

Learning about Bond Prices

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June 18, 2026

Learning and Bond price volatility:

- Many studies on stock price volatility, fewer on bond price volatility
- We learnt a lot about stock price behavior using models where agents learn about prices: *internal rationality*.
- Apply to learning about bond prices

Roadmap

- 1 Yields Data, Some Puzzles
- 2 Literature
- 3 A Model
- 4 Internal Rationality
- 5 A Simple Model of Learning about the Yield Curve
- 6 Model Fit

- Nominal US government bonds zero-coupon yields
- Sample 1983:12 to 2019:12.
- Monthly frequency
- Surveys: Reuters' Blue Chip Financial Forecasters survey on expected bond yields.

Flat Term structure of Yield Volatilities

Yearly maturities, leads, lags, forecasting horizons.

Table 1: Yields 1st and 2d moments

	1y	2y	3y	5y	7y	10y
$\hat{Y}^n$	379	407	431	471	500	532
$\hat{\sigma}_{Y^n}$	280	284	280	269	260	247
$\widehat{corr}(Y^n)_{-1y}$	0.89	0.90	0.91	0.92	0.92	0.92

Hard to reconcile with RE and reasonable serial correlation of expected inflation.

Failure of Expectations Hypothesis

$$Y_{t+1y}^{n-1} - Y_t^n = \alpha^n + \beta^n \frac{Y_t^n - Y_t^1}{n-1} + U_t^n \quad (1)$$

EH says β 's ≈ 1

Table 2: Failure of Expectations Hypothesis

	2y	3y	5y	7y	10y
$\widehat{\beta}^n$	-0.31	-0.50	-0.99	-1.28	-1.67

horizon= 1y here and in all tables below

Failure of EH: "Modern" version

Let Q_t^n be the price of a zero-coupon bond of maturity n .

Define excess bond returns

$$xr_{t+1y}^n = \log \left(\frac{Q_{t+1y}^{n-1}}{Q_t^n} \right) - \log \left(\frac{1}{Q_t^1} \right)$$

Or, in terms of log yields $Y_t^n \equiv -\frac{\log(Q_t^n)}{n}$

$$xr_{t+1y}^n = -(n-1)Y_{t+1y}^{n-1} + nY_t^n - Y_t^1$$

Failure of EH: "Modern" version

EH can also be expressed in terms of xr forecasts:

$$xr_{t+1y}^n = \alpha_n^{xr} + b_n^{xr} \frac{Y_t^n - Y_t^1}{n-1} + U_t^n$$

EH implies b^{xr} 's ≈ 0 .

Table 3: Failure of EH "modern" version

	2y	3y	6y	8y	11y
$\widehat{b}_n^{xr}$	1.37	3.1	10.84	15.53	24.61
$t - stat$	0.52	1.14	2.40	2.87	3.54
R^2	0.32	1.59	6.00	7.88	10.41

High volatility of excess bond returns

$$xr_{t+1y}^n = \log \left(\frac{Q_{t+1y}^{n-1}}{Q_t^n} \right) - \log \left(\frac{1}{Q_t^1} \right)$$

Table 4: Excess Returns volatility

maturity n	2y	3y	5y	7y	10y
$\widehat{\sigma}_{xr^n}$	127	239	521	676	895

Hard to explain with RE.

Excess volatility of yields, Intuition for RE failure

$$Q_t^n = Q_t^1 E_t^{\mathbb{Q}}(Q_{t+1}^{n-1}) = E_t^{\mathbb{Q}}\left(\prod_{i=0}^{n-1} Q_{t+i}^1\right)$$

where $\mathbb{Q}$ refers to "risk-neutral" distribution.

