# The Power of Long-Run VARS

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Second International Conference in memory of Carlo Giannini

### Introduction/Summary

- Question: Can structural long-run VARs be used to help us choose amongst models? (statistical size and power)
- Method: Simulate many dataset from macro models and ask how frequently these VARS reject different null hypotheses. (some false)
- Conclusions: long-run VARs can frequently reject false models. More useful than the previous literature conjectured.

# Statistical Properties

#### Size

- How often is a true null hypothesis mistakenly rejected?
- Researchers set a threshold for acceptable rejection rates.

#### Power

- How often is a false null hypothesis correctly rejected?
- A good test would have a high rejection rate for false models.
- Researchers less frequently report information about power.

### Literature on long-run VARs

- Simulate data from RBC model and estimate a long-run VAR on simulated data.
- Chari Kehoe and McGrattan CKM claim that structural VARs with long-run restrictions do "not allow a researcher to distinguish between promising and unpromising classes of models"
- Christiano Eichenbaum and Vigfusson CEV
  - A response and critique of CKM.

# CEV on size

- Simulate data from macro model.
- For each dataset, estimate impulse response and associated standard error. Construct confidence interval.
- Ask how frequently do we reject true model using these estimated confidence intervals.
- Results: Fairly good size, but we reject somewhat more often than the pre-declared threshold (nominal size).

### Power

- Less is known about power.
- Tests can have good size but poor power.
- This paper uses methods in CEV to study power. A sequel but a necessary sequel.
- Issues still to resolve about power, especially given claims by
  - Faust and Leeper
  - Chari Kehoe McGrattan.

Faust and Leeper

Faust and Leeper 1997 JBES

**Proposition 1.** Any test of  $H_0: a_{ijk} = 0$  has significance level greater than or equal to maximum power.

- This true only if you don't restrict the DGP.
- Faust and Leeper

F(1). The simplest solution they demonstrate is to assume that the model driving the data is a VAR with known maximum lag order, K. There is surely some K large enough to

## Discussing Faust and Leeper

- For my simulated datasets, VARs have power greater than size.
- Why? Does this contradict Faust and Leeper?
- No contradiction.
- Faust and Leeper: very general DGPs. In particular, possible spikes at frequency zero.
- Here: DGPs are parameterized macro models, more restricted than DGPs discussed in Faust and Leeper.

# Three DGPs Explored Here

- 1. Flexible price DSGE macro model with no real rigidities.
- 2. Flexible price DSGE models with varying levels of habit persistence and investment adjustment costs.
- 3. Sticky price and wage DSGE models with varying levels of real rigidities.

# First Set of Experiments

- Simulate data sets from a standard RBC macro model. (Many times)
- For each simulated data set, estimate impulse responses and several other statistics.
- For each statistic, ask how often, do you reject
  - True DGP (a standard RBC macro model.) (size)
  - Other macro models that are not the RBC model. (power)

### A Flexible Price DGE Model

- Similar to standard RBC model.
- Three estimated shocks
  - Technology shocks. (permanent)
  - Leisure shocks. (not permanent)
  - Investment tax shocks. (not permanent)
- Allow for habit persistence and investment adjustment costs that depend on the growth rate of investment.

