These should capture the main drivers of selection. If selection is on observables, controlling for them eliminates sample selection problems.

²⁷Note that even when we run regression (20) adding past exporting status as explanatory variable, the residual explains almost 40% of the variance. We report these results in Table 13.

²⁸Note that the firm-time and the destination-time component explain a much smaller fraction of the variance than in

Finally, we analyse the determinants of changes in export values of firms across destinations. Again, we conduct an ANOVA analysis, regressing midpoint export growth rates for each firm in each served destination between any two years, g_{ict} , on a set of firm-time (δ_{it}) and country-time (δ_{ct}) dummies as suggested by equation (5):

$$g_{ict} = \delta_{it} + \delta_{ct} + u_{ict}, \tag{22}$$

where the dummies have the usual interpretation. In Table 7 we report the fraction of the variance of the change in the intensive margin explained by the two sets of dummies as well as the residual variance, which represents the contribution of relationship shocks, for two samples: First, for the sample excluding entry and exit; second, for the sample including these observations. Results are very similar for both samples. Once more, demand and supply shocks have rather small explanatory power to explain intensive margin changes, representing respectively less than 1% and 17% of the total variance in the model. Relationship-specific shocks are instead what really matters for the growth rate of export value, explaining up to 82% of the total variation in growth rates. Although these shocks are important in all sectors, their explanatory power varies quite substantially across sectors from a minimum of 66% in "Printing and Publishing" to a maximum of 85% in the "Mechanic Equipment" industry. Moreover, supply shocks are much more important than demand shocks, in explaining growth rate variability.

Overall, the analysis reveals that relationship-specific shocks are key for explaining the dynamics of export relationships (extensive margin) and export values (intensive margin), while productivity and demand shocks are far less important. This corroborates the importance of learning models, which provide mechanisms for such shocks.

4.3 Small Export Values Upon Entry and the Role of Age

In this subsection we first document that export values are small when export relationships start and we then explore how export values change as relationships mature. The shock-augmented Melitz model predicts that firms should enter with large export values to overcome the fixed (sunk) cost hurdle and that export values are independent of the age of the export relationship. Differently, learning models highlight that initial export values should be small and that export values should grow with age conditional on survival.

the case in which export status is the dependent variable. The reason is that here we are differencing out average firm and destination characteristics.

We now take a closer look at export values in the first and in the last period of an export relationship. In Table 8 we report the average and median export values for all relationships, relationships that were created (terminated) in 1996 and for which this status persisted during the whole observation period and relationships that were created (destroyed) in 1996 but were destroyed (re-created) in some subsequent year. While the median export value for all relationships in 1996 was 27,796 euros, the median export value for relationships that were created in the same year was only 13,266 euros for those relations that survived for the rest of the sample period and 4,193 euros for those relationships that were destroyed in some subsequent year. Thus, initially export values are quite small. Moreover, those relationships that initially involve larger values are more stable. Similarly, those relationships that were destroyed permanently in 1996 involved median exports of 7,670 euros and those that were destroyed in 1996 and recreated in some subsequent year had median export values of 6,131 euros. Thus, also relationships that cease to exist tend to involve relatively small values.

To investigate whether entry values differ substantially across sectors, in Table 9 we report percentiles of export values upon entry by sector. The 10th percentile of export values upon entry varies from 659 euros in the "Printing and Publishing" sector to 1095 euros in the "Drugs" sector, while median export values vary between 3,129 and 7,359 euros (again in the same sectors). We conclude that the phenomenon of observing small entry values does not depend much on the specific sector we look at.

Moreover, even though starting relationships involve small export values, export values tend to increase as relationships mature. This is shown in Figure 2 where we report box plots of export values by age of the relationship.²⁹ Clearly, the median increases over time and the distribution becomes more left skewed, as some relations grow larger.³⁰ A more formal analysis of this phenomenon is reported in Table 10 where we regress export levels (in logs) on the age of the relationship as well as on firm-time and destination-time dummies for the sample excluding entry and exit.³¹ Again, we cluster standard errors at the firm-destination level. Results reveal that, conditional on survival, export values increase strongly with the age of the relationship. Depending on the sector, an increase in age of one year increases export values by 50-70%. Notice also that the fraction of variance in export values explained by the relationship-specific component is around 50 percent.

Finally, when we regress the growth rate of exports, g_{ict} , on age as well as on the usual dummies, we find a negative and significant coefficient ranging from -0.05 to -0.13 for different sectors, as reported

²⁹The box plot reports the median, the 25th and 75th percentiles as well as the minimum and maximum export values.

 $^{^{30}}$ We exclude the 5% largest observations from the plot because including them would make the graph unreadable.

³¹Results for the sample including entry and exit are very similar and are available upon request.

in Table 11.³² This result suggests that, conditional on survival, the growth rate of exporters is larger in the initial years of the export relationship and tends to decline with age. This result is consistent with the first version of learning (Arkolakis and Papageorgiou (2009)).

Summarizing, the fact that firms tend to enter destinations with very small values which increase as the relationship gets older is not consistent with the shock-augmented Melitz model (with or without sunk costs), while it lends support to learning models.

4.4 Persistence of Export Relationships and State Dependence

Both the Melitz model with sunk costs and the learning models predict that (some) export decisions will be persistent. We now document that even though we have previously presented evidence that a large fraction of export relationships are created and destroyed every year, at the same time there is a lot of persistence in export decisions, implying that creation and destruction of export relationships is not random.