Taking log-linear approximation

$$Y_t^n \approx E_t^{\mathbb{Q}}\left(\frac{\sum_{i=0}^{n-1} Y_{t+i}^1}{n}\right)$$

therefore we expect $\sigma_{Y_t^n} < \sigma_{Y_t^1}$

Excess volatility of excess bond returns, Intuition

Furthermore

$$\begin{aligned}xr_{t+1}^n &= nY_t^n - (n-1)Y_{t+1}^{n-1} - Y_t^1 \\ &\approx \left(\sum_{i=1}^{n-1} (E_t^{\mathbb{Q}} - E_{t+1}^{\mathbb{Q}})(Y_{t+i}^1) \right)\end{aligned}$$

so $\sigma_{xr_t^n}^2$ variance of a change in expectations, so it should be small.

Bond survey expectations

- Much recent work on surveys about asset prices
- Giacoletti, Laursen and Singleton (2020) and Singleton (2021) AFA Presidential Address (GLS) :
 - Describe investor expectations about bond yields
 - Investors (consensus) bond yield expectations are on average similar to those arising from best forecasts of an affine model with principal components

Adam, Marcet, Beutel (2017)

Let $\mathcal{E}$ denote survey expectations.

Consider a regression like EH but with expectations in left hand side

$$\mathcal{E}_t(xr_{t+1y}^n) = \alpha_n^{\mathcal{E}} + b_n^{\mathcal{E}} \frac{Y_t^n - Y_t^1}{n-1} + U_t^{n,\mathcal{E}}$$

Under RE $b_n^{xr} = b_n^{\mathcal{E}}$.

Related to Coibion and Gorodnichenko approach.

For $\mathcal{E}_t(xr_{t+1y}^n)$ interpolate BCFF to form monthly expectations 1 year ahead.

$$xr_{t+1y}^n = \alpha_n^{xr} + b_n^{xr} \frac{Y_t^n - Y_t^1}{n-1} + U_t^{n,xr}$$

$$\mathcal{E}_t(xr_{t+1y}^n) = \alpha_n^{\mathcal{E}} + b_n^{\mathcal{E}} \frac{Y_t^n - Y_t^1}{n-1} + U_t^{n,\mathcal{E}}$$

Table 5: RE Regressions

	2y	3y	6y	8y	11y
$\widehat{b}_n^{xr}$	1.37	3.1	10.84	15.53	24.61
$\widehat{b}_n^{\mathcal{E}}$	0.89	1.61	3.02	1.87	3.91
t-stat $b_n^{xr} = b_n^{\mathcal{E}}$	0.37	0.76	1.75	2.47	2.70

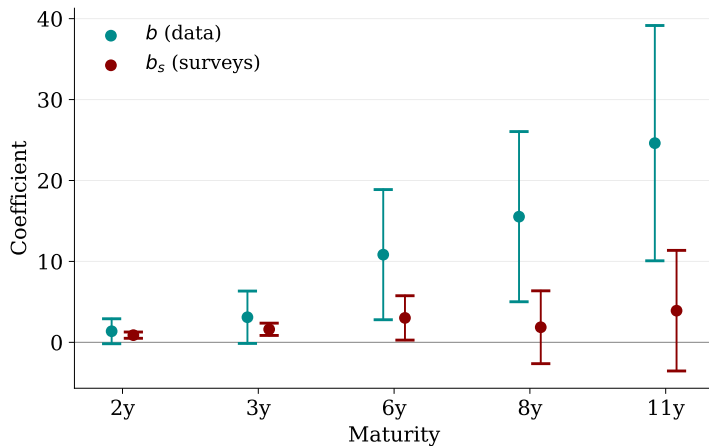
$\widehat{b}_n^{xr} > 0$ and $\widehat{b}_n^{xr} > \widehat{b}_n^{\mathcal{E}}$.

Both $\widehat{b}_n^{xr}$ and $\widehat{b}_n^{xr} - \widehat{b}_n^{\mathcal{E}}$ increase with n .

Surveys underestimate the role of the slope in predicting xr_{t+1y}^n

t - stat rejects RE, more strongly for long maturities

RE Regressions



RE Regressions, additional versions

Test is valid for any regressor in the information set at t .