### Estimate a Long-Run VAR

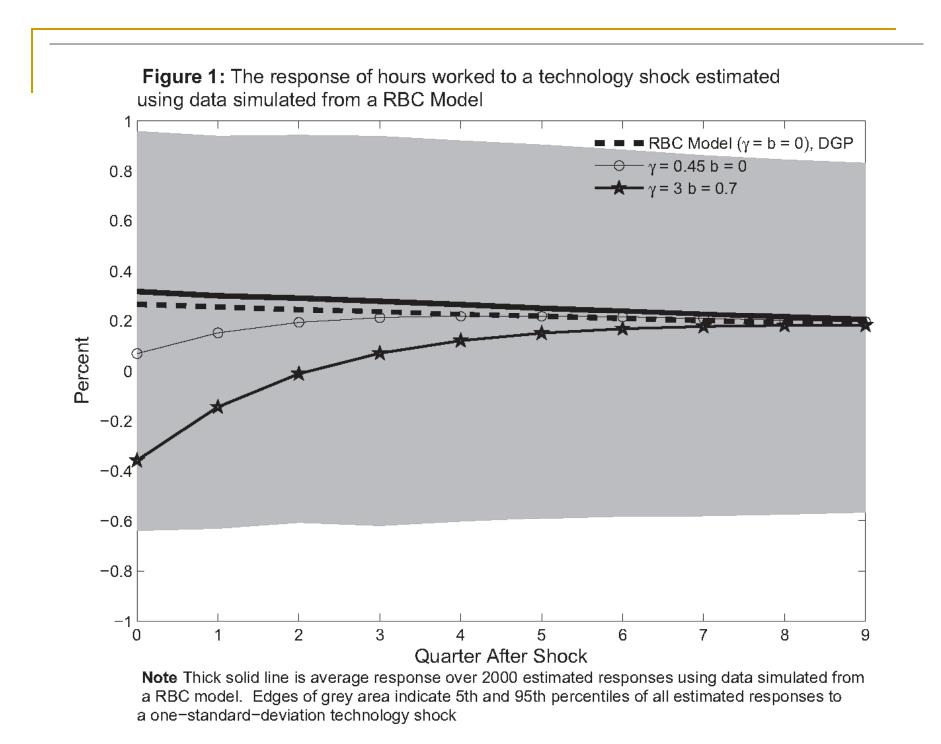
- Standard Estimation of a Long-run VAR.
- Estimate reduced-form VAR (four lags)
  - labor productivity growth
  - level of hours worked.
  - Investment to output ratio.
- Apply the standard long-run identification assumption that only technology shocks affect labor productivity in the long-run.
- Estimate standard error by bootstrap.

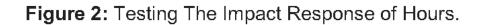
### First Exercise

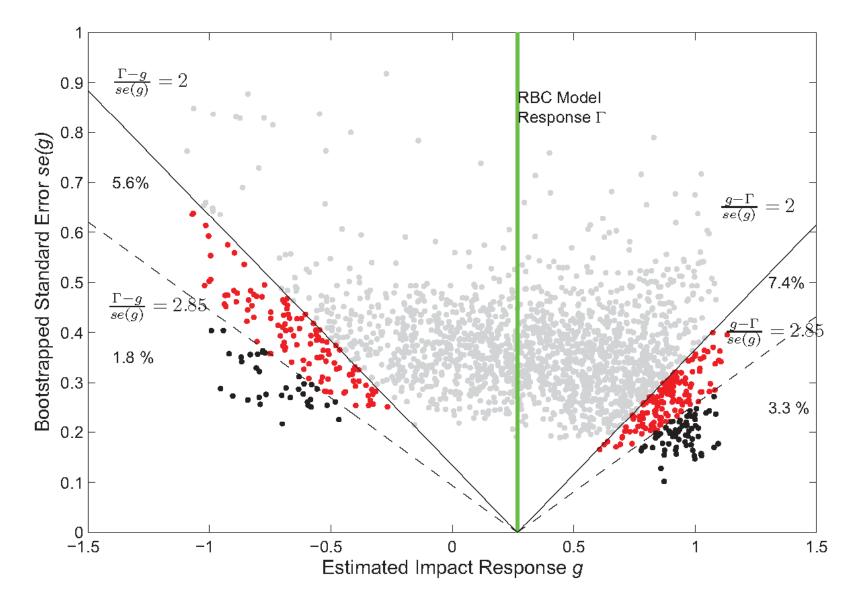
Simulate data from a standard RBC model.

For each simulation, Estimate VAR

- Calculate impulse response of hours to technology shock.
- Calculate Bootstrapped confidence interval.