First, we use a transition matrix to investigate persistence as well as the patterns of creation and destruction of export relationships in more detail. Each row of Table 12 refers to the firms which export to a given number of destinations, "0", "1", "2" and so on in 1995. Each column refers to the firms which export to a given number of destinations in the subsequent year (1996). Each cell reports the frequency with which firms that exported to a given number of destinations in 1995 transit to any of the column categories in the following year.³³ This means that the rows sum up to 100. The last row reports the frequency of exporters in each category in 1996.³⁴

Almost 60% of all exporters export to four destinations or less and only 6.57% of exporters export to 25 or more destinations. Moreover, the transition matrix is diagonal-dominant. This means that, given any initial number of export destinations, the probability of continuing to export to the same number of destinations is higher than the probability of changing the number of destinations. Non-exporters tend to integrate into the export market gradually, typically by entering one destination only, and firms that exported to only one destination tend to add or drop only one the year after. Indeed, this observation holds for all the categories considered: either firms continue to export to the same number of destinations, or they transit to the nearest category to the left or to the right. This

³²Again, results are reported for the sample excluding entry and exit.

³³The last 3 columns and rows aggregate the number of export destinations in a somewhat arbitrary way. However, results are robust to defining intervals differently.

³⁴Note that, as explained in the description of the data set, here we are considering those firms which export to at least one destination in at least one year in the time-span of our sample. Thus, the fraction of non-exporters in the population of all firms is much larger than the 22.62% reported here.

³⁵Similar evidence is provided in Eaton et al. (2004).

finding is in line with a recent literature that emphasizes learning across destinations (Albornoz et al. (2012)), which makes expansion from one destination to the next gradual, since firms learn about the appeal of their product to consumers in a given destination by exporting to nearby markets.

The previous table also shows that there is persistence in the number of relationships, since the probability of the number of relationships staying constant is much greater than the probability of the number of relationships changing from one year to the next. However, the transition matrix does not allow us to determine whether the identity of active relations is actually the same over time³⁶ and whether persistence is due to persistence of unobservable supply and demand shocks or due to state dependence of exporting decisions.

We now turn to a more formal analysis of persistence. To this end, we test for state dependence of exporting decisions by regressing the export status of a given firm on its past export status. Roberts and Tybout (1997) and Bernard and Jensen (2004) have interpreted state dependence of exporting as evidence for the presence of sunk costs. However, as outlined in the theory section, state dependence is equally consistent with learning models. We can improve upon the methodology employed by the previous authors, who only had information on firms' aggregate export status available, by exploiting the 3-dimensional nature of our panel. This allows us to control for both firm-time-specific (δ_{it}) and destination-time-specific (δ_{ct}) effects, i.e. for supply and demand shocks. We thus run the following linear probability model:

$$Y_{ict} = \beta_1 Y_{ict-1} + \delta_{it} + \delta_{ct} + u_{ict}, \tag{23}$$

for each sector, obtaining the results displayed in Table 13.³⁷ First, we find that state dependence is important in explaining firm export dynamics. The coefficient β_1 captures the effect of past export status on the current probability of exporting to a given destination. In fact, for an average sector, having exported to the same destination in the previous year increases the probability of exporting by around 67 percentage points (the coefficient is significant at the one per-cent level). This effect is similar for different sectors, ranging between 63% ("Electric and Electronic Equipment") and 72% ("Drugs, Soaps and Cleaners"). Second, even when controlling for past export status, the fraction of variance of export status unexplained by the model, while dropping substantially in magnitude (compare with Table 5), is still quite high. This is shown in Column (4) of the same table. The fraction of the

³⁶It may be that the number of export destinations remains constant but that the identity of export destinations changes. For example, a firm may export to Spain and Italy in 1995 and to Germany and Russia in 1996: in this case the transition matrix would report this observation on the diagonal since the number of active relations does not change from one year to the other.

³⁷We cluster standard errors at the firm-destination level.

variance of the dependent variable explained by the error term ranges from 32% ("Drugs, Soaps and Cleaners") to 48% ("Printing and Publishing"). This is further confirmation that it is important to model relationship-specific shocks in order to explain firm-level export dynamics convincingly.

Finally, we investigate whether state dependence of export relationships is systematically related to firm and destination characteristics, as both sunk-cost and learning theories would suggest. To shed light on this issue we therefore run the following linear probability model, interacting past export status with market size proxies, as well as firm productivity (measured as value added per worker):

$$Y_{ict} = \beta_1 Y_{ict-1} + \beta_2 Y_{ict-1} A_{ct} + \beta_3 Y_{ict-1} \phi_{it} + \delta_{it} + \delta_{ct} + u_{ict}, \tag{24}$$

where A_{ct} captures standard market size characteristics such as GDP, GPD per capita, and distance, while ϕ_{it} measures firm productivity. Results are reported in Table 14. We find that while the coefficient of past export status, β_1 , – which now measures the impact of past export status when all the interaction terms are zero – is negative, the interaction terms are strongly significant and have the expected signs.³⁸ The interactions with GDP, GDP per capita and productivity are positive, while the interaction with distance (which is an inverse measure of effective market size) is negative. While the precise coefficient magnitudes differ somewhat across sectors, the qualitative results are stable across sectors. These findings are expected and consistent with both the sunk-cost hypothesis and learning models.