We get a similar picture with alternative regressors that capture the slope:

- PC_t^2 : 2d Estimated Principal Component

- $\hat{s}_t$: Investor's estimate of underlying slope according to our model

RE Regressions, additional versions

- PC_t^2 : 2d Estimated Principal Component

- $\hat{s}_t$: Investor's estimate of underlying slope according to our model

Table 6: RE Regressions

regressor		2y	3y	6y	8y	11y
PC_t^2	$\hat{b}_n^{xr}$	0.06	0.13	0.43	0.61	0.89
	$\hat{b}_n^{\mathcal{E}}$	0.02	0.05	0.14	0.13	0.19
	t-stat $b_n^{xr} = b_n^{\mathcal{E}}$	0.85	1.11	1.98	2.57	2.87
$\hat{s}_{t-3}$	$\hat{b}_n^{xr}$	1.04	2.16	9.14	13.90	22.67
	$\hat{b}_n^{\mathcal{E}}$	0.41	0.57	0.28	-2.28	-2.86
	t-stat $b_n^{xr} = b_n^{\mathcal{E}}$	0.77	0.95	2.12	2.79	3.08
$\hat{s}_t, \hat{s}_{t-3}$	p-value $b_n^{xr} = b_n^{\mathcal{E}}$	0.70	0.59	0.09	0.01	0.00

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- Standard equilibrium asset pricing models have a hard time matching all previous facts.
- High volatility of xr documented since Shiller (1979) but very rarely addressed.
- Much progress has been made:
 - Under RE homogeneous agents: Time-varying reward for risk:
Dai and Singleton (2002), Wachter (2006), Gabaix (2008), Bansal and Shaliastovich (2013)
 - Bayesian/RE: heterogeneous beliefs but they know the pricing function:
Xiong and Yan (2010), Buraschi and Whelan (2022), Buraschi and Jiltsov (2007), Ehling Gallmeyer, Heyerdahl-Larsen and Illeditsch (2018).

- The old puzzles of Tables 1-4 remain hard to explain jointly.
Giglio et al (2018) argue standard models can't explain all those observations.
- The RE regressions in Tables 5-6 are a new puzzle.
- It matters whether the data needs RE or learning!!

Literature, Adaptive Learning, partial information and bond yields

- Sinha (2009), Piazzesi and Schneider (2006), Laubach, Tetlow, and Williams (2007) models of adaptive learning about how macro variables or inflation influence bond prices.
- Ellison Tischbirekz (2021) beauty contest.

This Project

Introduce *Internal Rationality*.

- investors have a *consistent system of beliefs about bond prices*
- investors are *fully rational* given their beliefs.

Agents' beliefs are "almost correct"

- close to RE beliefs
- model converges to RE
- "close" to the true distribution, (hard to reject using observed data)
- investors' beliefs match behavior of expectations surveys

Claim: IR is a "minor" deviation from standard paradigm:

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A model of bond prices

N zero-coupon nominal bonds of maturity $n = 1, \dots, N$

Bond prices $Q_t = [Q_t^1, \dots, Q_t^N]'$.

Homogeneous investors with $E_0^{\mathcal{P}} \sum_{t=0}^{\infty} \delta^t u(c_t)$

Investor beliefs represented by $\mathcal{P}$.

All bonds with $n > 1$ are resold one period after being purchased. Budget constraint:

$$P_t c_t + \sum_{n=1}^N Q_t^n B_t^n = w_t P_t + \sum_{n=1}^N Q_t^{n-1} B_{t-1}^n$$

$$Q_t^0 = 1$$

A model of Bond Prices

Assume bond limits:

$$\overline{B} \geq B_t^n \geq \underline{B} \quad (2)$$

Inflation process:

$$\log P_t - \log P_{t-1} \equiv \pi_t = \mu(1 - \rho) + \rho\pi_{t-1} + \eta_t^\pi$$

$$\eta_t^\pi \sim \mathcal{N}(0, \sigma_{\eta^\pi}^2), \text{ iid.}$$

Equilibrium conditions:

$$Q_t^n = \delta E_t^{\mathcal{P}} \left(\frac{u'(c_{t+1})}{u'(c_t)} \frac{P_t}{P_{t+1}} Q_{t+1}^{n-1} \right) \text{ for } n \geq 2$$

The short rate Y_t^1 is exogenous, given by

$$Y_t^1 = -\log(\delta) + \pi_t^e + \xi_t$$

ξ_t a mean-zero exogenous process (policy shock).

