# Discussion of the paper "Forecast evaluation of small nested model sets" by Kirstin Hubrich and Kenneth West

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#### The tests ...

• ... consider *m* regression models (i = 1, ..., m) each of them nesting a benchmark one (i = 0). The null hypothesis is equal predictive accuracy (EPA) across all models, while the alternative postulates that at least one model has a lower mean square prediction error than the benchmark

$$\begin{array}{rcl} H_{0} & : & \sigma_{0}^{2} = \sigma_{1}^{2} = ... = \sigma_{m}^{2}, \\ H_{A} & : & \max\left(\sigma_{0}^{2} - \sigma_{1}^{2}, ..., \sigma_{0}^{2} - \sigma_{m}^{2}\right) > 0. \end{array}$$

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• Why "small set" of nested models? White (2000)'s "reality check" was mainly intended for guarding against extensive data mining. How does test performance deteriorate as m gets larger?

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- Why "small set" of nested models? White (2000)'s "reality check" was mainly intended for guarding against extensive data mining. How does test performance deteriorate as m gets larger?
- Can use the test to select the "best" model?

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## The statistics...

 ... are based on comparing the average squared prediction error plus some adjustment (that -for nested models- helps to re-center the limiting distribution),

$$\widehat{f}_{i,t+1} = \widehat{e}_{0,t+1}^2 - \widehat{e}_{i,t+1}^2 + \left(\widehat{y}_{0,t+1} - \widehat{y}_{i,t+1}\right)^2$$
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• As the limiting distribution of  $\overline{f}_i \equiv P^{-1} \sum_{s=1}^{P} \widehat{f}_{i,t+s}$  is not too badly approximated by a Gaussian, two Wald-type statistics are proposed in the paper: a quadratic form in the vector of the  $\overline{f}_i$  's (called  $\chi^2 (adj)$  statistic) and the  $\widehat{z}$  statistic

$$\widehat{z} = \max\left(P^{1/2}\overline{f}_1/\sqrt{\widehat{v}_1},...,P^{1/2}\overline{f}_m/\sqrt{\widehat{v}_m}
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where  $\hat{v}_i$  is an estimate of the long-run variance of  $\hat{f}_{i,t+1}$ . The approximate null distribution of the  $\hat{z}$  statistic can be easily simulated.

Size-power tradeoff of testing against many alternative models

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# Size-power tradeoff of testing against many alternative models

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- However, as m gets larger the simultaneous comparison is "diluted" by adding a lot of randomness -> inevitable loss of power
- I do not necessarily share the opinion (in the empirical section) that unemployment does not really help predicting euro-area inflation because the forecasts from that model are not significantly better in a 5-model comparison (while they appear better in pairwise tests).

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Testing equal predictive ability and testing forecast encompassing ...

• ... is equivalent for the case of nested models.

$$\begin{aligned} \mathsf{M}_{0} &: \quad \widehat{y}_{0,t+1} = P(Y|X_{0}) \\ \mathsf{M}_{i} &: \quad \widehat{y}_{i,t+1} = P(Y|X_{0},X_{i}) \end{aligned}$$

If  $X_i$  does not have predictive power for Y then (a) the forecasts from  $M_0$  encompass those from  $M_i$  and (b) the two models have same predictive accuracy.

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$$\widehat{y}_{c,t+1} = (1 - \lambda) \, \widehat{y}_{0,t+1} + \lambda \widehat{y}_{i,t+1}$$

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• In fact, as recalled in the paper, the test of EPA based on the adjusted MSPE's is equivalent to a test of FE  $(H_0 : \lambda = 0)$ .

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- Getting critical values would be more complicated as cannot simply simulate from a multivariate normals with an estimated correlation structure. But a bootstrap approximation should go through, like in Hansen (2005). An idea of the order of magnitude of the power gain can be obtained looking at the simulated power functions computed in Busetti, Marcucci and Veronese (2009) for m = 1

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# Power functions of the MSPE, MSPE-adj and other FE tests (m=1)



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- The "joint FE test" of a benchmark model against several alternatives provided here should be useful also for the case of non-nested model. Asymptotic Gaussianity OK.
- Of course, EPA and FE are no longer equivalent. In particular a forecasting model can contain useful information even when its predictive accuracy is relatively bad (but FE by M<sub>0</sub> of M<sub>1</sub> implies that MSPE<sub>0</sub> ≤ MSPE<sub>1</sub>).

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- The "joint FE test" of a benchmark model against several alternatives provided here should be useful also for the case of non-nested model. Asymptotic Gaussianity OK.
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- It is also interesting that FE tests retain some advantage over the standard EPA tests for out-of-sample model selection. I have this example, taken from Busetti, Marcucci and Veronese (2009):

#### The set-up

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$$\begin{array}{lll} y_{t} = & \mu_{y} + \phi_{y} y_{t-1} + \beta x_{t-1} + \varepsilon_{t}, & \varepsilon_{t} & \sim IN(0,1) \\ x_{t} = & \mu_{x} + \phi_{x} x_{t-1} + u_{x,t} & u_{x,t} \sim IN(0,q_{x}^{2}) \\ w_{t} = & x_{t} + u_{w,t} & u_{w,t} \sim IN(0,q_{w}^{2}\sigma_{x}^{2}) \end{array}$$

So  $w_t$  and  $x_t$  are positively correlated with  $ho_{_{XW}}=1/\left(1+q_w^2
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So  $w_t$  and  $x_t$  are positively correlated with  $\rho_{xw} = 1/(1+q_w^2)$ . • Let  $M_X$  be the true model and  $M_W$  be a misspecified one.

$$M_X : P(Y|1, Y_{-1}, X) M_W : P(Y|1, Y_{-1}, W)$$

Assume that  $\beta \neq 0$ . The models are non nested (although, if  $\rho_{_{XW}} \rightarrow 1$  the two forecasts coincide)

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- The rejection frequencies of  $H_0$  are reported in the graph against the value of  $\rho_{xw}$  (if  $\rho_{xw} \rightarrow 1$  rejection frequencies $\rightarrow$ size).



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In practice, we may have an "economic" model that provides (slightly) worse predictions than others. The FE test can help discriminate whether the worse performance is just due to randomness or hot.
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- The paper provides an important contribution towards evaluating the out-of-sample performance of several (nested) models. The tests are neat and easily applicable!
- Perhaps one might design more powerful alternative tests but a cost of a substantial complication which could inhibit the actual use of them
- The idea of a joint FE test of a benchmark model against various alternatives should be kept in mind also in the context of non nested model comparisons