5 Conclusions

In this paper we have documented several stylized facts on export dynamics using a panel of French firms. Most changes in entry and exit decisions to export destinations as well as adjustments in export values cannot be explained by firm-specific or destination-specific shocks. Instead, they are driven by relationship-specific shocks, i.e. shocks that hit a given firm in a given export destination. Moreover, export values are small at the beginning of an export episode and increase over time if the export relationship is successful. Finally, export decisions are state dependent – past export behaviour is an important predictor for current export status.

As we explain in the paper, the combination of these stylized facts is more in line with a learning model of exporting than with a dynamic Melitz model with sunk costs. We have sketched two versions of learning models: one where exporters have to learn the demand for their product in a specific

³⁸Evaluated at the mean of the interaction terms past export status still has a positive effect on the current export probability.

market and another one where exporters rely on local partners for exporting. Both versions yield very similar predictions and are thus consistent with the data. However, only in the second version of the learning model should institutions in the destination market matter for export behaviour. Better legal institutions makes it more likely that contracts between exporters and importers will be respected and therefore increases export values and the probability that an export relationship will survive from one period to the next. While we have not investigated this channel in the present paper, Aeberhardt et al. (2011) provide evidence for this specific mechanism. Still, in order to assess the relative importance of the sunk-cost model and the different versions of the learning model quantitatively, one would have to structurally estimate a hybrid model that nests all the previous models. This task is left for future research.

References

Aeberhardt, Romain, Ines Buono, and Harald Fadinger, "Learning, Incomplete Contracts and Export Dynamics: Theory and Evidence from French Firms," *INSEE working paper*, 2011.

Albornoz, Facundo, Héctor Calvo Pardo, Gregory Corcos, and Emanuel Ornelas, "Sequential Exporting," *Journal of International Economics*, 2012, forthcoming.

Araujo, Luis and Emanuel Ornelas, "Trust Based Trade," CEP Discussion Paper, 2007, 820.

_ , Giordano Mion, and Emanuel Ornelas, "Institutions and Export Dynamics," CEPR Discussion Paper, 2012, 8809.

Arkolakis, Konstantinos and Theodore Papageorgiou, "Selection, Growth and Learning," mimeo, 2009.

Baldwin, Richard and Paul Krugman, "Persistent Trade Effects of Large Exchange Rate Shocks," Quarterly Journal of Economics, 1989, 419, 635–654.

Berman, Nicolas, Antoine Berthou, and Jérôme Héricourt, "Export Dynamics and Sales at Home," CEPR Discussion Paper, 2011, 8684.

Bernard, Andrew and Bradford Jensen, "Why Some Firms Export," Review of Economics and Statistics, 2004, 86, 561–569.

_ , _ , Stephen Redding, and Peter Schott, "Firms in International Trade," Journal of Economic Perspectives, 2007.

- _ , Jonathan Eaton, Bradford Jensen, and Samuel Kortum, "Plants and Productivity in International Trade," American Economic Review, 2003, 93, 1268–1290.
- _, Stephen Redding, and Peter Schott, "Multiple-Product Firms and Product Switching,"

 American Economic Review, 2010, 100, 70–97.
- Berthou, Antoine and Lionel Fontagné, "How Do Multi-product Exporters React to a Change in Trade Costs?," Scandinavian Journal of Economics, 2012, forthcoming.
- Bricongne, Jean-Charles, Lionel Fontagn, Guillaume Gaulier, Daria Taglioni, and Vincent Vicard, "Exports and sectoral financial dependence evidence on French firms during the great global crisis," ECB Working paper, 2010, 1227.
- Buono, Ines and Guy Lalanne, "La raction des entreprises franaises la baisse des tarifs douaniers trangers," *Economie et Statistique*, 2011, 435-436, 13-29.
- _ and _ , "The effect of the Uruguay round on the intensive and extensive margins of trade," Journal of International Economics, 2012, 86, 269–283.
- Das, Sanghamitra, Mark Roberts, and James Tybout, "Market Entry Cost, Producer Heterogeneity and Export Dynamics," *Econometrica*, 2007, 75 (3), 837–873.
- **Dixit, Avinash**, "Entry and Exit Decisions under Uncertainty," *Journal of Political Economy*, 1989, 97 (3).
- Eaton, Jonathan, Marcela Eslava, C. J. Krizan, Maurice Kugler, and James Tybout, "A Search and Learning Model of Export Dynamics," mimeo, 2008.
- _ , _ , Maurice Kugler, and James Tybout, "Export Dynamics in Columbia: Firm-Level Evidence," NBER Working paper, 2007, 13531.
- _ , Samuel Kortum, and Francis Kramarz, "Dissecting Trade: Firms, Industries and Export Destinations," American Economic Review, Papers and Proceedings, 2004.
- _ , _ , and _ , "An Anatomy of International Trade: Evidence from French Firms," *Econometrica*, 2011, 79, 1453–1498.
- Lawless, Martina, "Firm Export Dynamics and the Geography of Trade," Journal of International Economics, 2009, 77, 245–254.

Melitz, Marc, "The Impact of Trade on Intra-Industry Reallocations and Aggregate Industry Productivity," *Econometrica*, 2003, 71, 695–1725.