So bond limit is binding for $n = 1$ but assumed not to bind for $n > 1$.

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Investors' probability space under IR

$$\Omega \equiv \Omega_Q \times \Omega_{\pi, \xi}$$

where Ω_Q contains possible histories of $Q = [Q^n]_{n=1}^N$, and $\Omega_{\pi, \xi}$ contains possible histories of inflation π, ξ .

Investors have a belief system $\mathcal{P}$ with (consistent) probabilities on Ω .

RE is a special case when agents' beliefs about π, ξ are correct and they know the true pricing function determining Q .

Bayesian/RE a special case when agents' don't know everything about π, ξ but they know the true pricing function determining Q .

Model with learning about Q

However if agents don't know the pricing function prices are given by

$$Q_t^n = \delta E_t^{\mathcal{P}} \left(\frac{u'(c_{t+1})}{u'(c_t)} \frac{P_t}{P_{t+1}} Q_{t+1}^{n-1} \right) \text{ for } n \geq 2 \quad (3)$$

Need to state investors' beliefs about **joint** distribution of Q, π .

A common view:

”Investors with beliefs about bond prices Q are irrational.”

The reasoning:

”If investors are rational they can apply LIE to their own optimality conditions to obtain:

$$Q_t^n = \delta^n E_t^{\mathcal{P}} \left(\frac{u'(c_{t+n})}{u'(c_t)} \frac{P_t}{P_{t+n}} \right)$$

Knowledge of inflation and output process means

$$Q_t^n = \delta^n E_t \left(\frac{u'(c_{t+n})}{u'(c_t)} \frac{P_t}{P_{t+n}} \right) = Q_t^{n,RE}$$

This contradicts other price beliefs.

Therefore agents with beliefs about Q are irrational.”

But this common view is wrong

- ❶ If some agents are constrained LIE can not be applied with knowledge about inflation and output, it needs market knowledge.
(Adam, Marcet 2011)

- ❷ Under IR

$$Q_t^n = \delta^n E_t^{\mathcal{P}} \left(\frac{u'(c_{t+n}(Q^{t+n}))}{u'(c_t)} \frac{P_t}{P_{t+n}} \right) \neq Q_t^{n,RE}$$

so $E_t^{\mathcal{P}}$ depends on beliefs about Q .

(Adam, Marcet, Beutel, 2017)

Under IR investors' price beliefs are consistent with their information.

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Special case

- π unit root: ($\rho = 1$)
- Policy shock ξ iid, independent of π
- Risk neutrality
- Yearly model frequency
- Log linear approximation to Euler equation

A belief system about bond prices

We now specify our assumptions about the system of beliefs.
To build a "good" system of beliefs, we assume

- 1 Agents know process for (π, ξ) .
- 2 Agents' know yields at nearby maturities are related.
They express this belief with a perceived yield curve

$$Y_t^n = -\log \delta + \pi_t^e + \mathbf{s}_t n + u_t^n \quad (4)$$

- Investors know the intercept and how inflation affects yields
- but agents believe there is a time-varying, unobserved slope $\mathbf{s}_t$

(in blue we show equations that hold according to investors beliefs)

A Simple Model

Recall perceived yield curve

$$Y_t^n = -\log \delta + \pi_t^e + \mathbf{s}_t n + u_t^n$$

To complete investors' beliefs assume investors believe

$$\mathbf{s}_t = \mathbf{s}_{t-1} + w_t$$

Optimal filtering implies

$$E_t^{\mathcal{P}}(\mathbf{s}_t) \equiv \widehat{\mathbf{s}}_t = \widehat{\mathbf{s}}_{t-1} + g(s_{t-1} - \widehat{\mathbf{s}}_{t-1})$$

g is the "gain" parameter.