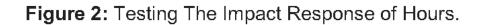


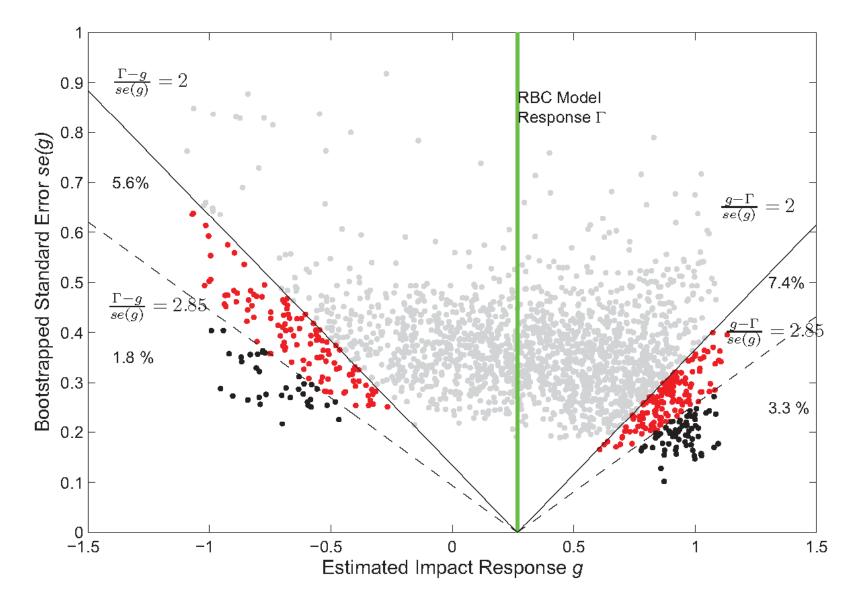




# Getting the Size Right.

- Each simulation of the model
- G(i) response se(i) standard error
- Γ true model response
- Find the 95 percentile of  $|G(i)-\Gamma|/se(i)$
- This is the size-adjusted critical value.
- Asymptotic critical value 1.96
- Here critical value is 2.96
- Rejection Rate Drops from 13 to 5 percent.





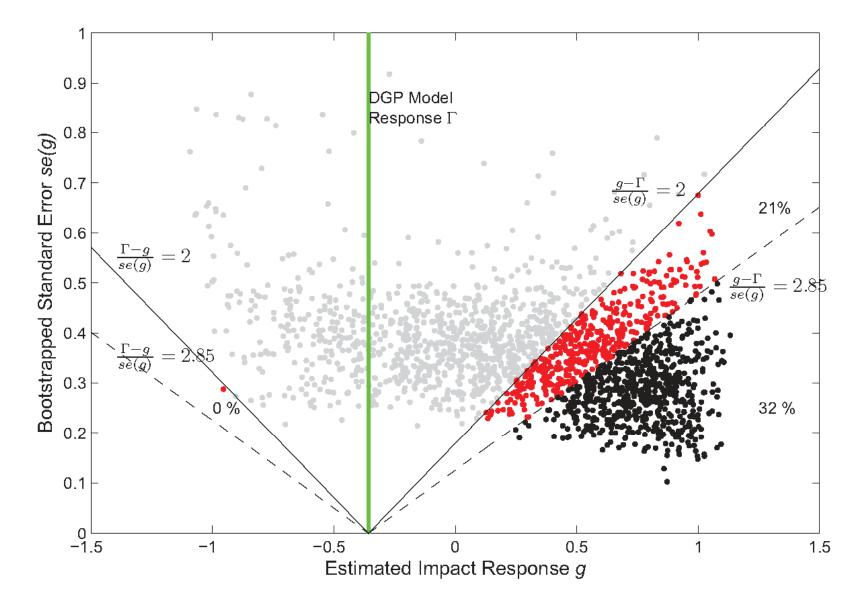
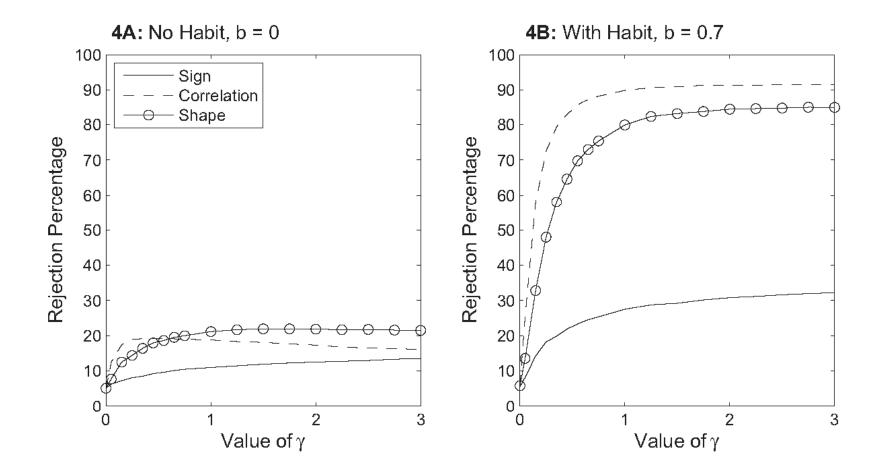


Figure 3: Testing The Impact Response of Hours using a false null hypothesis

### How do we assess power.

- How to interpret rejection rates? Is 32 percent good or bad? (Clearly, not great)
- Easiest to do relative comparisons.
- Test the correlation of output growth.
- Advantages
  - Used before in the literature.
  - No identification necessary. Just a summary statistic from the model.