Roberts, Mark and James Tybout, "The Decision to Export in Colombia: An Empirical Model of Entry with Sunk Costs," *American Economic Review*, 1997, 87, 545–564.

Rose, Andrew, "Do We Really Know That the WTO Increases Trade?," American Economic Review, 2004, 94, 98–114.

Ruhl, Kim and Jonathan Willis, "New Exporter Dynamics," mimeo, 2008.

Segura-Cayuela, Ruben and Josep Vilarrubia, "Uncertainty and Entry into Export Markets," Banco de Espana working paper, 2008, 0811.

A List of Countries

All the countries included in the analysis:

Antigua et Barbuda (AG), Albania (AL), Armenia (AM), Angola (AO), Argentina (AR), Austria (AT), Australia (AU), Azerbadjan (AZ), Barbados (BB), Bangladesh (BD), Burkina Faso (BF), Bulgaria (BG), Burundi (BI), Benin (BJ), Bolivia (BO), Brazil (BR), Botswana (BW), Belarus (BY), Belize (BZ), Canada (CA), Congo, Dem. Rep. (CD), Central African Republic (CF), Congo, Rep. (CG), Switzerland (CH), Cote d'Ivoire (CI), Chile (CL), Cameroon (CM), China (CN), Colombia (CO), Costa Rica (CR), Cuba (CU), Cyprus (CY), Czech Republic (CZ), Germany (DE), Denmark (DK), Dominique (DM), Dominican Republic (DO), Algeria (DZ), Ecuador (EC), Estonia (EE), Egypt (EG), Spain (ES), Ethiopia (excludes Eritrea) (ET), Finland (FI), Fiji (FJ), Gabon (GA), United Kingdom (GB), Granada (GD), Ghana (GH), The Gambia (GM), Guinea (GN), Equatorial Guinea (GQ), Greece (GR), Guatemala (GT), Guyana (GY), Guinea-Bissau (GW), Hong Kong (HK), Honduras (HN), Croatia (HR), Haiti (HT), Hungary (HU), Indonesia (ID), Ireland (IE), Israel (IL), India (IN), Islamic Rep. of Iran (IR), Island (IS), Italy (IT), Jamaica (JM), Jordan (JO), Japan (JP), Kenya (KE), Kyrgiz Republic (KG), Kazakhstan (KZ), Korea, Rep. (KR), Saint-Kitts and Nevis (KN), Cambodia (KH), Lebanon (LB), Santa-Lucia (LC), Sri Lanka (LK), Liberia (LR), Lesotho (LS), Lithuania (LT), Latvia (LV), Morocco (MO), Moldova (MD), Madagascar (MG), Macedonia (MK), Mali (ML), Macao (MO), Mauritania (MR), Malta (MT), Mauritius (MU), Malawi (MW), Mexico (MX), Malaysia (MY), Mozambique (MZ), Namibia (NA), Niger (NE), Nigeria (NG), Nicaragua (NI), Netherlands (NL), Norway (NO), Nepal (NP), New Zealand (NZ), Panama (PA), Peru (PE), Papua New Guinea (PG), Philippines (PH), Pakistan (PK), Poland (PL), Portugal (PT), Paraguay (PY), Romania (RO), Russian Federation (RU), Rwanda (RW), Seychelles (SC), Sweden (SE), Singapore (SG), Slovenia (SI), Slovak Republic (SK), Sierra Leone (SL), Senegal (SN), Sao Tome and Principe (ST), El Salvador (SV), Syrian Arab Republic (SY), Chad (TD), Togo (TG), Thailand (TH), Tunisia (TN), Turkey (TR), Trinidad and Tobago (TT), Taiwan (TW), Tanzania (TZ), Ukraine (UA), Uganda (UG), United States (US), Uruguay (UY), Uzbekistan (UZ), Saint-Vincent and the Grenadines (VC), Venezuela (VE), Vietnam (VN), Yemen (YE), South Africa (ZA), Zambia (ZM), Zimbabwe (ZW).

B Figures and Tables

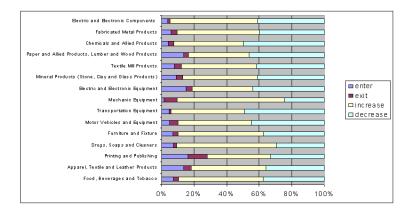


Figure 1: Intensive and extensive margin contributions to export growth by sector. The figure shows the contribution of new and destroyed relationships (extensive margin) and of continuing export relationships (intensive margin) to total sectoral export growth.

	Level	N	Mean	SD	Min	25th Pct.	Med.	75th Pct.	Max
export value (log)	firm - year - destination	871,794	10.4	2.3	0	8.7	10.3	11.9	22.2
productivity (log)	firm - year	122,680	5.6	0.5	0.5	5.3	5.6	5.9	12.5
number of destinations	firm - year	122,680	12	16.4	1	2	5	15	143
GDP (log)	country	144	17.3	2.1	11.8	15.7	17	18.9	22.9
GDP p.c. (log)	country	144	8.3	1.3	5.9	7.2	8.5	9.2	10.5
distance (log)	country	141	8.4	0.8	5.8	7.8	8.7	9.1	9.8

Table 1: Summary statistics. Data on export values are from Douanes. Export values are by firm-year-destination. Firm-level productivity is constructed from BRN data and is measured as value added per worker. Data on GDP and GDP per capita are from Penn World Tables (Mark 6.2), data on distance from Paris are from Rose (2004).