s_t is the best estimate of period- t slope given new info at t :

$$s_t = E_t^{\mathcal{P}}(s_t \mid Y_t^N, \dots, Y_t^1)$$

to be specified later. Then

$$E_t^{\mathcal{P}}(Y_{t+1}^n) = -\log \delta + \pi_t^e + \widehat{\mathbf{s}}_t n$$

Beliefs "near-rational"

Recall perceived yield curve:

$$Y_t^n = -\log \delta + \pi_t^e + \widehat{\mathbf{s}}_t n + \widehat{u}_t^n$$

Simple algebra, actual yield curve:

$$Y_t^n = -\log \delta + \pi_t^e + \widehat{\mathbf{s}}_t n \left(\frac{n-1}{n} \right)^2$$

(in brown equations that hold under the true distribution in the learning equilibrium)

Agents' beliefs are "quite good":

- recall agents know process for π, ξ
- when agents believe there is a time-varying slope there is a time-varying slope
- model converges to RE:

A simple estimate s_t (the perceived slope at t)

Investors should use all maturities Y_t^n to estimate $s_t = E_t^{\mathcal{P}}(s_t \mid Y_t^N, \dots, Y_t^1)$.

Additional assumption on beliefs:

Investors think there is a reference bond of maturity m with lower uncertainty

$$\text{var}^{\mathcal{P}}(u_t^m) \ll \text{var}^{\mathcal{P}}(u_t^n)$$

Then, GLS says it is optimal to use only $n = m, 1$ to obtain information about s_t :

$$s_t = \frac{Y_t^m - Y_t^1}{m - 1}$$

Learning Equilibrium with "simple" estimate s_t

Plugging in equilibrium yields

$$\begin{aligned}Y_t^n &= -\log(\delta) + \widehat{s}_t n \left(\frac{n-1}{n} \right)^2 + \pi_t^e \\s_t &= \frac{Y_t^m - Y_t^1}{m-1} \\\widehat{s}_t &= \widehat{s}_{t-1} + g(s_{t-1} - \widehat{s}_{t-1})\end{aligned}$$

So, under self-referential learning:

$$\uparrow s_t = \frac{Y_t^m - Y_t^1}{m-1} \implies \uparrow \widehat{s}_{t+1} \implies \uparrow Y_{t+1}^m \implies \uparrow s_{t+1} = \frac{Y_{t+1}^m - Y_{t+1}^1}{m-1} \dots$$

Convergence to RE: The T-map

When agents' perceived expectation (PLM) is

$$E_t^{\mathcal{P}} \frac{Y_{t+1}^m - Y_{t+1}^1}{m-1} = \widehat{s}_t$$

then the true expectation (ALM) is

$$E_t \frac{Y_{t+1}^m - Y_{t+1}^1}{m-1} \equiv \frac{m-1}{m} \widehat{s}_t \equiv T(\widehat{s}_t)$$

Convergence to RE

We have $T' < 1$.

Standard theorems in the learning literature imply that model converges to RE:

$$Y_t \rightarrow Y_t^{RE}$$

if either

- OLS instead of constant gain ($g = 1/t$) or
- $g \rightarrow 0$

(Marcet and Sargent (1989), Williams (2017), Evans Honkapohja (2002))

Fluctuations around RE

Furthermore, for $g > 0$, plugging in equilibrium value of s_t

$$\widehat{s}_t = \left(1 - \frac{g}{m}\right) \widehat{s}_{t-1} - g \frac{\xi_{t-1}}{m-1}$$

So $\widehat{s}_t$ very persistent and $E(\widehat{s}_t) = 0$.

The learning model fluctuates around RE.