Figure 4 Rejection Rates For Different Tests when True DGP is RBC Model.



Is there anything else to be done with VARS?

- These results are not overly supportive of using VARs to discriminate amongst models.
- Can we use VARs in different ways to choose between models.
  - Look at shape.
  - Look at other variables



- Look at the response this period and the response six periods later.
- Construct a confidence ellipse based on Wald test.

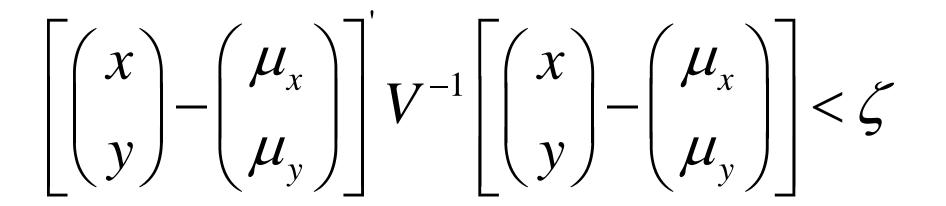


Figure 5: The Shape of The Hours Response

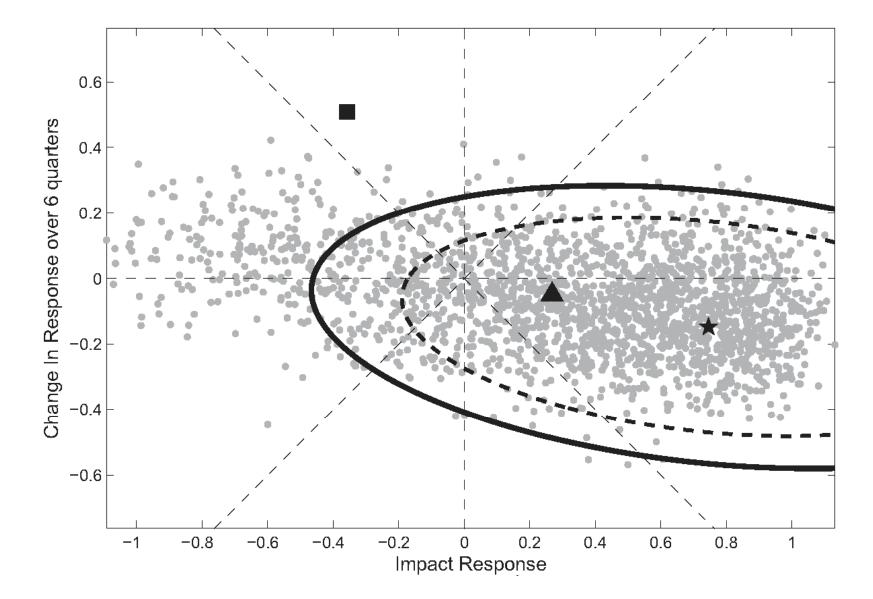
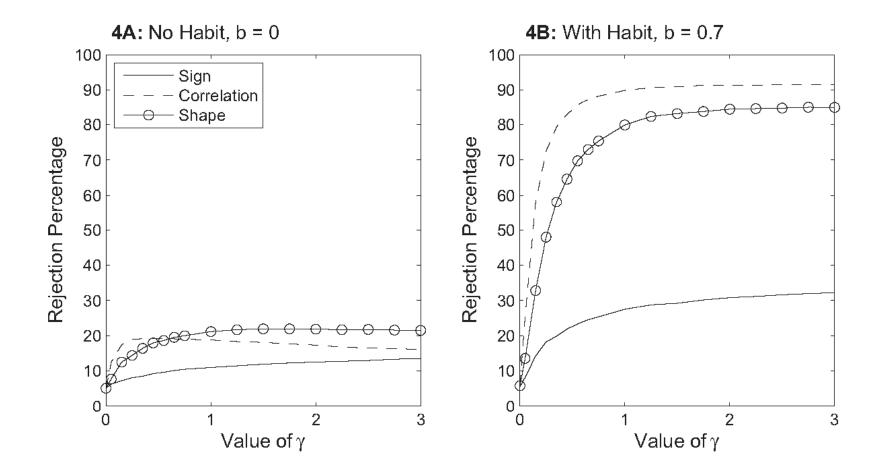
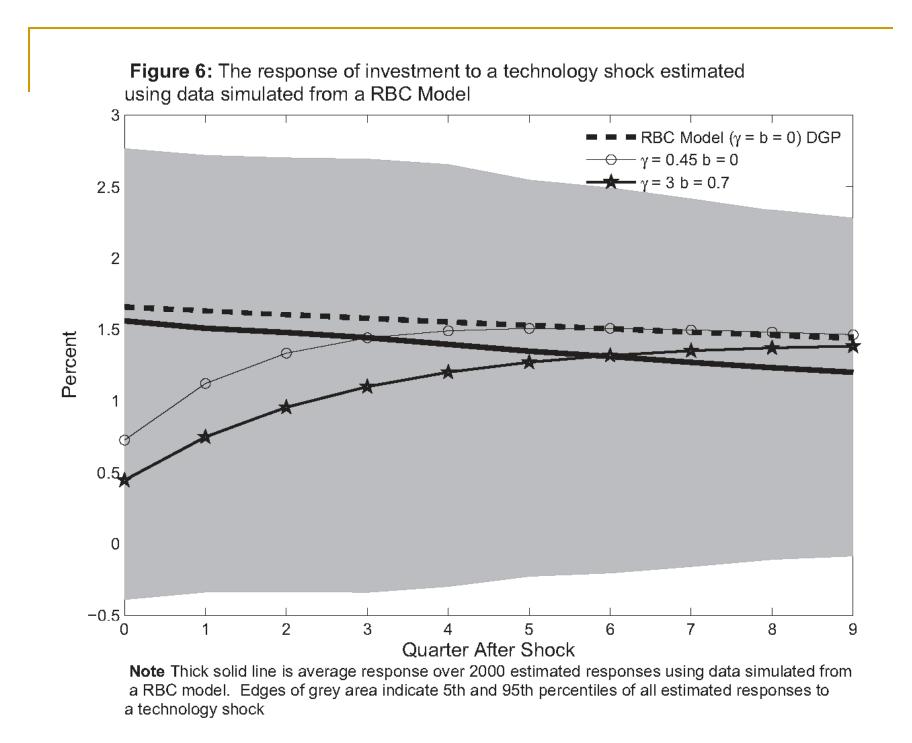


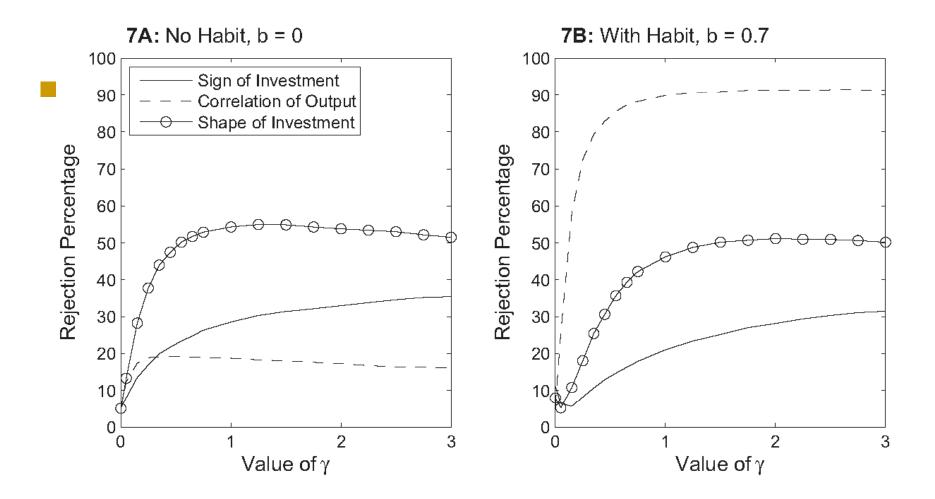
Figure 4 Rejection Rates For Different Tests when True DGP is RBC Model.