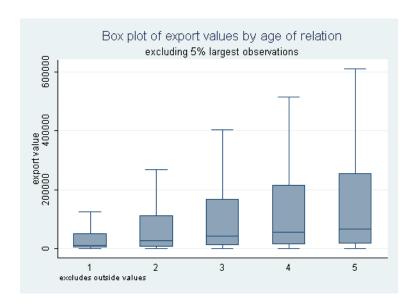


Figure 2: Export value distribution by relationship age. The figure shows box plots of export values on the vertical axis and relationship age on the horizontal axis.

year	number	new	as	ceasing	as	net entry into	as
	of exporters	exporters	percentage	exporters	percentage	export market	percentage
1995	18,382	-	-	-	-	=	=
1996	18,986	2,263	12%	1,659	9%	604	3%
1997	19,513	2,299	12 %	1,772	9%	527	3%
1998	19,950	2,164	11%	1,727	9%	437	2%
1999	19,996	2,003	10%	1,957	10%	46	0.2%

Table 2: Fluctuations in export status. The table presents entry and exit into exporting activity for the years 1995-1999.

year	relationships	created	as percentage	destroyed	as percentage	net creation	as percentage
1995	157,558	-	-	-	-	-	-
1996	167,279	43,629	27.7%	33,908	21.5%	9,721	6.2%
1997	177,513	45,715	27.3%	$35,\!481$	21.2%	10,234	6.1%
1998	183,595	44,721	25.1%	38,639	21.7%	6,082	3.4%
1999	185,849	43,394	23.6%	41,140	21.4%	2,254	1.2%

Table 3: Export relationships created and destroyed. The table reports creation and destruction of export relationships (defined by firm-destination) for the years 1995-1999.

	gross	net	enter	exit	increase	decrease
1996	0.100	0.011	7.3%	3.6%	48.1%	41.0%
1997	0.102	0.038	6.4%	2.7%	62.2%	28.8%
1998	0.097	0.016	5.0%	3.7%	53.4%	37.9%
1999	0.095	0.002	4.7%	3.5%	46.4%	45.4%

Table 4: Midpoint growth rates by year. Percentages explained by components.

ANOVA: fraction of variance of Y_{ict} explained by					
	δ_{ct}	δ_{it}	u_{ict}	Observations	
Food, Beverages and Tobacco	5.67%	23.91%	70.42%	1,569,500	
Apparel, Textile and Leather Products	9.09%	18.49%	72.43%	1,190,630	
Printing and Publishing	3.26%	15.33%	81.41%	1,407,440	
Drugs, Soaps and Cleaners	13.08%	21.88%	65.04%	405,880	
Furniture and Fixture	7.25%	19.93%	72.82%	1,301,590	
Motor Vehicles and Equipment	7.16%	21.38%	71.46%	334,340	
Transportation Equipment	7.89%	20.46%	71.46%	239,440	
Mechanic Equipment	7.13%	19.46%	73.42%	2,476,890	
Electric and Electronic Equipment	8.71%	21.01%	70.28%	1,005,940	
Mineral Products (Stone, Clay and Glass Products)	5.77%	23.22%	71.02%	724,160	
Textile Mill Products	9.68%	20.69%	69.63%	941,700	
Paper and Allied Products, Lumber and Wood Products	4.30%	20.04%	75.65%	1,222,020	
Chemicals and Allied Products	8.69%	22.15%	69.16%	1,662,210	
Fabricated Metal Products	4.51%	18.33%	77.16%	2,871,090	
Electric and Electronic Components	9.77%	21.47%	68.75%	558,450	

Table 5: Explaining the decision to export to a destination. ANOVA with linear probability model. The dependent variable is export status of firm i to destination c in period t. The explanatory variables are firm-time and destination-time dummies.

ANOVA: fraction of variance of C_{ict} explained by						
	δ_{ct}	δ_{it}	u_{ict}	Observations		
Food, Beverages and Tobacco	3.38%	3.52%	93.09%	1,255,600		
Apparel, Textile and Leather Products	6.40%	1.19%	92.41%	952,504		
Printing and Publishing	2.75%	4.21%	93.04%	$1,\!125,\!952$		
Drugs, Soaps and Cleaners	6.72%	0.00%	93.28%	324,704		
Furniture and Fixture	4.92%	2.46%	92.63%	1,041,272		
Motor Vehicles and Equipment	4.95%	1.74%	93.30%	$267,\!472$		
Transportation Equipment	5.29%	2.35%	92.36%	191,552		
Mechanic Equipment	4.63%	2.75%	92.62%	1,981,512		
Electric and Electronic Equipment	5.60%	2.46%	91.93%	804,752		
Mineral Products (Stone, Clay and Glass Products)	3.61%	3.08%	93.31%	579,328		
Textile Mill Products	5.99%	0.86%	93.15%	753,360		
Paper and Allied Products, Lumber and Wood Products	3.22%	3.66%	93.12%	977,616		
Chemicals and Allied Products	5.04%	1.45%	93.52%	1,329,768		
Fabricated Metal Products	3.15%	2.91%	93.94%	2,296,872		
Electric and Electronic Components	5.97%	1.10%	92.93%	446,760		

Table 6: Explaining changes in destination choice. ANOVA with linear probability model. The dependent variable is change in export status of firm i to destination c in period t. The explanatory variables are firm-time and destination-time dummies.