An Analytic Solution

A linear system of difference equations

$$\begin{aligned}\widehat{\mathbf{s}}_t &= \left(1 - \frac{g}{m}\right) \widehat{\mathbf{s}}_{t-1} - g \frac{\xi_{t-1}}{m-1} \\ Y_t^n &= -\log(\delta) + \widehat{\mathbf{s}}_t \frac{(n-1)^2}{n} + \pi_t^e \\ x r_{t+1}^n &= \left[2n - 3 + (n-2)^2 \frac{g}{m}\right] \widehat{\mathbf{s}}_t \\ &\quad + (n-1) \eta_{t+1}^\pi + \left[\frac{(n-2)^2 g}{m-1} - 1\right] \xi_t \\ E_t^{\mathcal{P}}(x r_{t+1}^n) &= -\xi_t\end{aligned}$$

$\text{var}(\widehat{\mathbf{s}})$ contributes to variances, more for higher maturities.

- $\text{var}(Y^n)$ has additional variance for long n .
- $\text{var}(xr^n)$ increases with n .
- Consistent with RE regression:
 - $\widehat{b}_n^{xr} > 0$
 - $\widehat{b}_n^{xr}$, increases with n
 - $\widehat{b}_n^{xr} > \widehat{b}_n^{\mathcal{E}}$
 - $\widehat{b}_n^{xr} - \widehat{b}_n^{\mathcal{E}}$ increases with n

Intuition for RE regressions

Recall PLM and ALM:

$$E_t^{\mathcal{P}}(Y_{t+1}^n) = -\log \delta + \pi_t^e + \widehat{\mathbf{s}}_t n$$

$$E_t(Y_{t+1}^n) = -\log \delta + \pi_t^e + \widehat{\mathbf{s}}_t n \left(\frac{n-1}{n} \right)^2 \left(1 - \frac{g}{m} \right)$$

So agents *over*predict the influence of $\widehat{\mathbf{s}}_t$ on future yields.

Agents *under*predict the influence of $\widehat{\mathbf{s}}_t$ on future

$$xr_{t+1}^n = -(n-1)Y_{t+1}^{n-1} + nY_t^n - Y_t^1$$

Consistent with $\widehat{b}_n^{xr} > \widehat{b}_n^{\mathcal{E}}$ in the data.

Intuition for RE regressions: general property of IR

Intuition for $b_n^{\mathcal{E}} \approx 0$

- EH holds, approximately, for the agents' belief system

Intuition for $b_n^{xr} > b_n^{\mathcal{E}}$

- Define T as $E_t(Y_{t+1}^n) = T(E_t^{\mathcal{P}}(Y_{t+1}^n))$
- Recall, for stability we need $T' < 1$
- if $T' < 1$ *actual* $E_t(Y_{t+1}^n)$ moves less with slope than $E_t^{\mathcal{P}}(Y_{t+1}^n)$
- Hence *actual* $E_t(xr_{t+1}^n)$ moves more with slope than $E_t^{\mathcal{P}}(xr_{t+1}^n)$
- therefore $b_n^{xr} > b_n^{\mathcal{E}}$ if model is stable.

Therefore $\widehat{b}_n^{xr}, \widehat{b}_n^{\mathcal{E}}$ consistent with IR and standard Euler Eq.

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Fitting a simple Calibrated model

- AR(1) inflation with $\rho < 1$.
- Calibrate to sample values $\sigma_{\pi^e}, \rho, \sigma_{\eta^\pi}, \sigma_\xi$ for $\xi = Y^1 - \pi^e$.
- Fiddled with with the free parameter g , selected $g = 0.3$ (yearly)

Second Moments RE and Learn for the Simpler Model

	3y			10y		
	data	RE	Learn	Data	RE	Learn
σ_{Y^n}	280	90	98	247	73	240
$\rho_{Y^n, -1y}$	0.91	0.94	0.94	0.92	0.94	0.97
σ_{xr^n}	239	214	224	895	295	622
b_n^{xr}	3.1	0	2.02	24.61	0	7.01
$b_n^{\mathcal{E}}$	1.61	0	1.57	3.91	0	3.97

Model III: π, ξ bivariate AR(1)