### Test Investment Rather Than Hours.

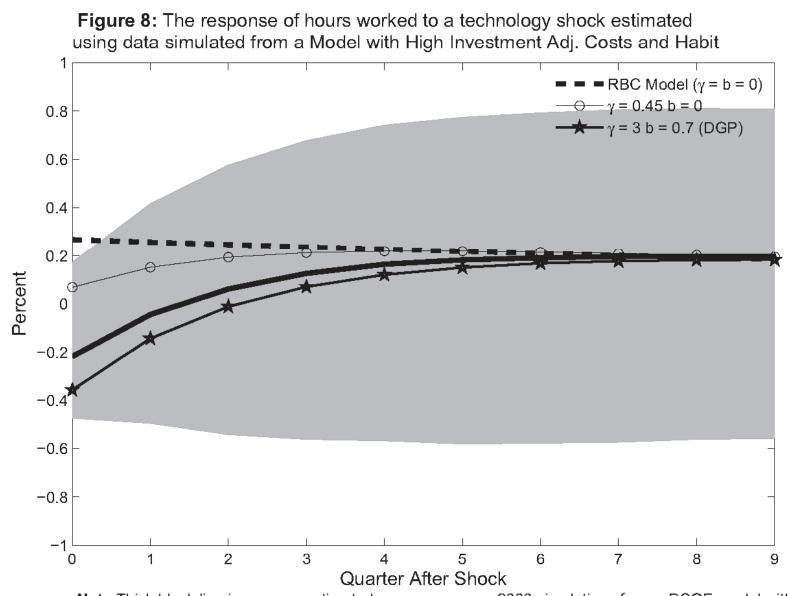
- Estimate VAR of labor productivity, hours and investment over output.
- Standard confidence interval too tight.
  Rejects 13 percent of the time.
- Size-adjusted critical value for sign 2.53
- Size-adjusted critical value for shape 13





Reversing the Role.

- Previous results simulated RBC model and studied rejection rates for macro models with real rigidities.
- Now, do the reverse.
  - Simulate Data from Macro Models with habit and investment adjustment costs.
  - Ask how often one rejects the RBC model.



**Note** Thick black line is average estimated response across 2000 simulations from a DSGE model with high real rigidities. Edges of grey area indicate 5th and 95th percentiles of all estimated responses to a one-standard-deviation technology shock

# Rejection Rates of the RBC Model

Model Parameters		Hours		Investment		Output
b	gamma	Sign	Shape	Sign	Shape	Correlation
0	3	11	100	24	100	0
0.5	0.5	20	43	70	88	79
0.5	1.5	33	100	85	100	92
0.5	3	41	100	87	100	92
0.7	0.5	27	25	83	58	95
0.7	3	58	100	89	100	100

# Sticky Price Models

- DGP (estimated by MLE)
- b=0.4 gamma = 2, theta = 0.75
- Test a variety of models

# When b = 0.4 and gamma = 2

Theta	0.15	0.35	0.55	0.75*	0.9
Hours					
Impact	2	2	3	5	8
Hours					
Shape	2	2	3	5	8
Investment					
Impact	5	5	5	5	8
Investment					
Shape	17	12	9	5	10
Wage					
Impact	32	20	11	5	4
Wage Shape	22	13	8	5	4