ANOVA: fraction of variance of	$f_{g_{ict}} \exp$	lained by		
Sample	δ_{ct}	δ_{it}	u_{ict}	Observations
excluding entry and exit	1.96%	15.47%	82.56%	536,777
including entry and exit	2.17%	15.50%	82.34%	863,404
by NES sector (excluding entry and exit)				
Food, Beverages and Tobacco	2.43%	19.94%	77.62%	44,583
Apparel, Textile and Leather Products	2.38%	18.00%	79.62%	40,709
Printing and Publishing	3.26%	30.72%	66.01%	17,436
Drugs, Soaps and Cleaners	4.55%	11.51%	83.93%	30,907
Furniture and Fixture	2.73%	17.62%	79.66%	38,622
Motor Vehicles and Equipment	6.86%	17.35%	75.79%	10,443
$Transportation \ Equipment$	6.61%	12.20%	81.19%	7,953
Mechanic Equipment	1.58%	13.21%	85.21%	78,901
Electric and Electronic Equipment	2.09%	13.49%	84.42%	39,007
Mineral Products (Stone, Clay and Glass Products)	4.04%	16.21%	79.75%	20,516
Textile Mill Products	2.55%	15.16%	82.29%	36.492
Paper and Allied Products, Lumber and Wood Products	2.93%	25.26%	71.81%	22,017
Chemicals and Allied Products	1.46%	14.91%	83.63%	68,338
Fabricated Metal Products	1.63%	22.75%	75.62%	56,718
Electric and Electronic Components	3.73%	12.41%	83.86%	24,135

Table 7: Explaining changes in export values. ANOVA. The dependent variable is midpoint growth rate export values of firm i to destination c in period t. The explanatory variables are firm-time and destination-time dummies.

1995-1996	Observations	average	median
All relations in 1995	157,558	640,997	28,084
Relations destroyed permanently in 1996	17,674	45,213	4,871
Occasionally destroyed relations (for 1996 only)	4,217	$54,\!867$	7,670
All relations in 1996	$167,\!279$	$630,\!214$	27,796
Relations created permanently in 1996	12,939	146,961	$13,\!266$
Occasionally created relations (for 1996 only)	13,929	24,125	4,193

Table 8: Average and median exports values (in Euros)

NES sector	Entry values by percentile				
	10th	30th	50th	70th	90th
Food, Beverages and Tobacco	963	2,911	6,522	15,314	56,098
Apparel, Textile and Leather Products	876	1,878	3,999	$9,\!497$	40,973
Printing and Publishing	659	1,497	3,129	8,583	$41,\!376$
Drugs, Soaps and Cleaners	1,095	3,201	7,359	19,742	84,750
Furniture and Fixture	866	1,877	3,927	9,783	36,942
Motor Vehicles and Equipment	882	2,577	6,715	16,362	85,082
Transportation Equipment	777	2,148	5,119	19,330	163,883
Mechanic Equipment	902	2,378	5,530	14,620	63,692
Electric and Electronic Equipment	955	2,470	5,793	13,980	$52,\!484$
Mineral Products (Stone, Clay and Glass Products)	953	2,347	5,971	14,779	60,394
Textile Mill Products	797	2,263	4,868	$11,\!697$	40,935
Paper and Allied Products, Lumber and Wood Products	892	2,928	6,636	14,695	$62,\!439$
Chemicals and Allied Products	979	2,456	5,723	14,023	54,882
Fabricated Metal Products	898	1,986	$4,\!514$	10,224	$38,\!354$
Electric and Electronic Components	854	1,953	5,151	14,729	63,951

Table 9: Entry values by percentile and 2-digit NES sector.

		Regression:	$q_{ict} = \beta_1 age_{ict}$	$+\delta_{it} + \delta_{ct} + u_{ict}$
	(1)	(2)	(3)	(4)
	age_{ict}	Adjusted R^2	Observations	Fraction of residual variance
Food, Beverages and Tobacco	0.51***	0.47	59,651	45%
	(0.011)			
Apparel, Textile and Leather Products	0.57***	0.40	54,733	52%
	(0.013)			
Printing and Publishing	0.45***	0.45	24,261	42%
	(0.017)			
Drugs, Soaps and Cleaners	0.48***	0.51	40,586	45%
	(0.012)			
Furniture and Fixture	0.46***	0.44	52,085	48%
	(0.011)			
Motor Vehicles and Equipment	0.57***	0.46	14,114	45%
	(0.027)			
Transportation Equipment	0.64***	0.42	10,719	49%
	(0.031)			
Mechanic Equipment	0.43***	0.35	107,002	57%
	(0.008)			
Electric and Electronic Equipment	0.43***	0.44	52,591	51%
	(0.012)			
Mineral Products (Stone, Clay and Glass Products)	0.48***	0.47	27,508	45%
	(0.016)			
Textile Mill Products	0.52***	0.41	48,838	52%
	(0.012)			
Paper and Allied Products, Lumber and Wood Products	0.55***	0.48	29,688	41%
	(0.015)			
Chemicals and Allied Products	0.50***	0.49	90,752	46%
	(0.008)			
Fabricated Metal Products	0.51***	0.46	76,480	44%
	(0.009)			
Electric and Electronic Components	0.52***	0.44	32,118	50%
	(0.016)			

Table 10: Intensive margin analysis with age (1). Excluding entry and exit. The dependent variable is the export value (in logs) of firm i to destination c in period t. The explanatory variables are the age of relationship, firm-time and destination-time dummies.