Now introduce

- Monthly model but yearly horizons and lags
- (π, ξ) bivariate AR(1), estimated to match observed moments (at yearly lags)
- Estimate g by MSM

Model III Moments

Statistic	Maturity	Data	Model	t-stat
Standard deviation yields σ_{Y^n}	1y	280.00	296.11	-0.72
	2y	284.04	275.48	0.36
	3y	280.63	258.46	0.91
	5y	269.98	246.55	0.95
	7y	260.52	267.36	-0.27
	10y	247.59	340.35	-3.85
Serial correlation yields ρ_{Y^n}	1y	0.89	0.91	-0.96
	2y	0.90	0.92	-0.66
	3y	0.91	0.92	-0.36
	5y	0.92	0.92	-0.29
	7y	0.92	0.94	-1.01
	10y	0.92	0.95	-1.79
Standard deviation excess returns σ_{x^n}	2y	127.11	101.27	2.23
	3y	239.72	201.37	1.91
	5y	521.33	384.68	2.77
	7y	676.19	571.83	1.58
	10y	895.49	917.47	-0.24

$$\hat{g}=0.0536$$

Model III: RE Regressions

Statistic	Maturity	Data	Model	t-stat
EH coefficient b_{xr}^n	2y	1.37	0.18	1.52
	3y	3.10	1.38	1.04
	6y	10.84	5.06	1.41
	8y	15.53	7.77	1.45
	11y	24.61	12.39	1.65
RE coefficient b_s^n	2y	0.89	0.09	4.05
	3y	1.61	0.12	3.84
	6y	3.02	0.18	2.03
	8y	1.87	0.19	0.73
	11y	3.91	0.19	0.98

Possible improvements

***- risk aversion

- a more elaborate perceived yield curve
- disagreement
- learning about factors

Conclusion

Surveys underpredict the effect of the slope on future yields.

We lay out a simple model of learning about the yield curve.

Improves remarkably fit of moments.

Explains RE regressions

$$\begin{aligned}Q_t^{RE,n} &= \delta^n E_t \left(\frac{P_t}{P_{t+n}} \right) \\Y_t^{RE,n} &= -\log \delta + \mu + \phi^n (\pi_t - \mu) \\ \phi^n &\equiv \frac{\sum_{j=1}^n \rho^j}{n}\end{aligned}$$

$$xr_{t+1}^n = \eta_{t+1}^\pi \sum_{j=1}^{n-1} \rho^j - \xi_t$$

Analytic formulae for simplest case $\rho = 1$

$$\sigma_{xr^n}^2 = \left[2n - 3 + (n - 2)^2 \frac{g}{m} \right]^2 \sigma_{\hat{s}}^2 \quad (5)$$

$$+ (n - 1)^2 \sigma_{\eta^\pi}^2 + \left[\frac{(n - 2)^2 g}{m - 1} - 1 \right]^2 \sigma_{\xi}^2 \quad (6)$$

$$\begin{aligned} b_n^{xr} &= \frac{\text{cov}(xr_{t+1}^n, \frac{Y_t^n - Y_t^1}{n-1})}{\sigma_s^2} \\ &= \frac{\left[2n - 3 + (n - 2)^2 \frac{g}{m} \right] \frac{n-1}{n} \sigma_{\hat{s}}^2 - \left(\frac{(n-2)^2 g}{m-1} - 1 \right) \frac{\sigma_{\xi}^2}{n-1}}{\left(\frac{n-1}{n} \right)^2 \sigma_{\hat{s}}^2 + \frac{\sigma_{\xi}^2}{(n-1)^2}} \end{aligned} \quad (7)$$

Analytic formulae for simplest case $\rho = 1$

For general π, ξ

$$b_n^{xr} - b_n^{\mathcal{E}} = \frac{(n-1)\text{cov}(E_t^{\mathcal{P}}(Y_{t+1}^{n-1}) - Y_{t+1}^{n-1}, s_t)}{\sigma_s^2},$$

$\text{cov}(E_t^{\mathcal{P}}(Y_{t+1}^n) - Y_{t+1}^n, s_t) > 0$, therefore $b_n^{xr} - b_n^{\mathcal{E}} > 0$.