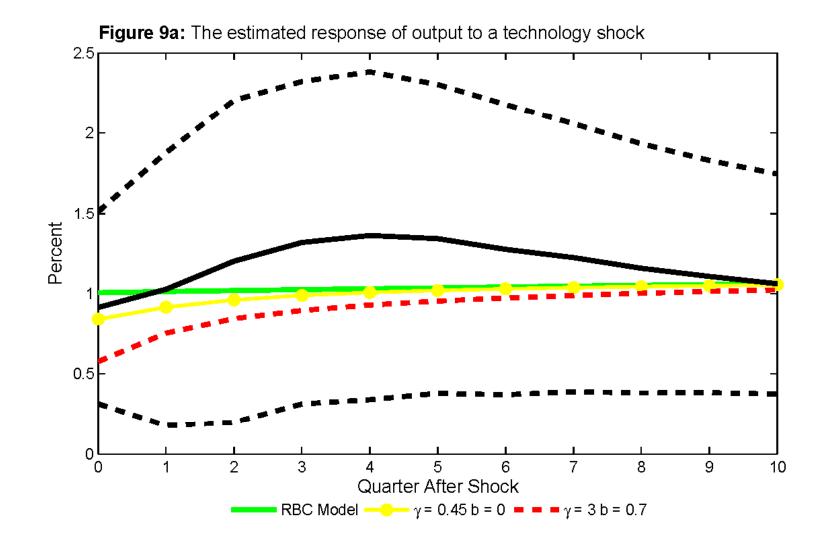
# When b = 0.4 and gamma = 0

Theta	0.15	0.35	0.55	0.75	0.9
Hours					
Impact	9	3	11	100	100
Hours					
Shape	13	3	13	100	100
Investment					
Impact	73	27	19	100	100
Investment					
Shape	61	23	26	100	100
Wage					
Impact	24	17	6	29	94
Wage Shape	17	11	6	43	100

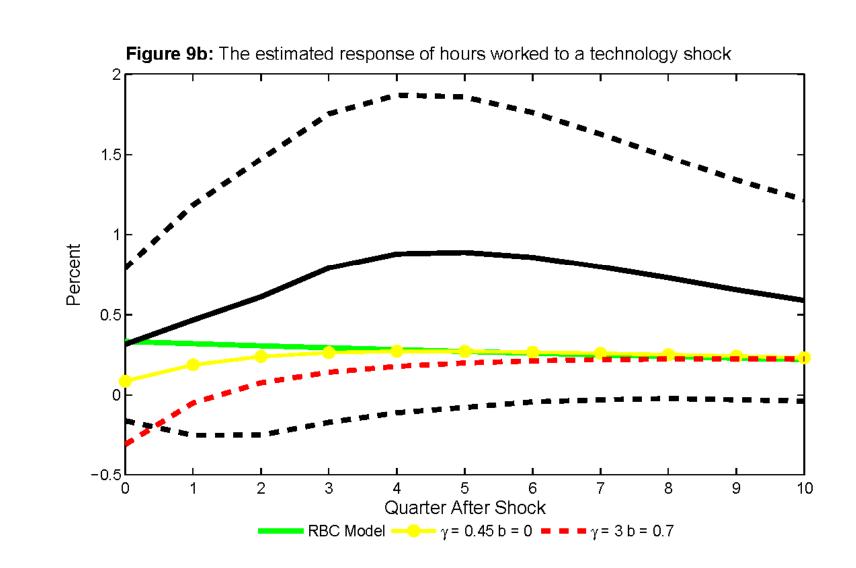
# Empirical Application

#### Estimate VAR Using U.S. Data

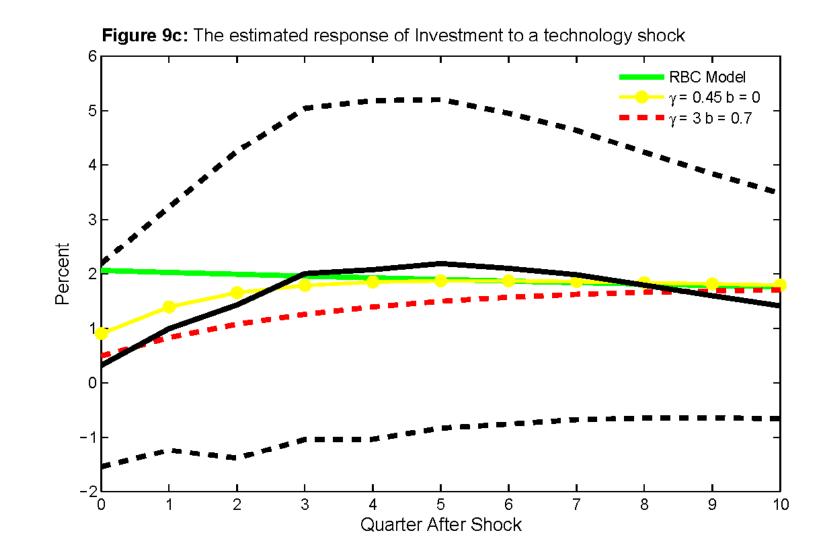
- Labor Productivity
- Hours Worked
- Investment
- **1959-2001**
- Use Critical Values from Simulation
  - Sign 2.8 rather than 2
  - Shape 10 rather than 6