^{***:} significant at the 1% level.
Standard Errors, clustered at the firm-destination level, are reported in parenthesis.
(4) Fraction of variance of q_{ict} explained by u_{ict} .

		Regression: $g_{ict} = \beta_1 a g e_{ict} + \delta_{it} + \delta_{ct} + u_{ict}$					
	(1)	(2)	(3)	(4)			
	age_{ict}	Adjusted R^2	Observations	Fraction of residual variance			
Food, Beverages and Tobacco (excluding entry and exit)	-0.107***	0.09	44,583	77%			
Apparel, Textile and Leather Products	(0.007) -0.114*** (0.007)	0.09	40,709	79%			
Printing and Publishing	-0.059*** (0.012)	0.13	17,436	66%			
Drugs, Soaps and Cleaners	-0.087*** (0.008)	0.09	30,907	84%			
Furniture and Fixture	-0.099*** (0.008)	0.08	38,622	79%			
Motor Vehicles and Equipment	-0.134*** (0.015)	0.10	10,443	75%			
$Transportation \ Equipment$	-0.0571*** (0.020)	0.05	7,953	81%			
Mechanic Equipment	-0.037*** (0.006)	0.04	78,901	85%			
Electric and Electronic Equipment	-0.051*** (0.008)	0.06	39,007	84%			
Mineral Products (Stone, Clay and Glass Products)	-0.092*** (0.01)	0.05	20,516	79%			
Textile Mill Products	-0.123**** (0.008)	0.08	36,492	82%			
Paper and Allied Products, Lumber and Wood Products	-0.132*** (0.010)	0.08	22,017	71%			
Chemicals and Allied Products	-0.116*** (0.006)	0.07	68,338	83%			
Fabricated Metal Products	-0.117***	0.08	56,718	75%			
Electric and Electronic Components	-0.0973*** (0.010)	0.06	24,135	83%			

^{***:} significant at the 1% level.

Table 11: Intensive margin analysis with age (2). Excluding entry and exit. The dependent variable is the midpoint growth rates of exports of firm i to destination c in period t. The explanatory variables are age of relationship, firm-time and destination-time dummies.

1995/1996	0	1	2	3	4	5	(6 to 10)	(11 to 25)	25 or more
0	63.23	27.66	5.54	1.75	0.71	0.19	0.47	0.29	0.15
1	27.26	49.74	14.07	5.45	2.13	0.67	0.61	0.08	0.00
2	9.30	26.30	32.62	17.29	7.37	3.81	2.97	0.34	0.00
3	3.65	11.65	21.88	27.86	18.55	8.33	7.68	0.33	0.07
4	1.72	3.53	11.21	21.81	24.91	16.55	19.31	0.95	0.00
5	0.81	1.39	4.75	13.66	17.01	21.64	38.54	2.08	0.12
(6 to 10)	0.31	0.66	1.04	2.60	4.33	9.53	62.94	18.39	0.21
(11 to 25)	0.00	0.03	0.09	0.13	0.19	0.32	9.89	81.64	7.71
25 or more	0.00	0.07	0.00	0.00	0.00	0.21	0.07	6.49	93.17
exporters (in%)	22.62	20.53	9.59	6.80	4.80	3.79	11.97	13.34	6.57

Table 12: Transition matrix between 1995 and 1996. The rows (columns) refer to the number of destinations a given firm export in 1995 (1996). The cells show the fraction of export relationships that change from one category to the next between 1995 and 1996.

Standard Errors, clustered at the firm-destination level, are reported in parenthesis. (4) Fraction of variance of g_{ict} explained by u_{ict} .

		Regression: Y	$Y_{ict} = \beta_1 Y_{ict-1}$	$+\delta_{it} + \delta_{ct} + u_{ict}$
	(1)	(2)	(3)	(4)
	Y_{ict-1}	Adjusted R^2	Observations	Fraction of residual variance
Food, Beverages and Tobacco	0.70***	0.63	1,255,600	36.97%
	(0,002)			
Apparel, Textile and Leather Products	0.65***	0.57	952,504	42.41%
	(0,002)			
Printing and Publishing	0.65***	0.51	1,125,952	48.21%
	(0,004)			
Drugs, Soaps and Cleaners	0.72***	0.67	324,704	32.38%
	(0,003)			
Furniture and Fixture	0.66***	0.57	1,041,272	42.19%
	(0,003)			
Motor Vehicles and Equipment	0.68***	0.6	267,472	39.52%
	(0,005)			
Transportation Equipment	0.64***	0.57	191,552	42.60%
	(0,006)			
Mechanic Equipment	0.64***	0.56	1,981,512	43.83%
	(0,002)			
Electric and Electronic Equipment	0.63***	0.56	804,752	43.42%
	(0,003)			
Mineral Products (Stone, Clay and Glass Products)	0.70***	0.62	579,328	37.45%
· · · · · · · · · · · · · · · · · · ·	(0,004)			
Textile Mill Products	0.68***	0.62	753,360	38.15%
	(0,003)			
Paper and Allied Products, Lumber and Wood Products	0.69***	0.6	977,616	41.71%
	(0,003)			
Chemicals and Allied Products	0.70***	0.63	1,329,768	36.30%
	(0,002)			
Fabricated Metal Products	0.70***	0.59	2,296,872	40.80%
	(0,002)			
Electric and Electronic Components	0.67***	0.61	446,760	38.47%
•	(0,003)		,	
*** · · · · · · · · · · · · · · · · · ·	<u> </u>			

***: significant at the 1% level.
Standard Errors, clustered at the firm-destination level, are reported in parenthesis.
(4) Fraction of variance of Y_{ict} explained by u_{ict} .