For simplest case

$$b_n^{\mathcal{E}} = \frac{\sigma_{\xi}^2}{\frac{(n-1)^3}{n^2}\sigma_{\hat{s}}^2 + \frac{1}{n-1}\sigma_{\xi}^2}$$

Model II: AR(1) but independent π and ξ

- Monthly model but yearly horizons and lags
- Serially correlated bivariate π and ξ each AR(1) but independent of each other, to match observed moments (at yearly lags)
- Estimate g by MSM

Model II Moments

Statistic	Maturity	Data	Model	t-stat
Standard deviation yields σ_{Y^n}	1y	280.00	273.75	0.28
	2y	284.04	274.94	0.38
	3y	280.63	277.47	0.13
	5y	269.98	285.35	-0.62
	7y	260.52	296.43	-1.42
	10y	247.59	318.00	-2.92
Serial correlation yields ρ_{Y^n}	1y	0.89	0.87	0.65
	2y	0.90	0.87	1.41
	3y	0.91	0.87	2.08
	5y	0.92	0.87	2.79
	7y	0.92	0.88	2.59
	10y	0.92	0.89	2.39
Standard deviation excess returns σ_{xT^n}	2y	127.11	104.86	1.92
	3y	239.72	213.80	1.29
	5y	521.33	433.49	1.78
	7y	676.19	655.58	0.31
	10y	895.49	994.95	-1.08

$$\hat{g}=0.0995$$

Model II: RE Regressions

Statistic	Maturity	Data	Model	t-stat
EH coefficient b_{xr}^n	2y	1.37	1.80	-0.54
	3y	3.10	4.22	-0.68
	6y	10.84	10.51	0.08
	8y	15.53	14.62	0.17
	11y	24.61	20.82	0.51
RE coefficient b_s^n	2y	0.89	0.04	4.28
	3y	1.61	0.06	4.00
	6y	3.02	0.11	2.08
	8y	1.87	0.14	0.75
	11y	3.91	0.18	0.98

Model III with RE: Moments

Statistic	Maturity	Data	Model	t-stat
Standard deviation yields σ_{Y^n}	1y	280.00	296.11	-0.72
	2y	284.04	275.48	0.36
	3y	280.63	258.46	0.91
	5y	269.98	246.55	0.95
	7y	260.52	267.36	-0.27
	10y	247.59	340.35	-3.85
Serial correlation yields ρ_{Y^n}	1y	0.89	0.91	-0.96
	2y	0.90	0.92	-0.66
	3y	0.91	0.92	-0.36
	5y	0.92	0.92	-0.29
	7y	0.92	0.94	-1.01
	10y	0.92	0.95	-1.79
Standard deviation excess returns σ_{x^n}	2y	127.11	101.27	2.23
	3y	239.72	201.37	1.91
	5y	521.33	384.68	2.77
	7y	676.19	571.83	1.58
	10y	895.49	917.47	-0.24

$$\hat{g}=0.0536$$

Model III with RE: RE Regressions

Statistic	Maturity	Data	Model IR	t -stat	Model RE	t -stat
EH coefficient b_{xr}^n	2y	1.37	0.07	1.65	-0.52	2.40
	3y	3.10	1.39	1.03	-0.96	2.46
	6y	10.84	5.95	1.19	-1.98	3.12
	8y	15.53	9.70	1.09	-2.49	3.36
	11y	24.61	16.60	1.08	-3.11	3.74
RE coefficient b_s^n	2y	0.89	0.09	4.05	-0.52	7.11
	3y	1.61	0.11	3.87	-0.96	6.65
	6y	3.02	0.13	2.06	-1.98	3.58
	8y	1.87	0.13	0.76	-2.49	1.90
	11y	3.91	0.11	1.00	-3.11	1.85