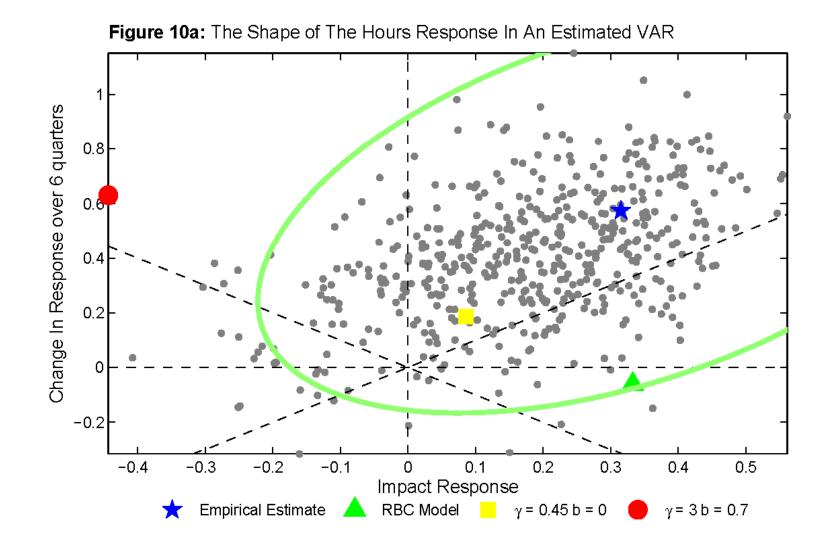
**Note** Thick black line is estimated response using a three variable VAR using U.S. data between 1954 to 2001. Edges of dashed areas indicate confidence interval of 2.8 standard deviations.



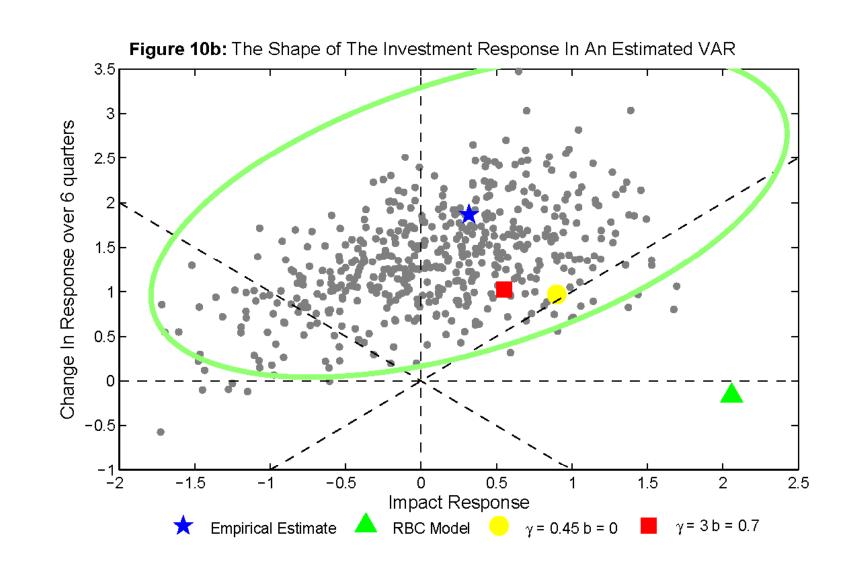
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**Note**Grey dots indicate responses from bootstrap simulations using empirical VAR. Blue ellipse indicates confidence interval around point estimate.



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# Conclusions

- Impulse responses from long-run VARs can reject false models.
- Rejection rates increase the further away the false model is from the true data-generating model.
- Testing share is more powerful than testing just the sign on impact.

# Conclusion (continued)

- Results should encourage us to find creative and new ways to test our models.
- Possible improvements
- Feve and Guay (2009)
- Gospodinov (2008),
- Kascha and K. Mertens (2009)
- E. Mertens (2008).

# Conclusions (continued)

- Overall, given these results on the power and size properties of long-run VARs, we conclude that
- VARs can be useful for discriminating between macro models and, therefore, should continue to be used in developing and testing business cycle theory.