Table 13: State dependence analysis. Linear probability model. The dependent variable is export status of firm ito destination c in period t. The explanatory variables are lagged export status, firm-time and destination-time dummies.

		Regression: $Y_{ict} = \beta_1$	$Y_{ict-1} + \beta_2 Y_{ict-1} A_{ct} + \beta_2 Y_{ict-1} A_{ct}$	$3Y_{ict-1}a_{it} + \delta_{it} + \delta_{ct}$	$+u_{ict}$			
	(1)	(2)	(3)	(4)	(5)	(6)	(7)	(8)
	Y_{ict-1}	$Y_{ict-1} \times \log(distance)$	$Y_{ict-1} \times \log(GDPp.c.)$	$Y_{ict-1} \times \log(GDP)$	$Y_{ict-1} \times \log(productivity)$	Adjusted R^2	Obs.	Residual Varianc
Food, Beverages and Tobacco	-0.196	-0.007	0.057	0.009	0.036	0.63	1,212,600	36.41%
	(0.040)***	(0.002)***	(0.003)***	(0.001)***	(0.003)***			
Apparel, Textile and Leather Products	0.051	-0.037	0.040	0.008	0.060	0.58	919,884	41.91%
	(0.044)	(0.002)***	(0.003)***	(0.002)***	(0.003)***			
Printing and Publishing	0.486	-0.050	0.013	0.008	0.043	0.52	1,087,392	47.62%
	(0.067)***	(0.004)***	(0.005)***	(0.003)***	(0.006)***			
Drugs, Soaps and Cleaners	0.314	-0.008	0.026	-0.001	0.040	0.67	313,584	32.24%
	(0.048)***	(0.002)***	(0.003)***	(0.002)***	(0.004)***			
Furniture and Fixture	-0.096	-0.030	0.049	0.014	0.043	0.58	1,005,612	41.55%
	(0.046)**	(0.002)***	(0.003)***	(0.002)***	(0.004)***			
Motor Vehicles and Equipment	-0.027	-0.035	0.027	0.022	0.050	0.61	258,312	38.82%
	(0.091)	(0.005)***	(0.006)***	(0.003)***	(0.009)***			
Transportation Equipment	-0.497	-0.017	0.041	0.026	0.066	0.58	184,992	41.59%
	(0.104)***	(0.005)***	(0.006)***	(0.003)***	(0.012)***			
Mechanic Equipment	-0.300	-0.025	0.034	0.200	0.073	0.57	1,913,652	43.12%
ž <u>ř</u>	(0.035)***	(0.002)***	(0.002)***	(0.001)***	(0.004)***			
Electric and Electronic Equipment	-0.411	-0.017	0.038	0.026	0.054	0.57	777,192	42.69%
i	(0.047)***	(0.002)***	(0.003)***	(0.002)***	(0.005)***			
Mineral Products (Stone, Clay and Glass Products)	-0.098	-0.026	0.031	0.021	0.050	0.63	559,488	36.82%
	(0.066)	(0.003)***	(0.004)***	(0.002)***	(0.006)***			
Textile Mill Products	-0.198	-0.020	0.055	0.013	0.046	0.62	727,560	37.67%
	(0.048)***	(0.002)***	(0.004)***	(0.002)***	(0.004)***			
Paper and Allied Products, Lumber and Wood Products	0.018	-0.026	0.018	0.017	0.062	0.58	944,136	41.34%
	(0.067)	(0.003)***	(0.004)***	(0.002)***	(0.006)***			
Chemicals and Allied Products	-0.011	-0.018	0.024	0.016	0.056	0.64	1,284,228	35.98%
	0.034	(0.002)***	(0.002)***	(0.001)***	(0.003)***			
Fabricated Metal Products	0.110	-0.030	0.037	0.013	0.038	0.59	2,218,212	40.25%
	(0.041)***	(0.002)***	(0.003)***	(0.001)***	(0.005)***			
Electric and Electronic Components	-0.205	-0.022	0.036	0.024	0.044	0.62	431,460	37.88%
•	(0.063)***	(0.003)***	(0.004)***	(0.002)***	(0.007)***		,	

***: significant at the 1% level.
Standard Errors, clustered at the firm-destination level, are reported in parenthesis.
(8) Fraction of variance of Y_{ict} explained by u_{ict} .

Table 14: State dependence analysis with interaction terms. Linear probability model. The dependent variable is the export status of firm i to destination c in period t. The explanatory variables are lagged export status, interactions of lagged export status with distance, GDP, GDP per capita and productivity, as well as firm-time and destination-time dummies.

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2010

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2011

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2012

A. ACCETTURO and G. DE BLASIO, *Policies for local development: an evaluation of Italy's "Patti Territoriali"*, Regional Science and Urban Economics, v. 42, 1-2, pp. 15-26, **TD No. 789** (January 2006).

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